

JGR Solid Earth

RESEARCH ARTICLE

10.1029/2021JB023498

Key Points:

- Statistical procedure to detect directivity effects in small-to-moderate earthquakes analyzing the aleatory residuals of the ground motion
- The directive events are the 36% of the data set and the distribution of rupture direction is aligned with the strikes of the faults system
- The observed directivity is a frequency-dependent band-limited phenomenon, which can occur up to five times the value of corner frequency

Supporting Information:

Supporting Information may be found in the online version of this article.

Correspondence to:

L. Colavitti,
leonardo.colavitti@ingv.it

Citation:

Colavitti, L., Lanzano, G., Sgobba, S., Pacor, F., & Gallovič, F. (2022). Empirical evidence of frequency-dependent directivity effects from small-to-moderate normal fault earthquakes in Central Italy. *Journal of Geophysical Research: Solid Earth*, 127, e2021JB023498. <https://doi.org/10.1029/2021JB023498>

Received 29 OCT 2021
Accepted 19 MAY 2022

Author Contributions:

Conceptualization: Leonardo Colavitti, Giovanni Lanzano, Francesca Pacor
Data curation: Leonardo Colavitti, František Gallovič
Formal analysis: Leonardo Colavitti, Giovanni Lanzano
Funding acquisition: Francesca Pacor
Investigation: Leonardo Colavitti, Giovanni Lanzano, Sara Sgobba, Francesca Pacor

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Empirical Evidence of Frequency-Dependent Directivity Effects From Small-To-Moderate Normal Fault Earthquakes in Central Italy

Leonardo Colavitti¹, Giovanni Lanzano¹, Sara Sgobba¹, Francesca Pacor¹, and František Gallovič²

¹Istituto Nazionale di Geofisica e Vulcanologia, Milano, Italy, ²Department of Geophysics, Charles University, Faculty of Mathematics and Physics, Prague, Czech Republic

Abstract Rupture directivity and its potential frequency dependence is an open issue within the seismological community, especially for small-to-moderate events. Here, we provide a statistical overview based on empirical evidence of seismological observations, thanks to the large amount of high-quality seismic recordings (more than 30,000 waveforms) from Central Italy, which represents an excellent and almost unique natural laboratory of normal faulting earthquakes in the magnitude range between 3.4 and 6.5 within the time frame 2008–2018. In order to detect an anisotropic distribution of ground motion amplitudes due to the rupture directivity, we fit the smoothed Fourier Amplitude Spectra (FAS) cleared of source-, site- and path-effects. According to our criteria, about 36% of the analyzed events (162 out of 456) are directive and the distribution of rupture direction is aligned with the strikes of the major faults of the Central Apennines. We find that the directivity is a band-limited phenomenon whose width may extend up to five times the corner frequency. The results of this research provide useful insights to parameterize directivity, to be explicitly implemented in future ground motion modeling and scenario predictions.

Plain Language Summary In seismology, directivity is one of the source phenomena that causes large spatial variability of earthquake ground motions and is related to the features of the rupture propagation along the fault. The importance of this effect is very well known for large-magnitude events ($M > 6$), while it is still an open issue for small-to-moderate events. The aim of this paper is to recognize directive events and quantify their strength in Central Italy, using a large data set of earthquakes in the magnitude range 3.4–6.5 that occurred between 2008 and 2018. We find that about 36% of the analyzed events are directive, caused by rupture propagations oriented along the NW-SE alignment of the Central Apennines fault systems. Furthermore, we find that directivity is a band-limited phenomenon and that as the directivity gets stronger, the frequency band becomes wider. Our contribution provides a useful insight with the possibility to improve the parametrization of directivity within the empirical seismic hazard assessment.

1. Introduction

Directivity effect of an earthquake is the focusing of the radiated seismic wave energy due to the rupture propagation along the fault (Anderson, 2007; Ben-Menahem, 1961; Boatwright, 2007; Joyner, 1991). Earthquake directivity represents the analogue of the Doppler effect for sound and light waves (Douglas et al., 1988; Pacor, Gallovič, et al., 2016), which shifts the frequency of a moving oscillator to higher frequency when the oscillator moves toward an observer, and lower frequency when it moves away.

This phenomenon, which represents one of the key factors in featuring the spatial distribution of the seismic shaking, produces azimuthal and spectral variations in the ground motion, that can be used to infer information on both the orientation of the fault plane and on the modes of rupture propagation (Abercrombie et al., 2017). Furthermore, the quantification of the directivity-induced amplifications has important consequences in seismic hazard assessment, in terms of ground-motion amplitude and associated variability (Chioccarelli & Iervolino, 2014; Spagnuolo et al., 2012). Although the importance of directivity is widely recognized for both seismological studies on earthquake sources and engineering applications, a clear picture of how strongly and how often it occurs is not yet available.

Methodology: Leonardo Colavitti, Giovanni Lanzano, Sara Sgobba, Francesca Pacor
Project Administration: Francesca Pacor
Resources: Leonardo Colavitti, Giovanni Lanzano, Sara Sgobba, Francesca Pacor, František Gallovič
Software: Leonardo Colavitti, Giovanni Lanzano
Supervision: Leonardo Colavitti, Giovanni Lanzano, Francesca Pacor
Validation: Leonardo Colavitti, Giovanni Lanzano, František Gallovič
Visualization: Leonardo Colavitti, Giovanni Lanzano, Sara Sgobba, Francesca Pacor, František Gallovič
Writing – original draft: Leonardo Colavitti
Writing – review & editing: Leonardo Colavitti, Giovanni Lanzano, Sara Sgobba, Francesca Pacor, František Gallovič

In the 90s, the directivity phenomenon was initially observed and modeled for large events, such as the two recordings of the 1992 Mw 7.3 Landers earthquake at Lucerne and Joshua Tree stations (Bernard et al., 1996; Pacor et al., 2005; Velasco et al., 1994) or the 1994 Mw 6.7 Northridge earthquake (Abrahamson & Somerville, 1996; Dreger, 1994; Somerville et al., 1996), since earthquake directivity was combined with the effects of the finite size of the fault. Over the years, with the development of the broad-band seismic networks, the data quantity and quality have improved considerably. For this reason, nowadays we are able to observe directivity even for moderate-sized (Boatwright, 2007; Calderoni et al., 2015; Convertito et al., 2016; Courboux et al., 2013; López-Comino et al., 2012; McGuire, 2004; Pacor, Gallovič, et al., 2016; Seekins & Boatwright, 2010; Tan & Helmberger (2010); Wen et al., 2015) and small earthquakes (Chen et al., 2010; Tomic et al., 2009; Yamada et al., 2005). Directivity in small earthquakes is not an isolated phenomenon, and recent studies of large data sets in California (Ross et al., 2020) and Japan (Yoshida, 2019) have shown that asymmetric ruptures (unilateral) seem predominant with respect to symmetric ones (bilateral), occurring in more than 60% of the cases analyzed.

One of the most common techniques to capture directivity effects of smaller earthquakes is to measure the duration of the source pulse (called Apparent Source Time Function) at each site and then model it by a line source (e.g., Tomic et al., 2009; Folesky et al., 2016 among the others). Trugman (2022) has recently proposed an innovative Bayesian technique to isolate source spectra using a generalized spectral decomposition inversion and applied it to magnitude M5 in the Southern California to infer some source properties, including directivity effects.

Much more debated is the question of a prevailing direction of rupture propagation in a given region or during a seismic sequence. On this aspect, the results are controversial, as are the physical causes that might determine a preferred direction. For instance, for the case of the 2012 Emilia seismic sequence in Italy, Convertito et al. (2013) observed that a variation of permeability might encourage the fracture in a preferred direction as a consequence of a local increase of the pore pressure and fluid flow rates: in this context, the source directivity can also enhance changes in permeability (López-Comino et al., 2021).

The effects of directivity on ground motion are magnitude and frequency dependent (Bernard et al., 1996). At low frequencies and for moderate-to-strong events, directivity can be responsible for coherent, potentially very destructive pulses with large amplitudes, while at high frequencies and for small-to-moderate events, the most evident effect is given by the shift of the corner frequencies that can result in high-frequency energy arrivals in short time intervals (Anderson, 2007). Modeling of the 2014 Mw 6 Napa earthquake by Gallovič (2016) suggested a preference of model with weak or no directivity effect at high frequencies to correctly explain the azimuthal variations of the strong ground motions. Based on empirical observations from seismic recordings, Pacor, Gallovič, et al. (2016) showed that high-frequency directivity can weaken due to source complexity even for small events.

In the context of empirical ground motion models (GMMs), directivity was clearly seen in the residuals of most earthquakes collected worldwide, such as in NGA-West (Rowshandel, 2010) or other recent near-source databases (Pacor et al., 2018; Sgobba, Felicetta, et al., 2021). Many researchers in the past have empirically modeled directivity effects by fitting the azimuthal variation of within-event residuals to get unbiased prediction (Sgobba, Lanzano, et al., 2021; Shahi & Baker, 2011; Somerville et al., 1997).

With the increasing dissemination of nonergodic approaches for ground-motion modeling in highly sampled areas (e.g., Baltay et al., 2017; Kuehn et al., 2019; Lavrentiadis et al., 2021; Sgobba, Lanzano, & Pacor, 2021), thanks to the use of repeated observations, it is now possible to decompose further the residuals into systematic effects, also known as random-terms, via a mixed-effect regression approach (Stafford, 2014), permitting better isolation of different sources of the model variability. Yet, in nonergodic models, directivity rupture effects are not repeatable in the sense that directivity from any earthquake does not allow predicting that of another event (Sahakian et al., 2019). Hence, the directivity effect cannot be systematically quantified and removed from the total variability of the ground shaking as random-terms, but it can be found in the leftover regression residuals (i.e., as the remaining record-to-record variability of observed ground motion from model predictions).

Therefore, the frequency dependence of the directivity effects in seismic recordings can be inferred by analyzing the azimuthal and frequency dependence of these residuals terms corrected for systematic contributions related to the source and the sites (e.g., Ameri et al., 2009; Ren et al., 2017) and, in a fully nonergodic approach, also to the path (Sgobba, Lanzano, & Pacor, 2021). Recently, this approach was adopted to recognize directivity

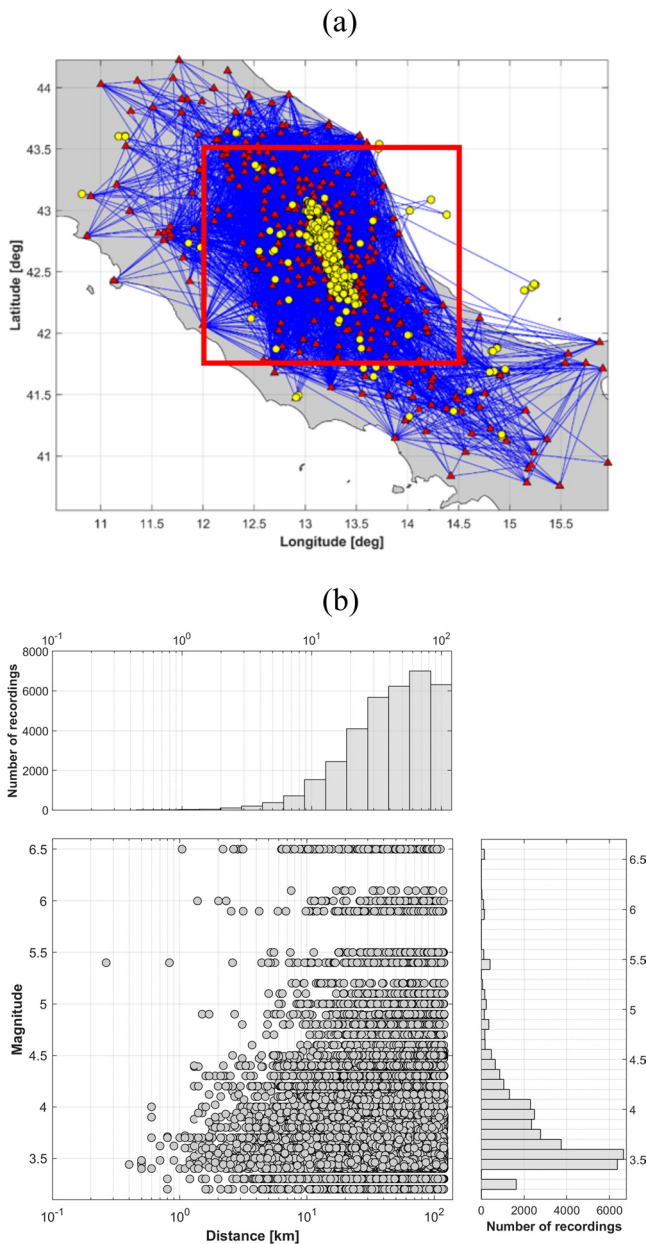


Figure 1. (a) Map of the coverage by ray-paths (lines) connecting events (yellow circles) and stations (red triangles). Red box indicates the study area; (b) Plot magnitude-distance from the epicenter. If the Joyner-Boore distance is not available, we use the epicentral distance.

effects during some events of the 2016–2017 Central Italy seismic sequence (Calderoni et al., 2015, 2017; Luzi et al., 2017; Ren et al., 2017; Wang et al., 2019), revealing significant rupture-propagation complexity, both for small and moderate earthquakes.

