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A meta-analysis of fault slip rates across the central Apennines

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Key Points:

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- We meta-analyzed throw measurements for active normal faults in the central Apennines;

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- Throw rates are overdispersed with respect to age and offset uncertainties; therefore, larger uncertainties are implied;

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- A robust null hypothesis is that most faults have throw rates that are quasi-constant through geologic time, and also along their traces, except for tapering at unconnected ends.

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14

15 [Abstract](#)

16 Several methods such as paleoseismic trenching, mapping of offset geomorphic markers, and dating of
17 scarp profiles have been used to determine slip rates of normal faults in the central Apennines.
18 Combining measurements obtained with different methods remains challenging because non-tectonic
19 processes can introduce noise or spurious signals that are elusive to quantify, and these influence slip
20 rate estimates. To this end, we meta-analyzed throw measurements with associated ages collected in
21 the central Apennines with several methods to quantify such erratic fluctuations and method-related
22 variances.

23 We show that throw rates are overdispersed with respect to nominal uncertainties in throw and age;
24 therefore, they are commonly affected by unmodeled noise processes. After comparing throw rate
25 distributions sampling the same faults with different techniques, no clear spatiotemporal patterns
26 appear, but only quasi-random noise. Assuming that field investigators sampled real tectonic features
27 (i.e., fault scarps), we find that such erratic throw rates indicate total uncertainties are two to three
28 times greater than the stated observation uncertainties. In this situation, a simple and robust null
29 hypothesis is appropriate. We propose that most faults should be assumed to have uniform throw rate
30 along their traces, except for possible tapering near unconnected ends. We also propose that models in
31 which throw rates are time-dependent (within the last 25 ka) are not yet justified. Then, relying on the
32 estimated total uncertainties, we determine the most probable long-term fault throw rate for each
33 active fault by combining different throw-rate probability density functions.

34 35 [Plain Language Summary](#)

36 Recent advances in paleoseismology are expected to provide crucial components for seismic hazard
37 models in the years to come. Different paleoseismological methods sample complementary features
38 along active fault traces. Thus, the most significant progress in seismic hazard analyses using
39 paleoseismological data will stem from advanced statistical methods comparing (and connecting)
40 outcomes of different techniques. However, comparing alternative techniques remains problematic,
41 given the limited number of feature measurements across the same active fault.

42 The available data for active faults across the central Apennines, collected with diverse methods by
43 different research groups, represents an opportunity to solve this issue due to their dense
44 concentration. Furthermore, such data richness motivated us to carefully investigate the spatial and
45 temporal variabilities to discriminate tectonic signals from noise.

46 We meta-analyzed available throw measurements (and their ages), inspected for disagreements among
47 features, and found evidence for significant data overdispersion. Hence, we argue that better
48 quantification of the total uncertainties is needed and that uniform throw rate (along each trace and

49 through the last 25 ka of geologic time) is still the most reasonable model for many scientific and urban
50 planning purposes.

51

52 1 Introduction

53 Recent seismic hazard models are increasingly relying on fault slip rates as the fundamental quantity for
54 translating the activity of a fault source model into earthquake rates (Field et al., 2014; Kastelic et al.,
55 2016; Murru et al., 2020; Visini et al., 2021). In this conversion, modelers are often tasked with selecting
56 different estimates and alternative methods to assess the fault slip rates and related uncertainties and
57 incorporate them into the seismic hazard analysis.

58 As for the central Apennines, most fault estimates are restricted to the past 25 kyr. Considering that
59 stable estimates can be retrieved after 5-10 earthquakes causing slip at the some point on some fault
60 trace, this time interval is barely sufficient to assess reliable long-term slip rates because the recurrence
61 interval for strong earthquakes on each fault is 1000-3000 yr (Galli, 2020).

62 A further complicating factor in determining long-term estimates is earthquake clustering. Episodic
63 behavior has been suggested for the central Apennines, alternating thousands of years of rapidly
64 accumulating displacement and similar periods of almost complete inactivity (Cowie et al., 2017).
65 However, such clustering models are inconsistent with periodic and quasi-stationary slip rates reported
66 for Mount Morrone (Gori et al., 2011), the Middle Aterno Valley fault (Falcucci et al., 2015), and other
67 faults of central Apennines (Galadini & Galli, 2000).

