Macroseismic intensity attenuation models calibrated in Mw for Italy

Bulletin of Earthquake Engineering

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a)



b)



c)



d)



e)



f)



Fig. A1 Spatial distribution of IDPs (left panels) and macroseismic intensity versus epicentral distance obtained from instrumental epicentre (right panels) of seven events with *Mw* greater than 6.0: a) 1908.12.28 *Stretto di Messina* (*Mw* 7.1); b) 1915.01.13 *Marsica* (*Mw* 7.0); c) 1980.11.23 *Irpinia-Basilicata* (*Mw* 6.8); d) 1930.07.23 *Irpinia* (*Mw* 6.6); e) 1920.09.07 *Garfagnana* (*Mw* 6.5); f) 1976.05.06 *Friuli* (*Mw* 6.4); g) 1962.08.21 *Irpinia* (*Mw* 6.2)



Fig. A2 Comparison of mean residuals and standard deviation (error bars) versus predicted macroseismic intensities for all macroseismic intensity attenuation model



Fig. A3 Comparison of mean residuals and standard deviation (error bars) versus epicentral distance for all macroseismic intensity attenuation models



Fig. A4 Comparison of mean residuals and standard deviation (error bars) versus magnitude Mw for all macroseismic intensity attenuation models



Fig. A5 Between event residuals versus Time and Mw of the calibration earthquakes

In CPTI15, various methods were applied to determine the magnitudes (Mw) of smaller earthquakes. This study utilised 119 events (electronic supplement Tab. A1), and the methods and time periods for Mw determination are as follows:

i) 31 earthquakes with Mw ranging from 3.8 to 4.8 between 1983 and 2005. The magnitude is calculated as the mean of Mw values determined from Ml, Ms, and mb estimates, with weights inversely proportional to their associated variances;

ii) 14 earthquakes with *Mw* in the range of 3.9 to 4.4 between 1983 and 1998. *Mw* is directly converted from local magnitude;

iii) 24 earthquakes with Mw ranging from 4.4 to 7.0. The magnitude is calculated as the mean of Mw values converted from Ms and mb estimates, with weights inversely proportional to their variances;

iv) 2 events with *Mw* ranging from 4.5 to 4.8. *Mw* is the mean of values converted from *Ms*, *mb*, and *Ml* Wood-Anderson estimates, with weights inversely proportional to their variances;

v) 2 events (1977-1980) with Mw ranging from 4.26 to 4.34. Mw is determined from Ml Wood-Anderson;

vi) The 1908 Messina earthquake has its Mw determined from S-waves amplitude;

vii) For 45 earthquakes with Mw ranging from 3.9 to 6.8, Mw is obtained from moment tensor solutions;

The use of diverse methods for calculating Mw may contribute to the observed bias in the between-event residuals.



Fig. 6b

Fig. A6a The map is plotted using 4 classes of Vs30 (m/s), following the EC8 classification (CEN, 2004); n=number of IDPs. **Fig. A6b** Distribution of residuals (coloured dots) versus Vs30 considering 90 m/s bins, its mean (square dots) and standard deviation (error bars), for the Log-Lin_10 model

Table A2 Coefficients and standard errors of the macroseismic intensity attenuation models analysed in the present study. Type 1 models (Eq. 4) are denoted by N=1, 2, 3, 7, 8 and type 2 models (Eq. 5) by N=4, 5, 6, 9. All models are fitted to 16,260 IDPs, except Log-Lin_{cut dist I=3} (N=7) which is fitted to 12,587 IDPs with $I \ge 3$. Column *Comp h* specifies if the pseudo focal depth *h* is a fixed (fix) or estimated (fit) parameter. The standard deviation of the intensity (σ_I), given by the root mean square of intensity residuals (Eq. 11), is provided for all models. For all type 2 models, the last column reports the standard deviation (σ_LogI) given by the root mean square of the log-intensity residuals for the additive model (5), as well as the standard deviation ($\sigma_Lobs/Icomp$) of the intensity residuals given by the theoretical formula (A1.4) for the equivalent multiplicative model (6)