In light of the above, it is clear that the literature studies on directivity modeling are mainly carried out with the aim of interpreting the spatial patterns of specific recorded seismic sequences or in order to find seismological features on simulated data. However, there is a lack of knowledge about the main statistical features related to directivity effects on a more extended data set covering different earthquake scenarios and wide frequency intervals, gaining insights to empirically model ground shaking and to provide more physical constraints to their final predictions. Our work thus aims to increase this knowledge by recognizing the statistical features of directivity on the basis of a very dense data set of records in Central Italy, which represents an almost unique natural laboratory for earthquakes occurring on normal faults (Boncio et al., 2000; Bonini et al., 2016; Di Luccio et al., 2010).

With this aim, we analyze the empirical frequency and azimuthal dependence of directivity imprinted in the distribution of the residuals between the observed and predicted data. In this work, we calibrate a fully nonergodic GMM, in terms of Fourier Amplitude Spectra (FAS). The relaxation of the ergodic assumption represents a relevant step toward the empirical modeling of directivity, as it allows to remove all the systematic components of variability, related not only to the event and site, but also to the source-to-site path and source region. Indeed, 60%–70% of the aleatory variance of the ground motion predictions can be associated with systematic effects (Abrahamson et al., 2019). The nonergodic technique thus leads to better isolating the directivity contribution in the random term of the remaining residuals. The approach is entirely based on empirical observations, thanks to the large amounts of high-quality seismic recordings following the two main seismic sequences of L'Aquila in 2009 (Chiarabba et al., 2009; Walters et al., 2009) and Amatrice-Norcia in 2016 and 2017 (Chiaraluce et al., 2017; Pizzi et al., 2017).

The paper is outlined as follows: first, the data set and the ground motion model used for the analysis are introduced, then directivity effects are modeled through an empirical approach based on residual decomposition, and the results are statistically analyzed. Finally, the main parameters controlling the phenomenon are identified for their potential implementation within ground motion predictive models in the field of engineering seismology such as for shaking scenarios and hazard assessments.

2. Data Set

The data set is the same as used by Sgobba et al. (2021) and consists of high-quality accelerometric and velocimetric waveforms related to stations and earthquakes located in Central Italy since 2008. The tectonic setting of this region is complex in terms of mechanical discontinuities and rheological properties (Carafa & Barba, 2011; Chiarabba et al., 2018), featuring mainly normal faults that caused several seismic sequences in the last 20 years (i.e., 1997–1998, Umbria-Marche Mw 6.0; 2009, L'Aquila Mw 6.1; 2016–2017, Amatrice-Visso-Norcia Mw 6.5; 2018, Muccia Mw 4.6). In order to guarantee more stable and robust analyses, we additionally set a threshold to at least 10 recordings for each station, according to the approach performed by Lanzano et al. (2017).

Figure 1a shows the high density of events and stations in the study region, featuring a dense azimuthal sampling of source-to-site ray paths. The overall data set analyzed in this study consists of more than 30,000 waveforms of

456 earthquakes in the magnitude range between 3.2 and 6.5, and 460 stations within 120 km from the epicenter, as demonstrated by the histograms in Figure 1b. The recordings' distribution is strongly characterized by low magnitudes (panel on the right of Figure 1b), with the most frequent value equal to 3.5 and with 70% of the recordings having a magnitude below 4. More than 80% of the bulk data is included in the distance range between 20 and 120 km (top panel on Figure 1b).

Starting from this data set, the study is conducted on the Fourier spectra smoothed using the Konno and Ohmachi (1998) algorithm (fixing the smoothing parameter b to 40) of the signal containing the S-phases. For each record, we use a distance-dependent energy criterion to identify S-wave time window (Pacor, Spallarossa, et al., 2016). As time windows of records close to the epicenter could be very short (generally less than 2 s), we fix a minimum length of 4 s in order to guarantee an acceptable spectral resolution above 1 Hz for the shortest time windows. Pre-event noise windows of the same length as the signal windows are used to compute the SNRs and a threshold equal to 3 is selected to remove noisy spectral ordinates.

3. Method

3.1. Nonergodic Ground Motion Model

The proposed GMM is developed for the geometric mean of horizontal peak ground accelerations (PGA) and 69 FAS ordinates in the frequency range logarithmically equispaced from 0.5 to 25 Hz. The model is calibrated via a mixed-effect regression (Bates et al., 2015), providing the estimation of the different repeatable effects on the seismic motion (i.e., source, site, and path), along with the associated aleatory variability.

We assume the same functional form proposed by Sgobba et al. (2021) for the response spectrum to describe PGA and spectral ordinates, Y , at each frequency $Y(f)$:

$$\log_{10} Y = a + F_M(M) + F_R(M, R) + \delta B_e + \delta S2S_{ref,s} + \delta L2L_{source} + \delta P2P_p + \delta W_0 \quad (1)$$

where a , $F_M(M)$, $F_R(M, R)$ represent the fixed effects, and δB_e , $\delta S2S_{ref,s}$, $\delta L2L_{source}$, $\delta P2P_p$ stands for zero-mean Gaussian-distributed random effects.

Parameter a is the offset and F_M describes the scaling with magnitude:

$$\begin{aligned} F_M(M) &= b_1(M_w - M_h) \quad \text{for } M \leq M_h, \\ F_M(M) &= b_2(M_w - M_h) \quad \text{otherwise} \end{aligned} \quad (2)$$

Here, b_1 and b_2 are calibrated coefficients obtained from nonlinear least-square regression, while M_h is the hinge magnitude fixed at 5.0. The term $F_R(M, R)$ is the other fixed term and represents the scaling with distance:

$$F_R(M, R) = [c_1(M_w - M_{ref}) + c_2] \log_{10} \frac{\sqrt{R^2 + h^2}}{R_{ref}} + c_3 \left(\sqrt{R^2 + h^2} - R_{ref} \right). \quad (3)$$

$F_R(M, R)$ is divided into a contribution due to the geometrical spreading also computed by nonlinear regression (including magnitude-dependent terms with coefficient c_1 , and magnitude-independent with coefficient c_2) and the anelastic attenuation (described by c_3 , which is typically regionally dependent). M_{ref} is the reference magnitude obtained from a preliminary nonlinear regression. R_{ref} is the reference distance fixed at 1 km, h is the pseudo depth fixed at 6 km, and R is the Joyner-Boore distance for events larger than 5.5, for which the fault geometry was defined. For lower magnitude events, the Joyner-Boore distance is assumed to be equal to the epicentral distance, since the small area of the rupture surface makes the event equivalent to a point-like source.

The random error terms with respect to the median prediction of the GMM in Equation 1 are defined as follows (from Stafford, 2014):

1. δB_e is the between-event error, zero mean, normally distributed residuals, which corresponds to the average bias of recordings of one particular earthquake with respect to the prediction of the fixed effects of Equation 1.
2. $\delta S2S_{ref,s}$ is the site-to-site term, defining the systematic bias of ground motions recorded at each station. This term is computed as a reference ground motion level that is observed on a set of reference rock sites, which

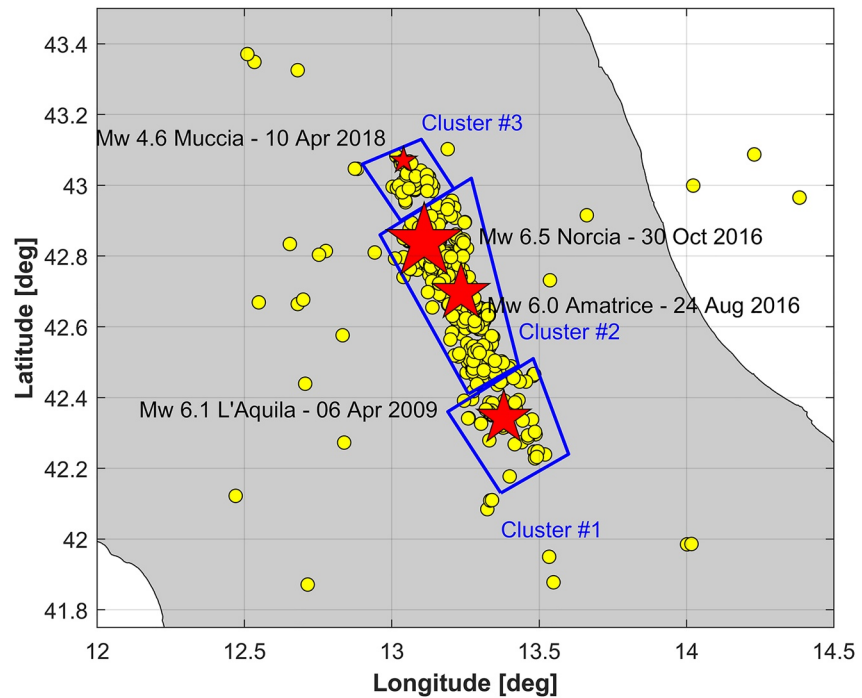


Figure 2. Source areas (with blue dashed boxes) in a zoomed view with respect to the one shown in Figure 1a. From South to North: Cluster #1 (main event: L'Aquila 06/04/2009 - 01:32 UTC), #2 (main event: Amatrice 24/08/2016 - 01:36 UTC and Norcia 30/10/2016 06:40 UTC), and #3 (main event: Muccia 10/04/2018 - 03:11 UTC). Red stars represent the mainshocks, while yellow circles are the other events with $M > 3.2$.

were previously detected in the study area according to several proxies based on geophysical, seismological, and geomorphological features (Lanzano et al., 2020).

3. $\delta L2L_{source}$ represents the source term, describing the systematic bias of the source regions (see further).
4. $\delta P2P_p$ is the path-term, denoting the systematic deviation along each source-to-site path (from each identified source region to the sites) and being related to anisotropy in the properties of the crustal propagation medium.

Details on both fixed terms and random terms can be found in Data Availability Statement.

The source regions that are necessary for the computation of systematic terms $\delta L2L_{source}$ and $\delta P2P_p$ were defined by Sgobba et al. (2021) by performing a spatial-temporal clustering of the events in the data set using the algorithm proposed by Reasenber (1985). In particular, these authors identified six clusters, three of these are located inside our study area (red box in Figure 1a) and are shown in Figure 2. These three clusters are (from South to North): cluster #1, including the 6 April 2009 Mw 6.1 L'Aquila, cluster #2, with the 2016 Mw 6.0 Amatrice and Mw 6.5 Norcia sequence, and cluster #3, including the 2018 Mw 4.6 Muccia sequence. One of the advantages of the nonergodic GMM over the ergodic one is the removal of effects due to the different characteristics of the sources, site effects not modelled in the GMM, or isotropic and anisotropic path effects.

As a result, the leftover residuals δW_0 reflect the remaining variability unaccounted by Equation 1, which should represent the purely random term unrelated to systematic effects. According to Villani and Abrahamson (2015), the nonsystematic effects are connected to the path contributions, while Kotha et al. (2019) associate these remaining record-specific residuals with the four lobes of the S-wave radiation pattern. In our data, the azimuthal pattern of the residuals emerges above the event corner frequency, suggesting its relation to finite source properties (directivity) and not to a point-source feature (radiation pattern).

Therefore, as previously seen in other works (Luzi et al., 2017; Ren et al., 2017; Türker et al., 2022), owing the extensive coverage of the present data set and taking into account the effects that can be modeled according to the fully nonergodic functional (Stafford, 2014), we investigate the azimuthal dependence of the δW_0 terms to recognize a potential signature of source directivity effects.