68 The findings of different studies on the along-strike variation of throw rates in the central Apennines are
69 also controversial and potentially misleading. There is increasing literature suggesting along-strike
70 variation and even proposals to use specific model throw-rate profiles for exploratory seismic hazard
71 assessments (Faure Walker et al., 2019). These data collections usually mimic the theoretical model of
72 the along-strike throw distribution of a growing isolated (unconnected) fault in a uniform medium
73 (Cowie & Scholz, 1992; Peacock & Sanderson, 1991) with variable throw rates. This model mainly relies
74 on detailed field mapping of active faults and related slip rates. However, geologic mapping can report
75 incomplete fault traces due to vegetation, grazing animals, plowing, urban development, soil creep,
76 landsliding, or other non-tectonic processes that hide parts of the fault traces. Also, any strike-slip faults
77 that may connect to the mapped ends of normal faults may remain undetected because they produce
78 no scarps.

79 Furthermore, it has been proven that climate change, topographic gradients, and the vertical and lateral
80 variation of geotechnical properties play crucial roles in unevenly amplifying, or possibly reducing, the
81 throw along fault traces. These sources of noise are active both in interseismic (Kastelic et al., 2017;

82 Stemberk et al., 2019) and coseismic phases, both in the Apennines (Di Naccio et al., 2019) and
83 elsewhere (Chen et al., 2012; Roback et al., 2018). As a result, the measured slip rate fluctuations can
84 reflect the influence of different noise sources reshaping the fault scarps more than a real (tectonic)
85 variation of long-term slip rate along the trace. This implies that the observational errors in throw
86 measurement at each single site might not capture the lateral overdispersion of throw measurements.
87 Thus, short-wavelength non-tectonic variations require an epistemic uncertainty in slip rate estimates
88 about three times bigger than the observational error (McPhillips & Scharer, 2018).

89 Several "static" models of seismogenic sources (Basili et al., 2008; Boncio et al., 2004) have been
90 proposed as an alternative to the growing-fault models because they relax the assumption of fitting, and
91 locally overshooting, the throw rates along the fault strike. Despite significant procedural differences,
92 these static models define a main fault ("seismogenic master fault" Boncio et al., 2004) or "composite
93 seismogenic source" (Basili et al., 2008) with a constant slip rate in space and time, but tapering to zero
94 at any unconnected fault tips. These major structures are considered continuous at depth, although
95 they might split into closely spaced fault segments at the surface.

96 Finally, different methods have been applied in the central Apennines to quantify fault slip rates.
97 However, alternative methods sample complementary features, and examining fault activity with
98 alternative techniques is expected to sample different spurious signals if non-tectonic processes are not
99 marginal. Unveiling this latter point is fundamental for any modeler. Similarly, combining rates acquired
100 with different methods remains the most challenging as an agreed method of synthesizing slip rates is
101 still missing (Morell et al., 2020). Therefore, the central Apennines represent an optimal case study due
102 to the temporal and spatial proximity of the relative measurement points of different methods. With
103 this goal, we meta-analyzed available throw measurements (and their ages), inspected for
104 disagreements among features, and found evidence for significant overdispersion among different
105 techniques of throw measurement. Hence, additional epistemic uncertainties are implied, and they may
106 be larger than observational uncertainties by a factor varying up to two/three for scarp profiling.

107 We warn against overfitting of throw rates, which invokes unsupported spatial or temporal fluctuations
108 in the last 25 kyr, and argue that a uniform throw rate (tapering to zero at any unconnected fault tips) is
109 still the most reasonable model for many scientific and urban-planning purposes.