N	Model [Type]	a Independent Term (error) [% error]	<i>b</i> Geometric Spreading (error) [% error]	<i>c</i> Anelastic Attenuation (error) [% error]	<i>d</i> Magnitude Coeff. (error) [% error]	h (km) (error) [% error]	Comp h	RMS_I	σ_LogI [σ_Iobs/Icomp]
1	Log-Lin 5	1,11	2,14	0,0054	1,41	5	fix	0,749	
	[Type 1]	(0,05) [4,5]	(0,03) [1,4]	(0,0002) [3,7]	(0,007) [0,5]			,	
2	Log-Lin_10	1,81	2,61	0,0039	1,42	9,87	fit	0,748	
	[Type 1]	(0,10) [5,5]	(0,07) [2,7]	(0,0003) [7,7]	(0,007) [0,5]	(0,56) [5,7]			
3	Log-Lin_16	2,86	3,26	0,0020	1,43	16	fix	0,754	
	[Type 1]	(0,06) [2,1]	(0,04) [1,2]	(0,0002) [10,0]	(0,007) [0,5]				
4	CRV5	-0,006	0,17	0,0004	1,35	5	fix	0,735	0,0657
	[Type 2]	0,007 [117]	(0,003) [1,8]	0,00002 [5]	(0,008) [0,6]				[0,13]
5	CRV9	0,032	0,19	0.0003	1,36	8,72	fit	0,731	0,0655
	[Type 2]	(0,01) [31,2]	(0,006) [3,2]	(0,00003) [10]	(0,008) [0,6]	(0,69) [7,9]			[0.13]
6	CRV ₁₆	0,125	0,25	0,0002	1,37	16	fix	0,738	0,0657 [0,15]
	[Type 2]	(0,008) [6,4]	(0,004) [1,6]	(0,00002) [10]	(0,008) [0,6]				[0,13]
7	Log-Lin _{cut dist I=3}	2,12	2,84	0,0051	1,45	11,3	fit	*0,771	
	[Type 1]	(0,18) [8,5]	(0,12) [4,2]	(0,0005) [10]	(0,010) [0,7]	(0,80) [7,1]			
8	Log Model	3,39	3,63	-	1,42	16,6	fit	0,751	
	[Type 1]	(0,069) [2]	(0,036) [1,0]		(0,007) [0,5]	(0,45) [2,7]			
9	Log Model [Type 2]	0,171 (0,008) [4.7]	0,29 (0,004) [1,4]	-	1,36 (0,008) [0,6]	16,2 (0,57) [3,5]	fit	[0,735]	0,0659 [0,16]

Appendix 1

A1. Some notes on the multiplicative model (6) and its logarithmic transformation (5)

Eq. (6) is a multiplicative model such as:

$$I = f(R_{epi}, Mw) \cdot \varepsilon' \tag{A1.1}$$

where the intensity *I* is positive and depends on moment magnitude *Mw* and epicentral intensity R_{epi} , and ε' is a random error.

To exploit computational methods for linear models, a double-log transformation is applied to model (6) and the linear model (5) is obtained:

$$Log I = Log f(R_{epi}, Mw) + \varepsilon$$
(A1.2)

where Log I spans the real line (Log denotes the base 10 logarithm) and ε is reasonably assumed to be an additive Gaussian/Normal error with zero mean and standard deviation σ , namely $\varepsilon \sim Normal(0, \sigma)$ and $Log I \sim Normal(Log f(R_{epi}, Mw), \sigma)$. The advantage of model (A1.2) is that it is linear and its coefficients can be estimated by standard computational methods developed under the Gaussian error hypothesis. Once model (A1.2) has been estimated, empirical probability rules for Normal distribution can be used to quantify the uncertainty on the predicted value (I_{comp}) of the intensity. In this study we consider $Prob(Log I_{comp} - \sigma \le Log I \le Log I_{comp} + \sigma) \approx 0.683$, that is the probability of the true value of Log I falling within the uncertainty range [$Log I_{comp} - \sigma, Log I_{comp} + \sigma$] is roughly 68.3%.

The model (A1.1) is easily obtainable from the estimated model (A1.2), after making some considerations on its error ε' and on its residuals. Since ε is assumed to be Normally distributed and $\varepsilon' = 10^{\varepsilon} = exp$ ($\varepsilon \cdot ln10$), it turns out that ε' is LogNormally distributed by definition. Moreover, mean and standard deviation of ε' can be obtained from those of ε as follows:

mean of
$$\varepsilon' = e^{(\sigma')^2/2}$$
 (A1.3)

standard deviation of
$$\varepsilon' = \sqrt{e^{(\sigma')^2}(e^{(\sigma')^2} - 1)}$$
 (A1.4)

where $\sigma' = \sigma \cdot ln 10$.

Consequently, the uncertainty on the intensity predicted from the multiplicative model (A1.1) has the following expression: $Prob(I_{comp} \cdot 10^{-\sigma} \le I \le I_{comp} \cdot 10^{\sigma}) \approx 0.683$.

Finally, the residuals in model (A1.2) are defined as the difference $Log I_{obs} - Log I_{comp}$ between the logarithm of the observed and predicted values of the intensity *I*; these residuals vary around 0, a value which indicates the exact match between observed and predicted intensity. As a consequence, the residuals in the multiplicative model (A1.1) are positive ratios $\frac{I_{obs}}{I_{comp}} = 10^{Log I_{obs}-Log I_{comp}}$, that vary around 1. In this case, a residual of 1 would imply that the observed intensity exactly matches the predicted one.