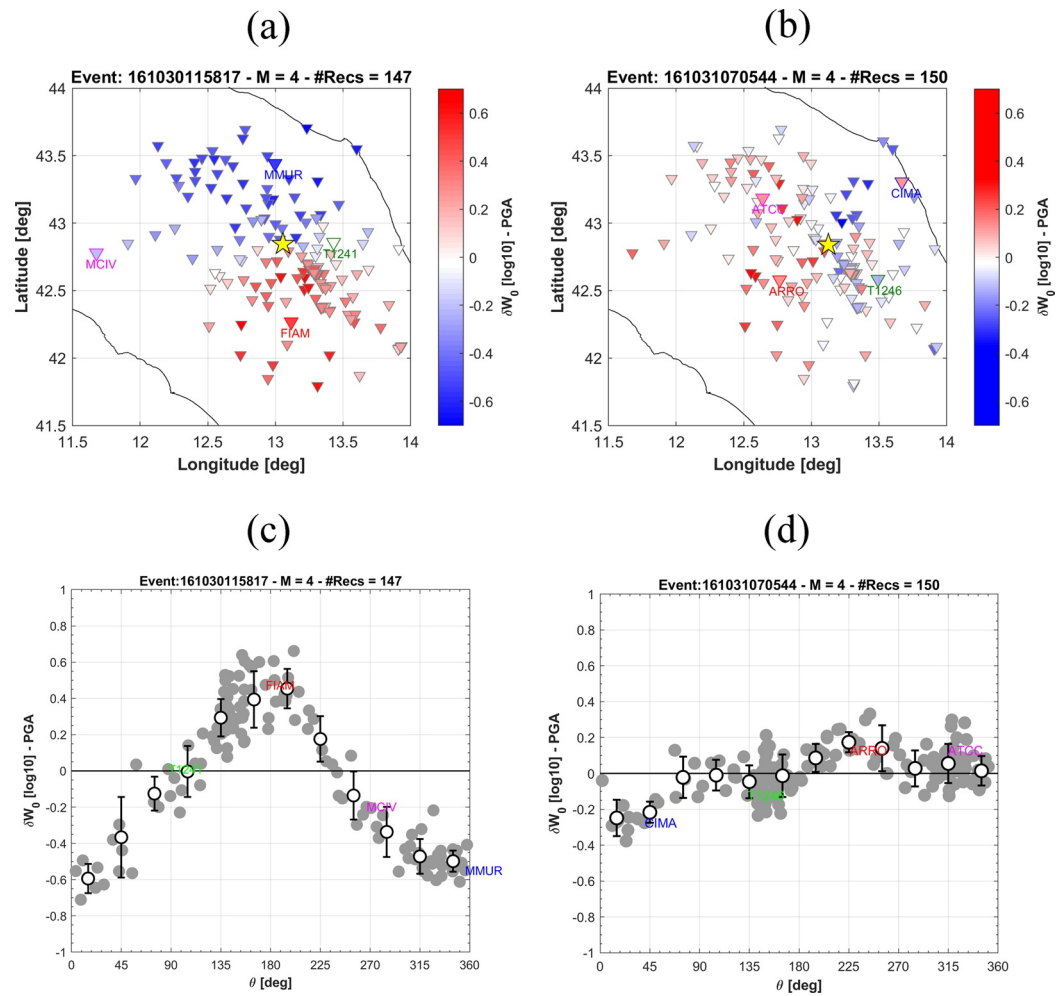


Figure 3. Spatial distribution of the PGA residuals δW_0 for the (a) M4 event of 30 October 2016 at 11:58:17, and (b) M4 event of 31 October 2016 at 07:05:44. Yellow stars are the epicenters, while reverse triangles represent the stations. Blue/red colors indicate negative/positive residual values (under/over-estimation). The labeled stations are used for illustration of the frequency dependence of the residuals in Figure 4. Locations in the four quadrants around the epicenter are highlighted. (c, d) Aleatory residuals δW_0 as a function of θ for the same events in (a) and (b), respectively.

3.2. Azimuthal Variations of the $\delta W_0(\theta)$ Residuals

For each event of the data set and for each frequency, we analyze the azimuthal variation of the aleatory residuals δW_0 . As an example, Figure 3 displays the spatial distribution of the PGA aleatory residuals for two M4 events of 30 October 2016 at 11:58:17 (Figures 3a) and 31 October 2016 at 07:05:44 (Figure 3b), both belonging to the Norcia sequence in Central Italy.

Figure 3a shows a clear spatial pattern with positive values of δW_0 to the South with respect to the epicenter and a well-defined azimuthal distribution (Figure 3c), with peak amplitude of 0.6 at around 180° . Conversely, for the event plotted in Figure 3b, the distribution of δW_0 values, ranging between -0.4 and 0.4 , are rather random without any predominant direction (Figure 3d). The trend of the aleatory residuals relative to the same events of Figure 3 for four selected stations, each one located in a different quadrant is plotted in Figure 4.

For the event of 30 October 2016 (Figure 4a), we observe that the station FIAM, located south of the epicenter, is characterized by positive residuals, while MMUR, lying in the northern sector, is affected by negative residuals on the entire frequency range of investigation. The remaining two stations (MCIV and T1241), located in the eastern and western sectors, have similar trends with values spanning around zero. In contrast, for the event of 31 October 2016 (Figure 4b), no significant variations of the δW_0 can be distinguished from the four investigated

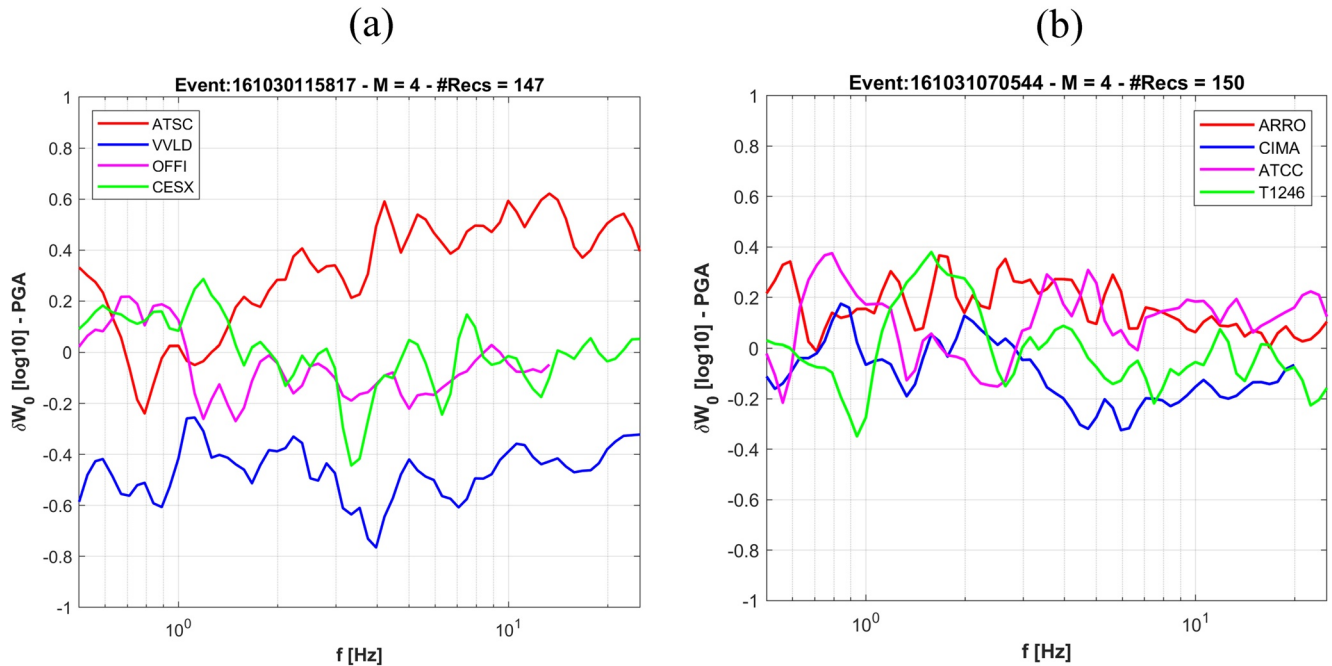


Figure 4. Trend of δW_0 versus frequency for the four stations shown in Figure 3, for (a) the M4 event of 30 October 2016 at 11:58:17, and (b) the M4 event of 31 October 2016 at 07:05:44.

stations. In the whole considered data set, the cases above are only two illustrative examples, but we have several earthquakes which show different behavior of δW_0 with respect to the frequency: some of these events are shown in Figure S1 in Supporting Information S1.

These observations are consistent with directivity effects that can generate spatial patterns, with the largest ground motion amplitude at sites located in the forward-direction with respect to the rupture propagation (Aki & Richards, 1980; Somerville et al., 1997). In light of this, the 30 October 2016 earthquake of Figures 3a and 3c may be classified as a unilateral rupture event, where FIAM station is in the forward-directivity direction, MMUR in the backward-directivity direction, whereas MCIV and T1241 are in neutral positions. On the other hand, the event on 31 October 2016 (Figures 3b and 3d) may be classified as a bilateral-rupture earthquake without pronounced source directivity effects.

3.3. Directivity Models for $\delta W_0(\theta)$

In order to detect and quantify the directivity effects from the aleatory residuals δW_0 , we introduce the following model:

$$\delta W_0 = \delta D2D(\theta) + \delta W_{\text{nodir}}, \quad (4)$$

where $\delta D2D(\theta)$ is the function that fits the residuals δW_0 for each event in the data set as a function of the source-to-site azimuth θ , while δW_{nodir} are the remaining residuals cleared of the directivity effect. Two functional forms are usually adopted to describe the dependence on θ of the $\delta D2D$ term:

1. A cosinusoidal function selected on the basis of the trend of residual distribution with the azimuth θ of the sites (Somerville et al., 1997):

$$\delta D2D(\theta) = A \cos(\theta - \theta_0), \quad (5)$$

where A is the amplitude of the fitting cosine function and θ_0 represents the azimuth at the maximum amplitude.

2. The directivity coefficient C_d based on a simple theoretical rupture propagation model (Ben-Menahem, 1961; Gallovič, 2016; Hirasawa & Stauder, 1965; Pacor, Gallovič, et al., 2016; Ruiz et al., 2011), which is introduced

to account for the rupture propagation on the fault plane. According to Boatwright (2007), the general expression of C_d for a bilateral rupture is the following:

$$C_d = \sqrt{\frac{k^2}{\left[1 - \left(\frac{v_r}{c}\right) \cos(\theta - \theta_0)\right]^2} + \frac{(1-k)^2}{\left[1 + \left(\frac{v_r}{c}\right) \cos(\theta - \theta_0)\right]^2}}, \quad (6)$$

where $\frac{v_r}{c}$ is the Mach number (ratio between rupture and shear-wave velocities, denoted as α in the following), while θ_0 is the azimuth of the rupture direction.

Parameter $k \in (0, 1)$ represents the relative portion of the rupture length in the direction θ_0 . In the case of a unilateral rupture propagating along a narrow fault at a constant velocity v_r , k is equal to 1 and the full expression of C_d reduced to that for the model of Haskell (1964):

$$C_d = \frac{1}{1 - \left(\frac{v_r}{c}\right) \cos(\theta - \theta_0)} \quad (7)$$

Corner frequencies for sites affected by directivity are given by the apparent corner frequency $f_a = f_c C_d(\theta)$, where f_c is proportional to the reciprocal of the rupture duration. Nevertheless, we aim to fit the amplitude of the source spectral ordinates that scale proportionally to C_d^n above frequency f_a , where the exponent n depends on the type of the source model. Omega-squared kinematic rupture models with single corner frequency (e.g., k-squared model by Herrero & Bernard, 1994) suggest $n = 2$, while models with two corner frequencies (e.g., Haskell model, with constant slip and rise time) suggest $n = 1$. In case of purely stochastic models, the directivity disappears and n becomes null (see Gallovič, 2016).

In this second approach, the corrected event-site and path residuals of Equation 4 are fitted by the following function for each frequency:

$$\delta D2D(\theta) = \log_{10} C_d^n(\theta) - \langle \log_{10} C_d^n(\theta) \rangle, \quad (8)$$

where $\langle \rangle$ denotes azimuthal averaging. The fitting variables are the angle of the rupture direction θ_0 , the ratio k , the Mach number α and the exponent n .

3.4. Tuning of Model Parameters

In order to determine the fitting parameters, we calibrate the C_d model using the Levenberg-Marquardt optimization algorithm (Marquardt, 1963), which is commonly adopted to solve nonlinear least squares problems. In the optimization tests, we assume that the explanatory parameters vary within a fixed range of values: n in the interval between 0 and 2, k between 0.6 and 1, α from 0.5 to 1, and θ_0 in the range 0–360°.