110

111 [2 Study area and geological background](#)

112 Our study area extends for more than 150 km along the central Apennines (Figure 1). This mountain
113 range was originally a northeast-verging fold and thrust belt, but now the axial portion is undergoing
114 regional uplift and NE-SW active extension since Late Neogene-Early Quaternary (Carafa & Barba, 2013;
115 Cavinato & Celles, 1999). Extension at 2.5-3.5 mm/a (Carafa & Bird, 2016; Carafa et al., 2020; Devoti et

116 al., 2017) shaped a complex landscape featuring previous thrust-belt ranges, glacial and fluvial
117 landforms, karst plateaus, and fluviolacustrine basins. In the central Apennines, the NW-SE and NNW-
118 SSE striking normal faults dissect and overprint older compressional features, drive the deposition in
119 intramountain basins (e.g. Fucino and Sulmona basins), and control the Late Pleistocene (< 126 ka)-
120 Holocene continental deposits and landforms. The average fault trace length is greater than 10 km, and
121 these are usually interpreted as ground-breaking tectonic faults, although investigations along the strike
122 and in-depth are scarce, and the role of exogenic processes perturbing the tectonic signal remains
123 underexplored.

124 In many cases, the faults separate footwall carbonates from unconsolidated deposits such as slope
125 debris and colluvium eluvium or alluvial fan deposits in the hanging wall (Blumetti et al., 1993; Galadini
126 & Galli, 2000), and their trace is lost where the relief ends. In this setting, bedrock fault scarps are locally
127 well-exposed and mapped at different elevations along the steep carbonate bedrock of the footwall.
128 These latter features have been documented to favor slope instabilities mimicking a tectonic signal
129 (Galadini, 2006; Kastelic et al., 2017). In other cases, fault planes are not visible at the ground surface,
130 but they leave dislocated landforms and stratigraphic markers or earthquake horizons (Galli et al., 2008).
131 Evidence of fault activity is supported by displaced Pleistocene and Holocene alluvial fans, slope deposits
132 (Galli et al., 2010; Gori et al., 2007; Miccadei et al., 2002; Villani et al., 2015), and smooth upland
133 surfaces mainly created around the time of the Last Glacial Maximum (Giraudi, 1995; Pantosti et al.,
134 1996; Papanikolaou et al., 2005; Roberts & Michetti, 2004).

135 [3 Review of applied methods to determine paleoearthquake throw and cumulative throw rate in](#) 136 [the central Apennines](#)

137 We collected the cumulative throw rates from a thorough literature review on geologic/geomorphic
138 data along 59 normal fault traces in the central Apennines, also considering alternative versions of data
139 contested by competing research teams.

140 In an extensional setting like the central Apennines, offsets are usually constrained by vertical
141 separation of displaced stratigraphic layers or geomorphic features. For this reason, all of our performed
142 calculations and results have been focused on the relative vertical offset (throw). We calculated the
143 throw from slip (when necessary) using the reported dip angle in the few cases analyzing the offset
144 along the fault plane (see Benedetti et al., 2013, as an example). The most common geomorphic
145 markers used as piercing points to calculate fault offsets are fluvial terraces, alluvial-glacial fan surfaces,
146 moraine crests, valleys, and remnants of peneplains. The offset age is bracketed by several Quaternary
147 dating methods, such as radiocarbon and cosmogenic radionuclide. In many cases, the dislocation
148 events are constrained by correlation with regional subplanar surfaces or stratigraphic layers of known
149 age (Blumetti & Guerrieri, 2007; Galadini, 1999; Giraudi & Frezzotti, 1997). For different faults (e.g.

150 Aremogna-Cinque Miglia, Sangro-Mt. Marsicano, Barisciano-San Pio delle Camere, Campo Felice, Campo
151 Imperatore, Assergi), we opted to analyze the throw rates of each splay and not merge them in a single
152 polyline to highlight eventual spatial variations.

153 We organized the collected data into two different datasets. The first one (Table S1) contains 403 throw
154 features older than 2 ka as younger rates can be affected by non-marginal elastic-strain changes (see
155 section 4.3 “Definition and determination of the net elastic offset ΔE ”). This dataset is meta-analyzed to
156 determine the uniform throw rate along 59 normal faults. The second dataset (Table S2) contains all the
157 dislocations resulting from single paleoearthquake throws (150 measurements) regardless of the age of
158 the event as this allowed us to determine the distribution of coseismic throws along a set of 18
159 representative normal faults. This dataset is used to estimate the coseismic slip distribution of all active
160 faults in the central Apennines (see section 4.3 “Definition and estimation of the net elastic throw ΔE ”).

161 We divided the cumulative throw rate measurements into four classes according to the used
162 methodology.

163 3.1 Paleoseismic trenching (TRCH)

164 In the last three decades, paleoseismic trenching (from now on, the TRCH technique) identified several
165 previously unknown paleo-earthquakes in the central Apennines (Boncio et al., 2010; Cinti et al., 2011;
166 D'Addezio et al., 2001; Galadini & Galli, 1999; Michetti et al., 1996; Moro et al., 2013; Salvi et al., 2003;
167 Saroli et al., 2008; Serva et al., 1986). Throw rates were measured for different studied faults by digging
168 trenches in Quaternary continental deposits, mapping their walls, and determining earthquake
169 displacements. In detail, paleoseismological trenches are focused on measuring earthquake horizons,
170 defined as coseismic throws or deformed landforms and sediments. The dislocation measurements
171 often represent minimum throw estimates due to previous erosional events altering the preservation of
172 stratigraphic layers in the footwall. In other cases, layers are not recognizable within the hangingwall or
173 are buried below it. Throw age is often bracketed between the age of the pre- and post-earthquake
174 layers and has been determined by dating collected samples (e.g., shells, charcoal, paleosoils) or tephra
175 layers by radiocarbon, thermoluminescence, and archeological methods. The total number of TRCH
176 measurements is 46 in the throw rate dataset (Table S1) and 81 in the coseismic throw dataset (Table
177 S2).

178 3.2 Measuring geomorphic markers post Last Glacial Maximum (LGMA method)

179 The Last Glacial Maximum (LGMA) extensively shaped the landscape resulting in erosive landforms (e.g.,
180 cirques or glacial valleys) and deposits (till, moraine, or glacial fans) on a regional scale (Giraudi, 1988,
181 1989, 1995). The last age at which these climate-induced features dominated over tectonic scarp-
182 formation constrains the later cross-cutting activity of faults in the central Apennines. Last Glacial
183 Maximum data refer to dislocated subplanar or sublinear geological and geomorphic markers exposed

184 since this climate change (Giraudi & Frezzotti, 1995; Piccardi et al., 1999; Roberts & Michetti, 2004). This
185 data subset often lacks direct dating of the geomorphic marker because it assumes the observed offset
186 has occurred since the demise of the Last Glacial Maximum (15 ± 3 kyr, Faure Walker et al., 2010).
187 Furthermore, investigators often believe that the sampled parts of the footwall and hangingwall slopes
188 are intact because recent exogenic processes on faults scarps are not expected to perturb the tectonic
189 signal (Tucker et al., 2011). Under this assumption, the scarp height variations along strike are purely
190 tectonic features because they represent cumulative throws. With 232 data points, the LGMA is the
191 most common sampling technique in our data collection. For many data, the offset is routinely
192 estimated by the intersection of the fault scarp topography with the surrounding surface morphology
193 with meter ruler accuracy, which define the observational error of the measurement (Bubeck et al.,
194 2015; Cinti et al., 1992; Faure Walker et al., 2009; Faure Walker et al., 2010; Morewood & Roberts,
195 2000; Pizzi & Pugliese, 2004; Wilkinson et al., 2015).