While the fitting of the cosine function is straightforward (Sgobba et al., 2021a), the choice of variables in Equation 6 should be treated carefully considering a strong trade-off among some of the fitting parameters. In Figure 5 and Table 1, we show an example of the fitting of the residuals for the event of Figure 3a, which is obtained at three different frequencies (results for other sample events are reported in Figure S2 in Supporting Information S1).

In fact, although we well resolved the angles related to the direction of propagation θ_0 , as shown in Pacor, Gallovič, et al. (2016), similar fitting functions are obtained for different combinations of α and n . Several approaches to deal with this issue have been suggested. Convertito et al. (2017) suggest that the best value is 0.8, after performing a set of inversions for different values ranging between 0.2 and 0.9, selecting as the best model the one resulting in the smallest residual. Ren et al. (2017) suggest that the optimum Mach number α can vary in a range from approximately 0.5 to 1.0, and strongly depends from the value of exponent n considered, as smaller values of n typically result in larger values of the Mach number.

The determination of the optimal values of n still remains ambiguous. In order to automatize the method, the estimation of the values of α , k , and n is done by considering the bulk of the events, although a more proper option would be to do an event-by-event calibration. Therefore, after the calibration of several tests, we applied a

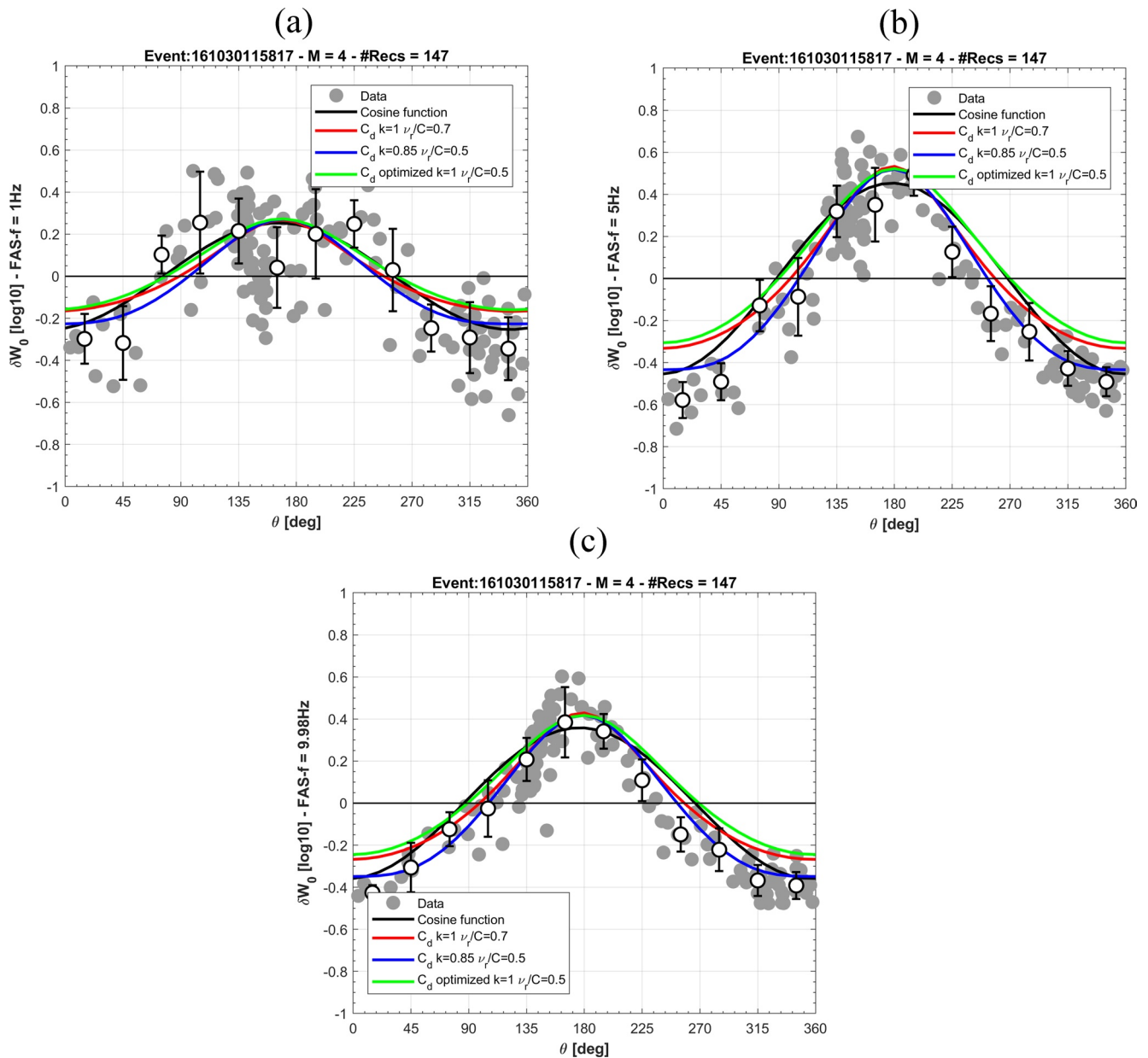


Figure 5. Azimuthal variation of the aleatory residuals for a Mw 4.0 event of 30 October 2016, 11:58:17 at (a) $f = 1$ Hz, (c) $f = 5$ Hz, (e) $f = 10$ Hz. In black, fit with cosinusoidal function (case I), in red fit with $k = 1$ and $\alpha = 0.7$ (case II), in blue $\alpha = 0.5$ and $k = 0.85$ (case III), while with the green curve the fit is estimated with an optimization algorithm (in this case, k tends to 1 and α is close to 0.5, case IV).

sort of compromise between the resulting values of the inversion algorithm (see Figures S3 and S4 in Supporting Information S1) and the parameters suggested by literature to have reasonable value even for a physical point of view. The best performance in the majority of the events being able to also capture the trend of the residuals also in the anti-directive azimuths was obtained when fixing $k = 0.85$ and $\alpha = 0.5$ (blue curve in Figure 5 and case III in Table 1). Since these values are also found to be consistent with common seismological observations (Ren et al., 2017), we decide to adopt this fitting function for the final C_d model.

In Figure 6, we inspect the trend of the directivity parameters as a function of frequency for the two events shown in Figure 3 (for other examples see Figure S5 in Supporting Information S1). As we can observe in Figure 6a, the M4.0 event of 30 October 2016 at 11:58:17 is directive for frequencies greater than 1 Hz as the proxy for the determination coefficient R^2 is consistently greater than 0.5 (third plot at the bottom). For this event, the

Table 1
Values of Amplitude (A or n), Azimuth θ_0 , Coefficient of Determination R^2 and Stand σ for Different Cases Reported in Figure 5; (a) $f = 1$ Hz, (b) $f = 5$ Hz, (c) $f = 10$ Hz

a)				
Case	A or n	$\theta_0(^{\circ})$	R^2	σ
I	0.2534	166.34	0.5338	0.1884
II	0.5770	168.52	0.4436	0.2050
III	1.1736	169.74	0.5024	0.1988
IV	0.9062	168.41	0.5186	0.1996

b)				
Case	A or n	$\theta_0(^{\circ})$	R^2	σ
I	0.4537	178.17	0.8200	0.1410
II	1.1488	178.66	0.8011	0.1415
III	2.2469	179.06	0.8732	0.1289
IV	1.7347	179.87	0.8108	0.1452

c)				
Case	A or n	$\theta_0(^{\circ})$	R^2	σ
I	0.3586	176.26	0.8118	0.1169
II	0.9248	178.19	0.8123	0.1079
III	1.8064	178.87	0.8786	0.1007
IV	1.3872	179.17	0.8200	0.1154

directivity amplifications is larger than 0.2 and 0.4 if we observe the amplitude A from cosine fitting (black curve) and the value of n from the C_d model (blue curve), respectively (Figure 6a at the top). In particular, if we look at C_d model, we find a maximum value of n around 2.5 corresponding to about 4–5 Hz. The θ_0 angle is very stable and independent of frequency, which is a typical and expected behavior for other directive events.

The direction of the rupture propagation does not change with frequencies (Figure 6a in the middle), while the coefficient of determination R^2 follows the frequency trend of the amplification level and thus can be considered a proxy of the significance of the directivity effects in the ground motions. The M4 event of 31 October 2016 (Figure 6b) shows instead small amplification (<0.4) at all frequencies both for cosine and C_d models, the θ_0 is unstable and R^2 is consistently below 0.5, indicating that no directivity occurs for this event.

We observed that the C_d function provides slightly better description of the azimuthal variation of the residuals compared to the simpler cosinus-type relationship. The former reproduces better the asymmetry observed in ground motion amplification of directive sites and the de-amplification of non-directive sites. Moreover, it is more related to the physics of the phenomenon, and thus we prefer the C_d model in the subsequent analyses.

4. Results

We apply the above-described method to all the events of the data set. About 10% of the investigated events (47 out of 456 events) is characterized by median values over frequencies of coefficient of determination R^2 and the value of the exponent n for the function C_d are strongly correlated (Pearson coefficient ρ is always greater than 0.7 and, for frequency values greater than 1 Hz, greater than 0.9), as shown by the graph of the correlation coefficient computed for each frequency analyzed in this study (see Figure S6 in Supporting Information S1). Since in this research we want to focus on directive events only, we introduce a selection criteria described in the following section.

4.1. Identification of Events With Significant Directivity

Considering that for each event the phenomenon of directivity exhibits variable intensity in different frequency bands of variable length (Figure 7), we introduce some assumptions for identifying a directive event. We designate an event as directive if:

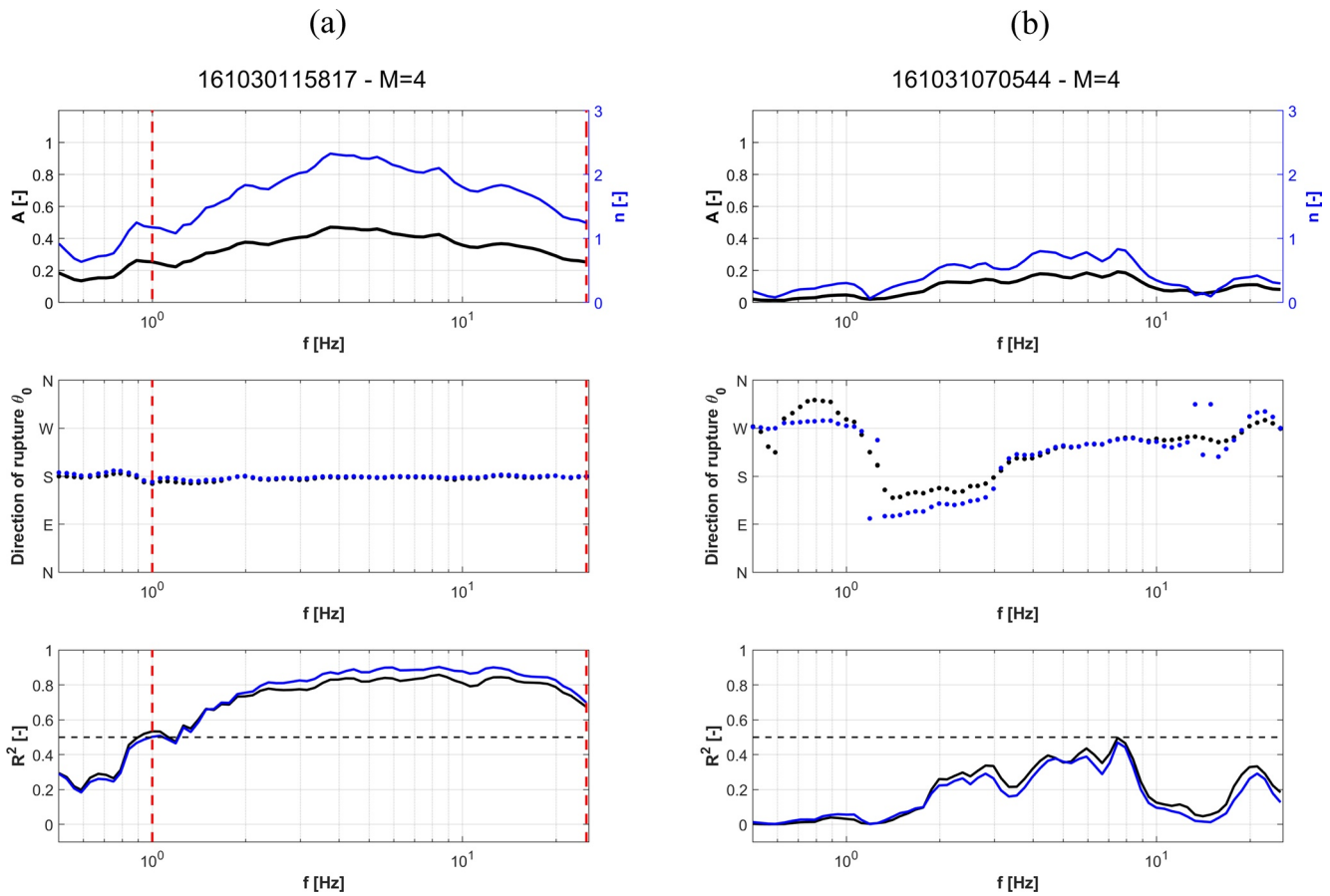


Figure 6. Directivity amplitude (top); direction of rupture propagation (middle); coefficient of determination (bottom) plotted versus frequency. The black curve represents the fitting results with the cosine model, the blue curve with the C_d model. Red lines represent the minimum and maximum frequency within which the proxy R^2 is larger than 0.5. (a) M4 event of 30 October 2016 at 11:58:17; (b) M4 event of 31 October 2016 at 07:05:44.