196 3.3 Geochemical dating (CHEA method)

197 Dating methods deriving from measurements of terrestrial cosmogenic radionuclide concentrations in
198 rocks and sediment are commonly used in geomorphic and neotectonic studies (Amos et al., 2007;
199 Balco, 2011; Cyr & Granger, 2008; Darvill, 2013; Granger et al., 2013). In the central Apennines, the
200 surface exposure dating techniques have been applied to footwall carbonates scarps to recover the fault
201 slip histories and paleoearthquakes (Benedetti et al., 2013; Carcaillet et al., 2008; Schlagenhauf, 2009;
202 Schlagenhauf et al., 2010; Tesson et al., 2016). This dating method relies on measuring the production of
203 ^{36}Cl cosmogenic radionuclide due to cosmic ray-induced reactions often combined with the
204 concentration of rare earth elements undergoing dissolution and recrystallization after their
205 exhumation. Thus, the fault scarp exposure at a specific site is recovered by modeling the geochemical
206 element concentration from the top to the base of the fault ribbon and several meters below the
207 ground surface. The variation of geochemical concentration is then linked to the fault slip history at the
208 measurement point under the assumption of hangingwall wedge stability and absence of hillslope
209 processes, as for LGMA method. We will refer to this method as CHEA and 73 measurements were
210 collected in the throw rate dataset (Table S1) and 69 in the coseismic throw dataset (Table S2).

211 3.4 Measuring long-term (>20 ka) markers (LONG class)

212 All these three techniques investigate the cumulative throw of the past 20 kyr. For the estimates relying
213 on greater feature ages, we defined the long-term (LONG) class. Generally, they consist of dislocated
214 geological/geomorphologic markers and the reported features are suspended high on the footwall and
215 correlated with deposits buried or outcropping at lower altitudes in the hanging wall (Blumetti et al.,
216 2013; Calamita et al., 2000; Falcucci et al., 2015; Galadini & Galli, 2000; Galadini et al., 1998; Galli et al.,
217 2015; Gori et al., 2007; Gori et al., 2011; Patruno & Scisciani, 2021; Pizzi et al., 2002; Pizzi et al., 2010;
218 Pucci et al., 2016; Pucci et al., 2019; Roberts et al., 2010). The total number of LONG data is 52.

219 4. Definition of long-term throw rates in the central Apennines

220 Inspecting the literature while searching for available data, we realized that two main issues are
221 afflicting the long-term throw rate estimations in the central Apennines.

222 The first is the need for a standard probabilistic framework for estimating long-term throw rates from a
223 single offset feature. Usually, authors have provided cumulative fault offsets ΔF since a fixed time
224 from the present, which we refer to as Δt (Table S1). In detail, for the central Apennines, stated offsets
225 ΔF at the trace of the investigated faults are reported as bounded by two measurements, as a best
226 estimate or as a lower/upper limit. Stated ages Δt are described as the best estimate or bounded by
227 two dates or as a lower limit. Suppose the offset and age uncertainties ($\sigma_{\Delta F}$ and $\sigma_{\Delta t}$, respectively) are
228 unambiguously defined. In that case, the form of ΔF and Δt distributions are explicit, and this allows
229 for a complete treatment of both in a probabilistic sense. However, we needed to make some
230 assumptions about offset and age distributions when their uncertainties weren't reported explicitly. In
231 the following subsections we describe how we determined the ΔF and Δt distributions and their
232 general trends.

233 The second issue is related to the definition of the long-term fault offset rates. Commonly, slip rates
234 have been estimated in the central Apennines by dividing the observed displacement of the faulted
235 landform by its age, that is $\Delta F/\Delta t$. However, suppose, as expected, the shear stress around the fault is
236 not equal in the starting and final instants of the investigated period. In that case, a more correct
237 equation to recover the long-term throw rate is $L = \Delta D/\Delta t$ with the total far-field displacement ΔD
238 defined as

$\Delta D = \Delta F + \Delta E$	(1)
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239 where ΔE is the (positive or negative) offset variation of elastic origin needed to recover the shear
240 stress on the fault to its starting value at Δt . In that case, neglecting ΔE could result in an “apparent
241 slip rate” which can be remarkably different from the real long-term one regardless of the slip history in
242 the investigated period. This distinction becomes less significant for older offset features, whose offset
243 is accumulated over more seismic cycles (McCalpin, 2009), and it can be neglected only for a stably-
244 slipping frictionless fault.