1. The coefficient of determination R^2 is greater than 0.5 for at least 10% of the frequencies investigated (in our case, 7 out of 69 values).
2. The standard deviation of the angle of the direction of propagation θ_0 is smaller than 20° in the same frequency range.

Figure 7 shows the values of n as a function of the frequencies analyzed for each event within the magnitude range between 3.7 and 4.0 (for the behavior of n with respect to the other magnitude classes, see Figure S7 in Supporting Information S1). Since the frequency ranges analyzed are logarithmically equispaced, the upper part of the graph on x -axis shows the \log_2 of the frequency, while the corresponding value of the frequency content is shown at the bottom of the x -axis. After the selection criteria, the directivity frequency-bandwidth can contain gaps, which are in some cases relatively narrow. In order to guarantee the continuity of the frequency-bandwidth with respect to the physics of the phenomenon, gaps composed of less than five consecutive intervals are filled with the average values in the missing frequency ranges, assuming a tolerance in the R^2 criterion, namely $R^2 > 0.45$ for closing the gap.

4.2. Directivity Event Distributions

According to the criteria defined in 4.1, about 36% of the analyzed earthquakes (162 out of 456) can be classified as directive (details about the list of events can be found in Data Availability Statement). The number of the directive events seems independent of magnitude, being distributed as the entire data set, as illustrated by the histograms reported in Figure 8a. Figure 8b shows the percentage of directive events as a function of frequency

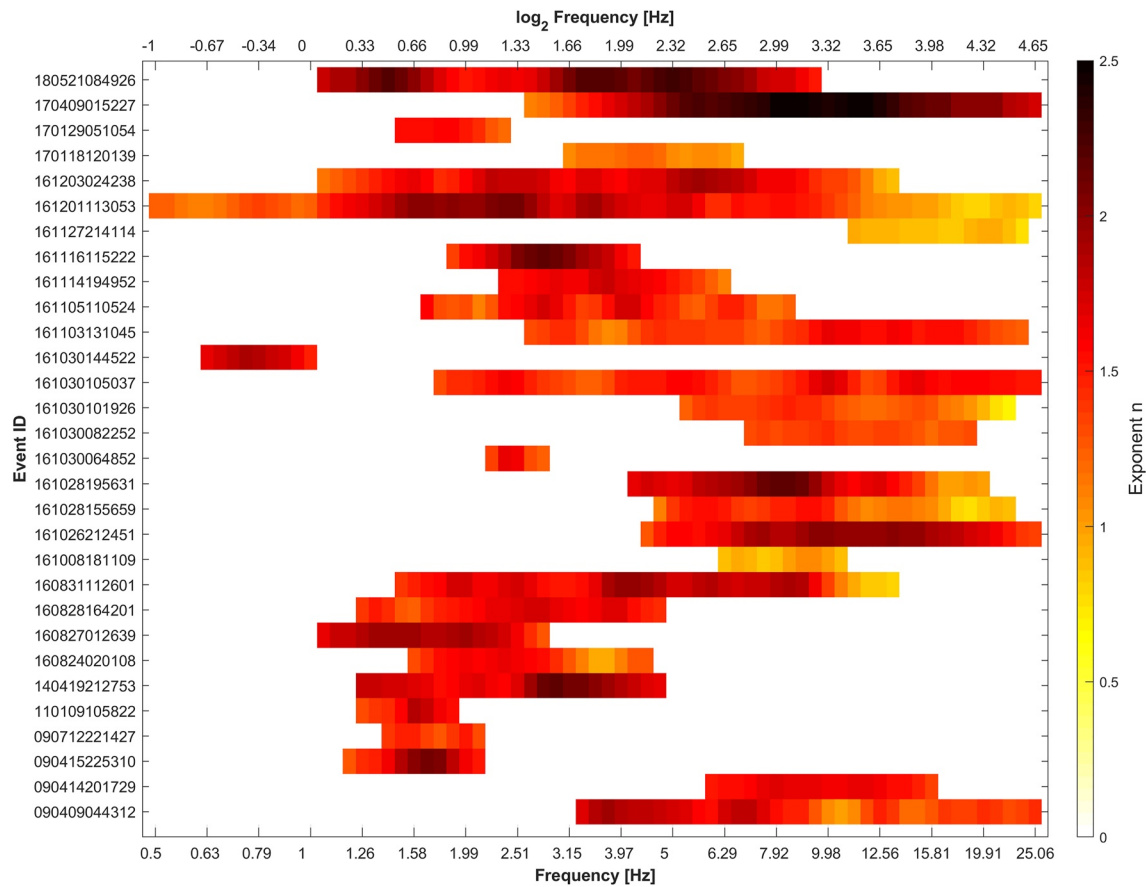


Figure 7. Directivity frequency bandwidth of the events in the magnitude range from 3.7 to 4.0. Each frequency interval of the bandwidth (on x -axis) is colored according to the value of amplitude n . Since the frequency range is logarithmically equispaced, at the top part of x -axis is shown the \log_2 frequency, and the corresponding frequency content is shown at the bottom part of x -axis.

for different magnitude ranges (black lines, all events). At low frequencies, this percentage is less than 6% of the total, then increases to about 20% in the range 5–10 Hz, and finally reduces to 10% at higher frequencies. The same trend is visible for events smaller than 4.1 (green and blue lines in Figure 8b), that represent the majority of the catalog. Nevertheless, as the magnitude increases, these effects start to shift to lower frequencies, although the trends are more scattered due to the reduction of the number of events. Indeed, directivity for moderate events (blue and red curves in Figure 8b) is mainly observed at intermediate frequencies in the band 1.5–10 Hz. For events with magnitude greater than 5 (magenta curve in Figure 8b), more than 30% of the total events exhibit remarkable directivity mainly at lower frequencies ($f \leq 1$ Hz), while this percentage drops starting from 2 Hz.

4.3. Directivity Frequency Band

As clearly observed in Figure 7 and also broadly discussed in other works (Chen et al., 2014; Hirano & Yagi, 2017; Pacor, Gallovič, et al., 2016), the directivity occurs in variable frequency bands that seem to depend on the magnitude of the event. To this purpose, from each directive, we estimate the minimum f_{\min} and the maximum frequencies f_{\max} corresponding to the values where $R^2 > 0.5$ and plot them as a function of magnitude in Figure 9. There, we also report the corner frequency f_c (blue stars) of the events, estimated by Bindi et al. (2018) who applied a generalized spectral inversion technique (GIT) to a similar data set in Central Italy.

In order to highlight the dependencies among the various frequencies and magnitude, we plot in Figure 9 the median values computed on the half-magnitude bins. While the scatter between f_{\min} and f_{\max} is relatively large (red and green symbols in Figure 9), binned median values (red and green solid lines in Figure 9) exhibit a decaying trend with increasing magnitude. The decay rate is roughly similar to that of the corner frequency, suggesting

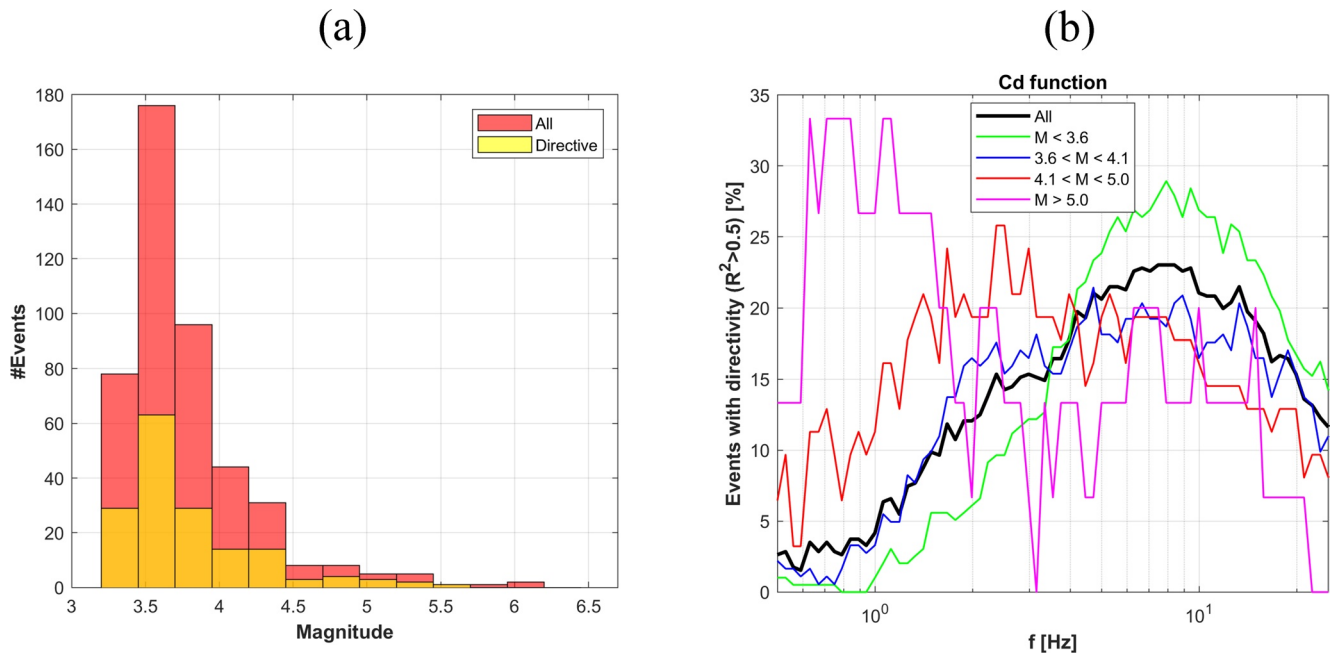


Figure 8. (a) Number of directive events (yellow) compared to the event distributions (red) of the study data set, grouped in magnitude bins. (b) Percentage of the events with $R^2 > 0.5$ out of the total number of events as a function of frequency for different magnitude intervals (see legend).

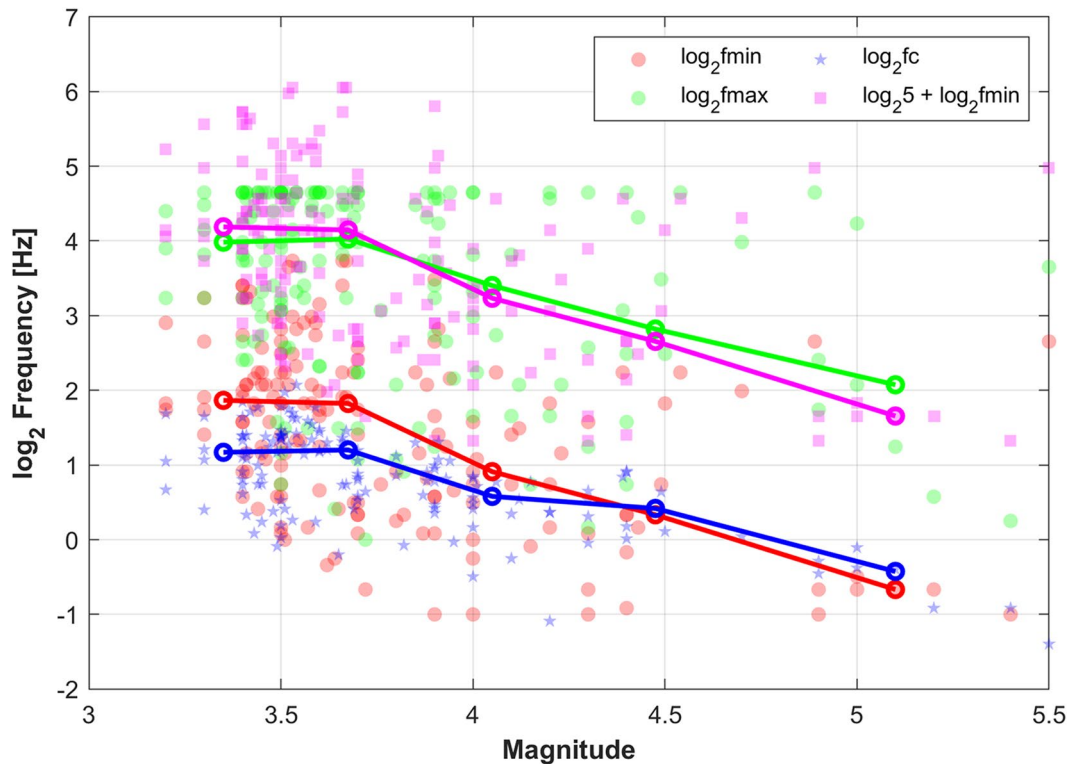


Figure 9. Minimum frequency ($\log_2 f_{min}$, red dots) and maximum frequency ($\log_2 f_{max}$, green dots) where the directivity effect is present plotted as a function of magnitude. Blue stars are the corner frequency from Bindi et al. (2018) and magenta squares are the values obtained by multiplying the minimum frequency by a factor of 5. Solid lines represent the median values over magnitude bin width.

a relation of the directivity bandwidth on the source properties (Trugman et al., 2021). Frequencies f_{\min} and f_c attain approximately the same values for events with magnitude greater or equal than 4 (although they are only a small proportion of the total events), while for smaller magnitudes the relationship between $\log_2 f_{\min}$ and $\log_2 f_c$ seems constant and equal to about 1.6 \log_2 units. We point out that the directivity effect changes the apparent corner frequency as a function of azimuth (Bindi et al., 2020; Chen et al., 2014; Motazedian & Atkinson, 2005; Pacor, Gallovič, et al., 2016). In our analysis this is manifested by increasing directivity amplitude (n) and the coefficient of determination R^2 in the region around the event corner frequency. Nevertheless, in average, f_{\min} and f_c are close to each other.

Directivity bandwidth, defined by the ratio between $\frac{f_{\max}}{f_{\min}}$ in \log_2 scale (i.e., number of octaves), is similar for both small and moderate events. This bandwidth is, on average, about 2.3 octaves (i.e., $\frac{f_{\max}}{f_{\min}} \sim 5$), although this observation may be masked by the limited number of events with $M > 4$ and the maximum frequency of the data set at 25 Hz.

This factor is estimated ad-hoc for the Central Italy data set, but there are also theoretical reasons for having the directivity effect scale-dependent. As pointed out in Pacor, Gallovič, et al. (2016), it seems reasonable that at short scales (approximately 5–7 times shorter wavelengths than the rupture size), the rupture propagation is complex, not exhibiting coherent rupture propagation in a single direction and radiating complex wavefield at higher frequencies. Similarly, the wave propagation due to the complexities of the 3-D velocity model may play a significant role in destroying the directivity effect with the seismic waves coming from various parts of the fault becoming less coherent with increasing frequency, although Pacor, Gallovič, et al. (2016) did not observe significant dependence of the residuals on the source distance. We note that the method and the high-quality data set we use is efficient even at high-frequencies with generally stable at least up to 20 Hz, where the signal-to-noise ratio is still high.

In order to validate the directivity detection of each event, we compare our results with those of Wang et al. (2019), who also analyzes the source rupture directivity in the 2016–2017 Central Italy sequence by an independent technique based on a two-step nonparametric generalized inversion technique [GIT]. Figure S8 in Supporting Information S1 demonstrates a good agreement of inferred directivity frequency bands.

4.4. Directivity Amplitude

As mentioned above, the value of n depends on the type of source model considered, ranging from 0, which represents an incoherent nondirective rupture, to 2, which corresponds to a coherent rupture with the strongest directivity effect. For each event analyzed in this work, we take into account both the peak value n_{\max} and the median value n_{med} , computed within the frequency band where $R^2 > 0.5$.

We note that the uncertainty of the n value varies from frequency to frequency, but is relatively small since the ratio between the standard error (SE) and n is less than 0.5 for 70% of the analyzed directive events for frequency equal to 1 Hz, reaching 97% for frequencies 5 and 10 Hz. A detailed uncertainty analysis is given in Figure S9 in Supporting Information S1 for several characteristic frequency bands.

As we can observe from the histograms of Figure 10, the median values n_{med} span between 0.26 and 2.1 (median 0.69, mean 0.77), while the maximum values n_{\max} vary from 0.8 to 2.8 (median 1.4, mean 1.44). We also note that the largest values of n (above 2) can be alleviated by considering larger rupture velocity (i.e., Mach number) in the fitting procedure. The amplitude of directivity is independent of magnitude, but shows a rather evident linear correlation with the bandwidth of the directivity effects (Figure 11). In particular, the higher is the value of n_{med} , the wider is the frequency band where directivity is observed (see Figure S10 in Supporting Information S1 for an alternative plot with n_{\max}), although the scatter increases.

As shown in Figure 11, the linear fit between n_{med} and the ratio $\frac{f_{\max}}{f_{\min}}$ in logarithmic scale (base 2), is established by the relation:

$$\log_2 \frac{f_{\max}}{f_{\min}} = 2.7427 * n_{\text{med}} - 0.1457 \quad (9)$$

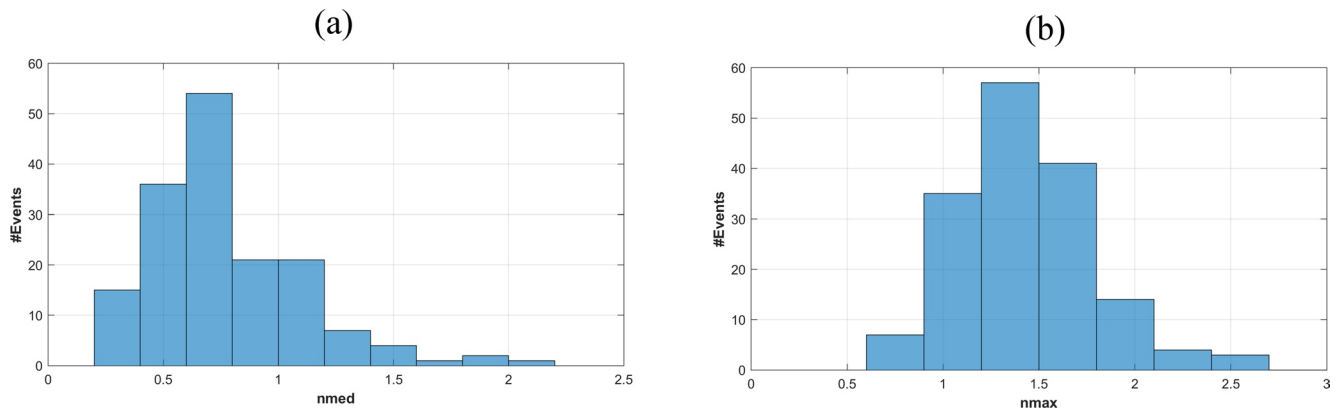


Figure 10. (a) Median and (b) maximum value of the exponent n for the events with directivity among all observed frequency ranges.

The correlation coefficient r , which measures the strength of the linear relationship between n_{med} and bandwidth is 0.7607. In any case, it is important to point out that this relation is valid only for our choice of the fixed parameter values ($k = 0.85$ and $\alpha = 0.5$).

4.5. Direction of Rupture Propagation

In Figure 12, we show the median value of the direction of rupture propagation, θ_0 , over the frequency range where we observe directivity. The distribution is bimodal with two peaks around 150° (SE) and 330° (NW), which are aligned with the fault strikes of the Central Apennines events (D'Amico et al., 2013; Tinti et al., 2016; Improta et al., 2019; Vignaroli et al., 2020).

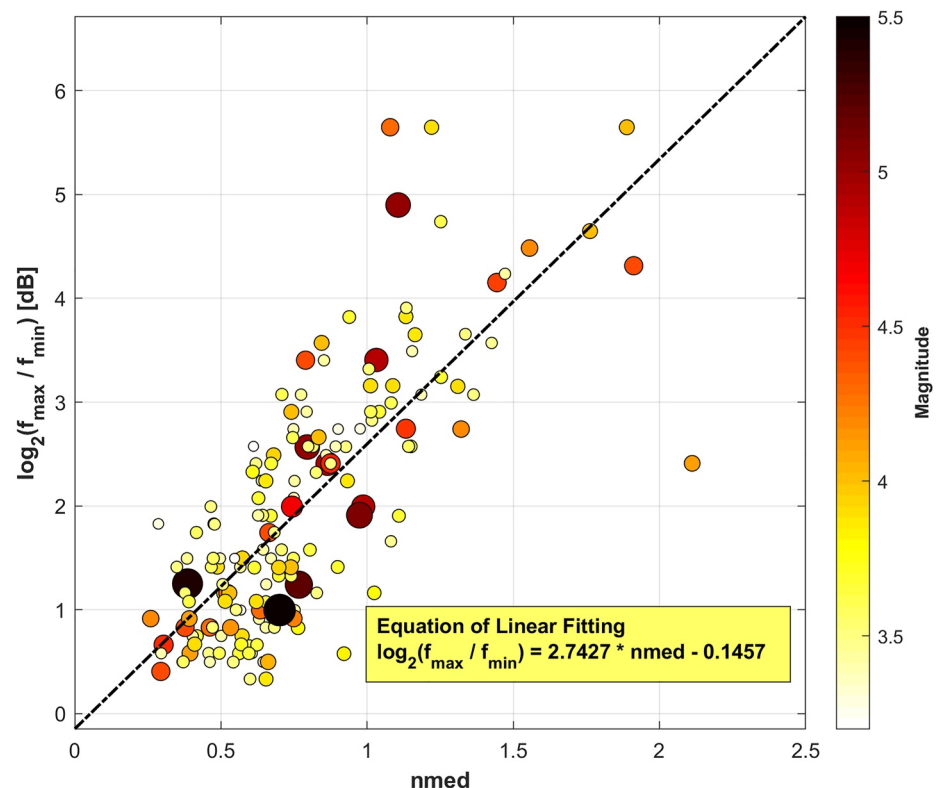


Figure 11. Frequency bandwidth against values n_{med} . The size of the circles is proportional to the magnitude; black dashed line represents the fit curve. The yellow box shows the empirical linear relation between the frequency bandwidth and n_{med} .

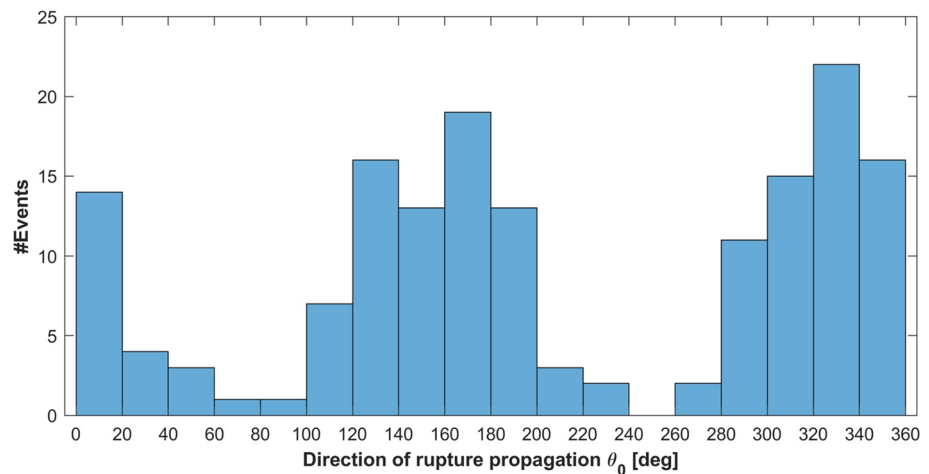


Figure 12. Median value of the direction of rupture propagation for all the directive events.