245 Commonly, ΔE is unknown and deeply problematic to estimate; thus, it is overlooked in slip rate
246 estimates of the central Apennines, resulting in a potentially misleading value. In subsection 4.3
247 “Definition and estimation of the net elastic throw ΔE ” we report how we handled this problem.

248

249 4.1 Single throw distributions in the central Apennines

250 In our dataset (Table S1), 287 data have a single “best estimate” throw ΔF . In 64 estimates, the stated
251 throw is bounded by two limits; in 50 it is a lower (or minimum) limit, and in 2 cases it is an upper limit.
252 These data usually refer to a vertical difference between two correlated peneplain remnants (Gori et al.,
253 2007; Miccadei et al., 2002) or the displacement sum of two splays connected at depth (Galadini & Galli,
254 1999; Roberts & Michetti, 2004).

255 The uncertainty is directly reported in 55 measurements but can be indirectly deduced for 53 CHEA
256 measurements, because the investigators furnished the standard deviation of the dip angle. We defined
257 the standard deviation as one-half of the increment due to increasing (or decreasing) the rightmost
258 significant digit by unity for data without quoted standard deviations. As an example, the inferred
259 standard deviation of an offset measurement of 7.4 m is 0.05 m.

260

261 4.1.1 Stated offset is the best estimate

262 For the 287 offsets reported as best-estimates we assumed a Gaussian distribution. As there is no
263 evidence of tectonic inversion during the Late Pleistocene/Holocene on the studied faults, it is
264 appropriate to truncate (and renormalize) the Gaussian distributions so that they are limited to positive
265 throws. That is, if we consider an offset f' , given the measured offset ΔF_n and related standard
266 deviation $\sigma_{\Delta F_n}$, its probability is given by:

$P_{\Delta F}(f') = N \left(\begin{cases} 0; & f' < 0 \\ \frac{1}{\sigma_{\Delta F_n} \sqrt{2\pi}} \exp\left(\frac{-(f' - \Delta F_n)^2}{2\sigma_{\Delta F_n}^2}\right); & f' \geq 0 \end{cases} \right) \quad (2)$
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267 where $N(\)$ is a dimensionless normalizing operator generically defined as

$N(y(x)) \equiv \frac{y(x)}{\int_{-\infty}^{+\infty} y(x') dx'}$	(3)
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268

269 4.1.2 Stated offset is bounded by two limits

270 For the 64 measurements with stated throw bounded by two limits, we assume that this range
271 corresponds to the $3\sigma_{\Delta F_n}$ interval (87%-confidence interval) and that the mean of the two limits is ΔF_n .
272 Both $\sigma_{\Delta F_n}$ and ΔF_n are then used in equation (2) to determine the distribution of ΔF (see Figure 2).

273 *4.1.3 Stated offset is a lower limit*

274 For 50 measured offsets the investigators warned that the actual throw is higher than the reported
 275 minimum ΔF_{\min} . This occurs because the coseismically-displaced layer is found only in the hanging wall
 276 but not in the footwall, from which it has been eroded. In other cases, datable layers are suspended
 277 over the footwall but buried under an unknown thickness of younger deposits in the hanging wall
 278 (Galadini & Galli, 2000; Galli et al., 2015). In this case, we can define:

$$p_{\Delta F}(f') = N \begin{cases} 0 & f' < 0 \\ p_{\Delta F}^{\text{prior}}(f') \frac{1}{2} \left[1 + \operatorname{erf} \left(\frac{f' - \Delta F_{\min}}{\sqrt{2} \sigma_{\Delta F_{\min}}} \right) \right] & f' \geq 0 \end{cases} \quad (4)$$

279 where the probability density function $p_{\Delta F}^{\text{prior}}$ is the prior estimate of long-term throw (during the
 280 measurement epoch) for the 59 faults and it is determined by bootstrap as explained in subsection 4.4
 281 “Prior slip rates and bootstrap”. See Figure 2 for the plot of the equation (4).