To support and complement this point, we compare the median of θ_0 with the strike parameters inferred by Herrmann et al. (2011). Figure 13a demonstrates a good match between our rupture and the strikes of the closest plane (fault plane or auxiliary plane): approximately 83% and 45% of the directive analyzed events have a difference less than 45° and 30° , respectively. Nevertheless, some events, mostly with small magnitudes ($M_w < 4.0$), differ significantly from the bisector at 45° , by showing a variation larger than 90° . We point out that the apparent prevalence of rupture directivity in the along-strike direction might be biased due to the ease of detection. Indeed, possible up-dip (or down-dip) rupture propagation along a dipping fault might remain undetected since our analysis is most sensitive to the horizontal rupture propagation.

We also compare our results in terms of direction of rupture propagation with Wang et al. (2019) for 10 most directive events. As shown in Figure 13b, except for 1 event where the values mismatch by approximately 180° , there is a good correspondence between the values of rupture directions obtained by different and independent methods. Similar agreement was found also for the few common events in Pacor, Gallovič, et al., 2016 from the L'Aquila sequence.

We try to find a possible relation between the preferential rupture direction (SE or NW) and seismogenic areas or elements connected to the structural geology. We first consider division into source areas following the cluster classification of Sgobba et al., 2021a (Figure 2), and we do not find any preferred rupture direction. After that, we consider a division based on the recent results from the RETRACE-3D project (Di Bucci et al., 2021, www.retrace3d.it) focused on the revision of all the available geological and geophysical data in the area of the 2016–2018 Central Italy seismic sequence. In particular, we take into consideration one of the major geological structures of the Central Apennines, the Sibillini thrust system (Figure 14a), which consists of a west-dipping, regional ramp-flat structure over thrusting (Lavecchia, 1985; Pierantoni et al., 2013; Porreca et al., 2020) and divides the investigated zone in the hanging and footwall areas to the north and south, respectively. According to Pizzi et al. (2017), the Mt. Sibillini thrust played a key role as a structural barrier at depth, controlling the rupture of the Mw 6.2 Amatrice event of 24 August 2016 (Chiaraluce et al., 2017) and the initiation of the Mw 6.5 Norcia mainshock of 30 October. In particular, the area struck by the Mw 6.5 seismic sequence has a geo-structural architecture delimited to the East by the deep portion of the Miocene-Pliocene Sibillini thrust system, which according to Buttinelli et al. (2021) is the fault that better matches the hypocentral depth and average rupture dip of the Mw 6.5 mainshock.

In general, according to our results, about 60% of the events lie in the cluster at the North to the Mt. Sibillini, while the remaining ones belong to the hanging wall. Histograms in Figure 14b and relative maps in Figures 14c and Figure 14d suggests that for the hanging wall, the events have a fairly equal distribution of unilateral propagation toward North-West and South-East. Regarding the earthquakes in the footwall, we found a preference of the northward direction of the rupture propagation (44 vs. 19 events at the cluster North and South equivalent to about 70% and 30%, respectively).

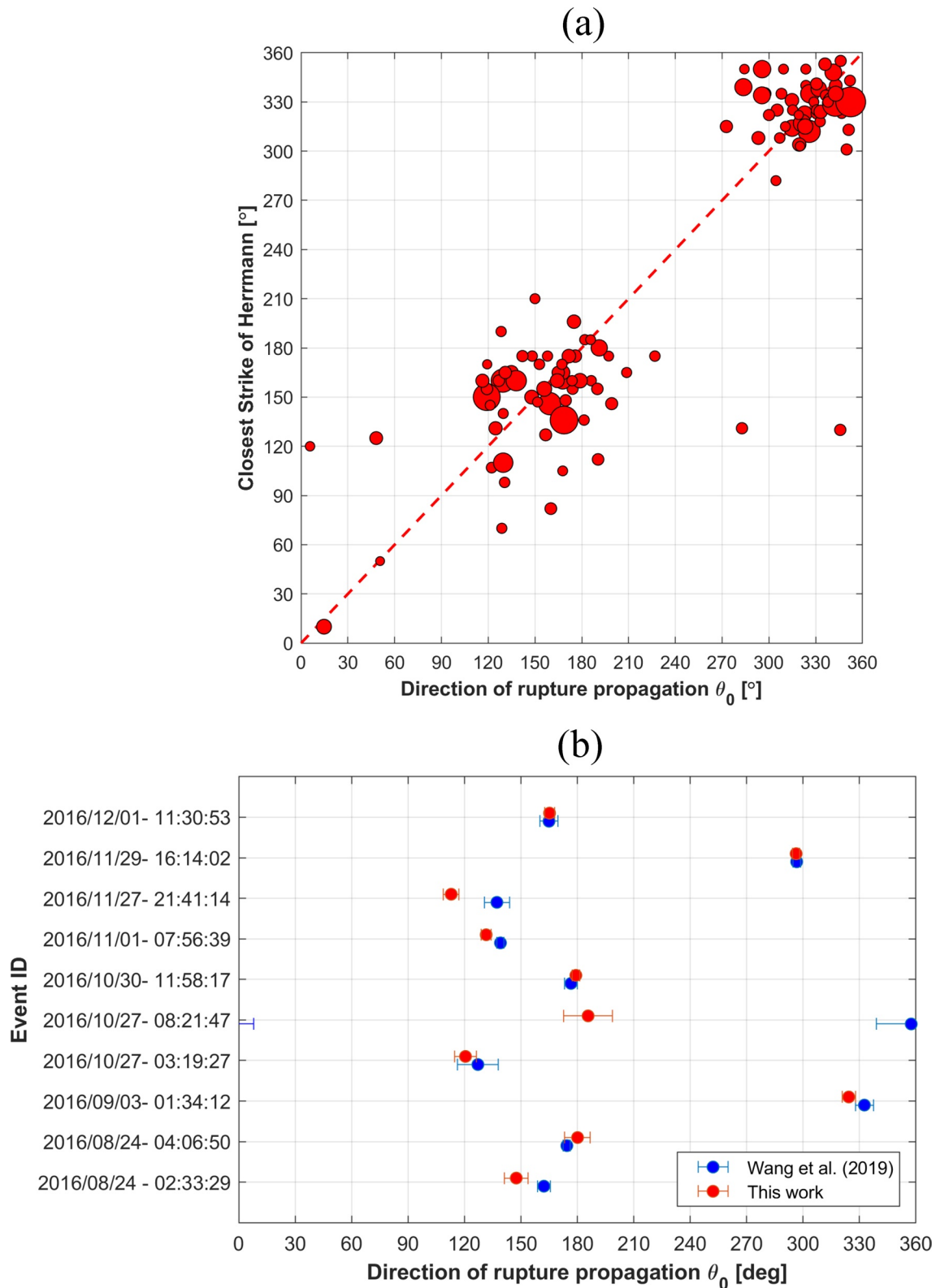


Figure 13. (a) Comparison between the median of the direction of rupture propagation θ_0 and the closest strike inferred by Herrmann et al. (2011) for the analyzed directive events. The size of red dots is proportional to the moment magnitude, while the red dashed line represents the bisector at 45°. (b) Direction of rupture propagation for the 10 most directive events according to the work of Wang et al. (2019). Blue points are the angles θ_0 determined by Wang et al. (2019) and red ones are those estimates in this research. Error bars express the mean ± 1 SD within the frequency band where we observe directivity.

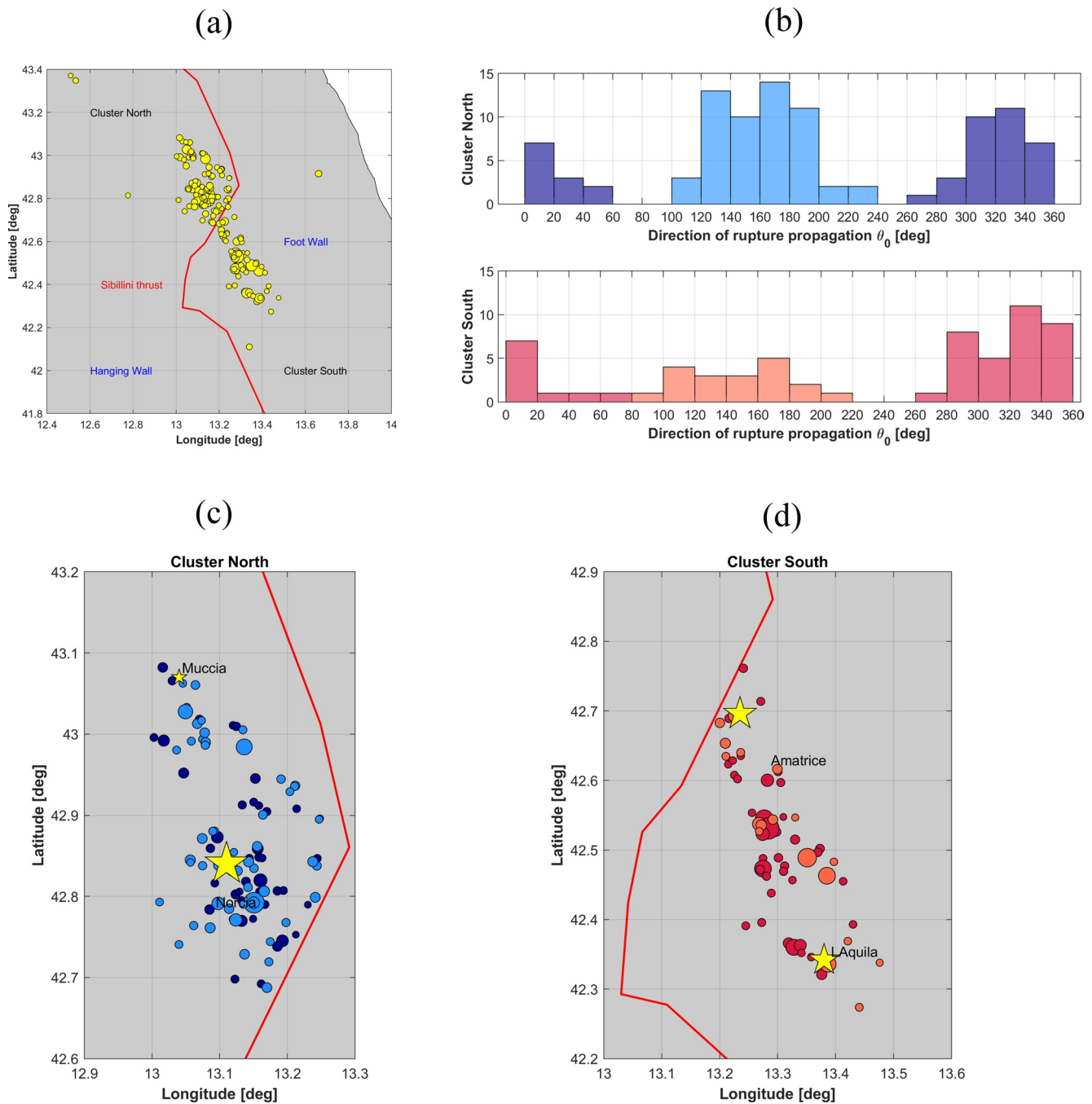


Figure 14. (a) Map of the events (yellow dots, size is proportional to the magnitude) divided in a cluster at the North and at the South with respect to the Mt. Sibillini thrust (red line). (b) Distribution of the median values of the direction of rupture propagation for the two clusters: bluish colors for the northern cluster, reddish colors for the southern cluster. (c) Events with directivity in the cluster North and (d) in the cluster South. For the cluster North, events with direction of rupture propagation toward North (270°–90°) are in light blue while the ones with direction toward South (270°–90°) are in dark blue. For the cluster South, events toward North are in light red and toward South in dark red. Yellow stars are the mainshocks for the analyzed earthquakes.

This preferential northward direction of the footwall, that are evident in particular from the aftershocks of the L'Aquila and Amatrice-Visso-Norcia sequences, is in agreement with the findings of Calderoni et al. (2015) and Calderoni et al. (2017), respectively, who investigated the source directivity from azimuthal variation of records using the spectral ratio technique.

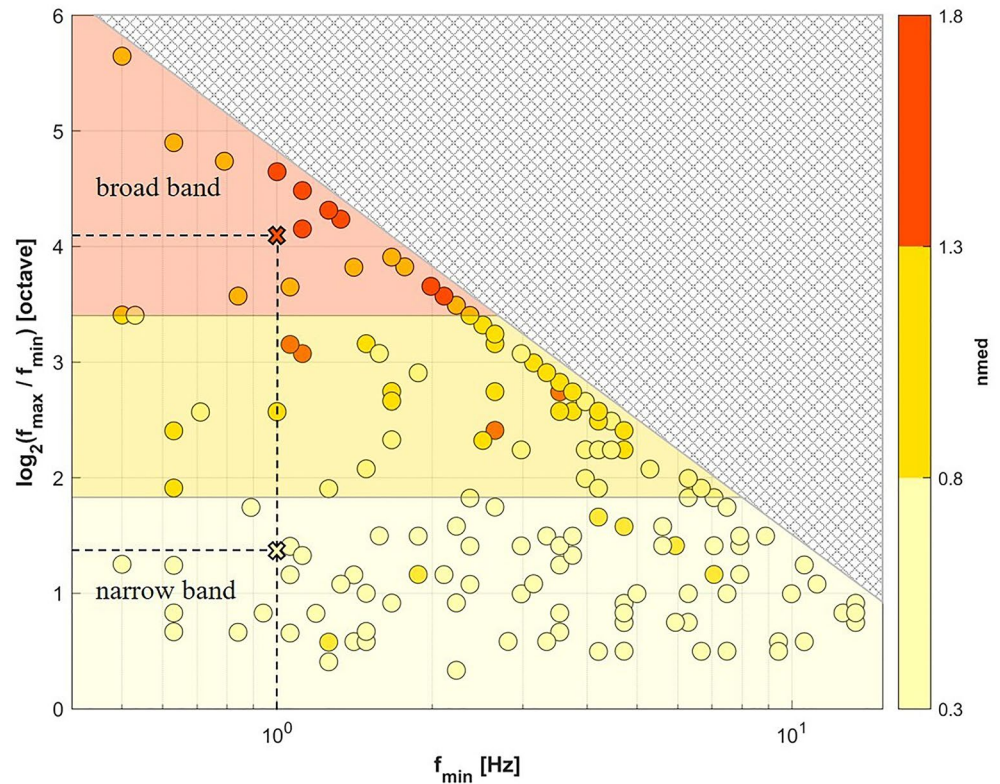


Figure 15. Relation between minimum frequency f_{\min} and bandwidth, defined as $\log_2 \frac{f_{\max}}{f_{\min}}$. The color of the circles represent the value of n_{med} . Three domains are shown in the graph, defined according to the 75th percentile. With crosses, it shows how we can classify an event with Mw 4.6, where $f_{\min} = f_c$, equal at about 1 Hz (see Figure 9). Light-yellow cross shows an event with weak directivity (input value $n_{\text{med}} = 0.55$) where $\log_2 \frac{f_{\max}}{f_{\min}}$ is equal to 1.36 octaves; red cross represent the case with high directivity (input value $n_{\text{med}} = 1.55$) and $\log_2 \frac{f_{\max}}{f_{\min}}$ equal to 4.1 octaves. The large gray triangle with mesh net in the upper right portion of the graph defines the limit of the existence domain of our frequency interval, where $f_{\max} = 25$ Hz.

5. Discussion

In the following, we introduce a first-order classification of the directive events, based on the relations among the frequency bandwidth $\Delta f = \log_2 \frac{f_{\max}}{f_{\min}}$, the amplitude, proportional to n_{med} and the minimum frequency f_{\min} at which the directivity occurs.

To this aim, we plot the distribution of the events on a $f_{\min} - \Delta f$ graph (Figure 15), where each data point is colored following of the values of n_{med} . To group the data points in classes representing the “strength” of the directivity, we select three threshold levels following the n_{med} distribution reported in Figure 10. The weak-directivity class is relative to $n_{\text{med}} < 0.8$ (corresponding to the 65th percentile of the distribution), the moderate-directivity class is in the range between 0.8 and 1.3 (around 65th and 95th percentiles) and high-directivity class is given by $n_{\text{med}} > 1.3$.

One can observe in Figure 15 that these three classes are not randomly distributed, but are localized on specific portions of the graph: the weak directivity events are mainly in the bottom part, the moderate ones in the middle, and the high ones lie on the top. This pattern suggests the definition of three domains, delimited by different values of directivity amplitude and frequency bandwidth. The limiting values of Δf are set to the 75th percentile of the corresponding distributions evaluated for each of the three directivity classes, for which we assumed that the median amplitude n_{med} is uniformly distributed over the frequency range.

As we can observe in Figure 15, the first domain is characterized by weak-directivity and small values of Δf , extended up to 1.8 octaves (i.e., narrow band), the second one is a “transition” zone, up to 3.4 octaves, formed by

the majority of events with moderate directivity and intermediate values of Δf up to 3.4 octaves and a third zone (>3.4 oct), where we mainly find events characterized by high-directivity and larger Δf (broadband).

Such a pattern seems to confirm the existence of a physical correlation between the frequency bandwidth and the strength of directivity so that more directive events tend to show broadband features in contrast to weaker ones that are more bounded in frequency. It is worth to be noted, however, that these findings are strictly related to the statistics of the investigated data set for Central Italy and that we are also limited in exploring the directivity strength of the smallest events due to the inherent limit of observation, which reaches up to a maximum frequency of 25 Hz (the unexplored domain is displayed with a gray triangle with mesh net in Figure 15).

As an example of application, we consider a scenario earthquake with Mw 4.6 in the study area. Following the seismological relation by Morasca et al. (2019) calibrated for Central Italy, this magnitude corresponds to a corner frequency f_c approximately equal to 1 Hz, that can be assumed equal to f_{\min} according to the trend illustrated in Figure 9. Based on the above findings, this f_{\min} may correspond to different frequency bandwidths according to the level of n_{med} (black dashed lines in Figure 15). For instance, an event with weak-directivity ($n_{\text{med}} = 0.55$) is characterized by a directivity bandwidth ranging from f_{\min} equal to 1 Hz up to 3 Hz ($\Delta f = 1.55$ octaves according to Equation 9). On the other hand, for a high-directivity event with $n_{\text{med}} = 1.55$, we get a value of bandwidth $\Delta f = 4.1$ octaves (red cross in Figure 15), which means that in this case directivity appears from 1 to about 18 Hz.

6. Conclusions

In this work, we have shown a statistical overview aiming to provide a better understanding of the main features of the rupture directivity based on empirical evidence of seismological observations. The study exploits the exceptionally dense data set of Central Italy compiled after the main sequences of L'Aquila in 2009 and Amatrice-Visso-Norcia from 2016 to 2018 with a large number of recordings and broad azimuthal coverage.

We take advantage of this data set to calibrate an ad-hoc fully nonergodic FAS-GMM adopted as a reference model for the target area. The nonergodicity enables removal of event-, site- and path-related components of variability in the ground motion model to isolate the aleatory effects related to source directivity. The aleatory residuals exhibit a clear azimuthal pattern and frequency dependence, representing the signature of source directivity, in agreement with previous studies (Pacor, Gallovič, et al., 2016; Luzi et al., 2017; Sgobba et al., 2021a). The residuals are well fitted by the physics-based directivity function $C_d^n(\theta)$, including their asymmetry in amplification of directive and anti-directive sites.

Parameters of the $C_d^n(\theta)$ fitting function are the Mach number α (ratio between the rupture velocity and the shear-wave velocity), the k parameter (the portion of the rupture length in the propagation direction), and the angle θ_0 representing the direction of the rupture propagation. The strength of the azimuthal variability of ground motions is proportional to exponent n , that is a parameter describing the “strength” of the directivity. Optimization tests confirmed the existence of a strong trade-off among the directivity parameters. We thus fixed some of them at typically accepted values in seismological applications: $k = 0.85$, $\alpha = 0.5$. We prefer to allow n to possibly depend on frequency, as it is theoretically more acceptable than frequency dependence of the rupture velocity.

Statistical investigation of directivity parameters for the whole data set demonstrated strong dependence of the exponent n and the coefficient of determination R^2 . On this basis, we classify the events in the data set as “directive” and “non-directive” according to R^2 as a statistical proxy, assuming that the directivity effect is significant when R^2 is greater than 50%. The directive events represent 36% of the whole data set. The main results are:

1. The azimuthal distribution of the directive events is bimodal with two peaks at 150° and 330° , corresponding to the NW-SE structural orientation of the Central Apennines fault systems. Although the prevailing directivity directions are aligned with the strikes of the earthquakes only, events with possibly up-dip or down-dip rupture propagation would remain rather undetected by our method.
2. No preferential direction of directivity is present in the whole data set. Nevertheless, the Amatrice sector is dominated by events with N-NW rupture propagation (70%). We speculate that the main structural discontinuity of the region (i.e., the Mt. Sibillini thrust) may place constraints on the location of the nucleation point, as shown also by Calderoni et al. (2017) for the 2016–2017 seismic sequence of Central Italy. However, the present results need more extensive correlations with rheological parameters to be confirmed.

- The observed directivity is a band-limited phenomenon. It occurs always above the corner frequency up to a maximum value f_{\max} . We propose a relation between the directivity amplitude n and the directivity bandwidth (Equation 9). For example, for a moderate directivity ($n = 0.9$), the bandwidth spans up to five times the minimum frequency f_{\min} . We note that the result is partially only an extrapolation for smaller events, because we are limited by the maximum investigated frequency of 25 Hz.

We point out that our observed band-limited directivity effect cannot be easily explained by a simple azimuthally dependent stretching of the apparent source times functions identified by time-domain methods (e.g., Abercrombie et al., 2017; Warren & Silver, 2006; Yin et al., 2021). Indeed, when we move from the time to the frequency domain, the apparent duration stretching causes contraction of the whole spectrum, including a shift of the corner frequency f_c . The directivity would thus affect the spectral amplitudes at all frequencies beyond the minimum apparent f_c . It means that the observed limited directivity bandwidth should be due to other phenomenon, possibly related to small-scale complexity of the rupture evolution.

Our quantification of the frequency-dependent directivity effect on the ground motions represents a first step toward its empirical parameterization, since it connects the bandwidth of significant directivity with the fundamental scenario parameters. In particular, we can use the directivity function $C_d^n(\theta)$ for predictive purposes, by setting θ_0 and n and calculating the directivity bandwidth from Equation 9. The minimum directivity frequency can be approximated by the corner frequency calculated from the stress drop and magnitude using seismological scaling laws.

This effort to incorporate the directivity contributions into empirical GMMs was carried out because past investigations performed in the NGA-West2 project (Spudich & Chiou, 2008; Spudich et al., 2014) led to the conclusion that “none of these models has a functional form that transitions smoothly from large to small magnitude and that describes the small-magnitude or short-period directivity.” Our results may provide useful insights to move toward parametrization of directivity within the empirical simulation of shaking scenarios in Central Italy, as the ones proposed by Sgobba et al. (2021) in the nonergodic framework. Nevertheless, we clarify that the proposed approach reproduces well the characteristics of the data in the target region while more extensive research is needed to generalize these findings.

Conflict of Interest

The authors declare no conflicts of interest relevant to this study.

Data Availability Statement

All the data used for Central Italy (CI_dataset) can be found at the page: <https://shake.mi.ingv.it/central-italy/>. In particular, for this analysis, coefficients and uncertainty terms of the nonergodic FAS ground motion model CI-FAS_GMM6 (Lanzano et al., 2022, https://doi.org/10.13127/ci_dataset/CI-FAS_GMM) and the directivity characteristics of the 162 events CI-FAS_directivity (Colavitti et al., 2022; https://doi.org/10.13127/ci_dataset/CI-FAS_directivity) are publicly available. The data on the strikes of fault and auxiliary plane, with which we compared the median values obtained in this research are available at https://www.eas.slu.edu/eqc/eqc_mt/MECH.IT.

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Acknowledgments

The authors are very grateful to Prof. Daniele Spallarossa (University of Genova, Italy) for making the data set used in this work available. This research is supported by Istituto Nazionale di Geofisica e Vulcanologia (INGV) in the frame of the project *Pianeta Dinamico* (Working Earth) and INGV-DPC Agreement B1 2019-2021—Geosciences for the Understanding of the Dynamics of the Earth and the Consequent Natural Risks (CUP code D53J19000170001), founded by the Italian Ministry of University and Research (MIUR) in the Task S3–2021—Seismic attenuation and variability of seismic motion. Moreover, this study has been partially funded by INGV-DPC Agreement B1 2019–2021, with the goal of promoting research activities in the field of seismic hazard in Italy. The authors are very grateful to the Editor Prof. Rachel E. Abercrombie, an anonymous reviewer and Prof. Yefei Ren for their valuable comments that improved this manuscript. Open Access Funding provided by Istituto Nazionale di Geofisica e Vulcanologia within the CRUI-CARE Agreement.

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