282 *4.1.4 Stated offset is an upper limit*

283 When the measured offset is defined as an upper limit ΔF_{\max} , the probability density function for the
 284 offset f' is a boxcar smoothed around the upper limit

$$p_{\Delta F}(f') = N \begin{cases} 0; & f' < 0 \\ \frac{1}{2\Delta F_{\max}} \operatorname{erfc} \left(\frac{f' - \Delta F_{\max}}{\sqrt{2} \cdot \sigma_{\Delta F_{\max}}} \right); & f' \geq 0 \end{cases} \quad (5)$$

285 To this aim the complementary error function (erfc) is used (Figure 2).

286 *4.2 Single age distributions*

287 The investigators have reported the feature age Δt in different ways: 107 measurements are bounded
 288 by two dates, 293 are reported as the best-estimate and 3 as lower limit. If not explicitly reported, we
 289 consider $\sigma_{\Delta t}$ as one-half of the increment caused by increasing the rightmost significant digit by unity.

290

291 *4.2.1 Two dates bound stated age*

292 Many offsets have age constrained by dated stratigraphic successions. For example, the age of
 293 earthquakes in paleoseismological trenches is often bounded by the youngest dislocated deposits (age
 294 Δt_{\min}) and the oldest sealing/overlap sedimentary units (age Δt_{\max}). Similarly, for long-term estimates,
 295 Δt is bounded by two ages obtained by radiocarbon dating of collected samples (e.g., shells, charcoal,
 296 paleosoils) from two sedimentary or tephra layers. In these cases, the probability density function of age

297 t' to represent the starting time of the observed features is expressed by a smoothed-boxcar function
 298 given as:

$p_{\Delta t}(t') = N \left\{ \begin{array}{ll} 0; & t' \leq 0 \\ \frac{1}{\Delta t_{\max} - \Delta t_{\min}} \frac{1}{2} \left[1 + \operatorname{erf} \left(\frac{t' - \Delta t_{\min}}{\sqrt{2} \sigma_{\Delta t_{\min}}} \right) \right] \frac{1}{2} \operatorname{erfc} \left(\frac{t' - \Delta t_{\max}}{\sqrt{2} \sigma_{\Delta t_{\max}}} \right); & t' > 0 \end{array} \right. \quad (6)$	
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299 *4.2.2 Stated age is the best-estimate*

300 In the vast majority of best-estimates, their value is dictated by the regional determination of the ice
 301 retreat following the Last Glacial Maximum, generally reported to occur 15±3 kyr BP. However, we
 302 noted that different authors refer to the same major climate change (the end of the Last Glacial
 303 Maximum), reporting different values and standard deviations, thus assuming different distributions.
 304 Note that the age of these features is often not experimentally determined but just assumed by field
 305 investigators. These are the common drawbacks of secondary sources summarizing the previous
 306 literature in the region without performing any new (primary) investigation. In a statistical study of
 307 western US fault-offset database (Bird, 2007), the error rate of such secondary sources was shown to be
 308 about three times that of primary literature sources. However, we do not want to risk contaminating our
 309 study by inserting these diverse climate models; therefore, we homogenized the Δt and $\sigma_{\Delta t}$ of Bubeck et
 310 al. (2015); Faure Walker et al. (2009); Faure Walker et al. (2019); Morewood and Roberts (2000);
 311 Papanikolaou et al. (2005); Roberts and Michetti (2004); Wilkinson et al. (2015) to the values suggested
 312 by Faure Walker et al. (2010) (15±3 kyr BP), defining the following probability density function of age t'
 313 (see Figure 2 for its plot)

$p_{\Delta t}(t') = N \left\{ \begin{array}{ll} 0; & t' < 0 \\ \frac{1}{\sigma_{\Delta t} \sqrt{2\pi}} \exp \left(\frac{-(t' - \Delta t_n)^2}{2\sigma_{\Delta t}^2} \right); & t' \geq 0 \end{array} \right. \quad (7)$	
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314 *4.2.3 Stated age is a lower limit*

315 Galadini and Galli (2000), relying on a series of stratigraphical considerations, determined a throw of
 316 200-340 m for a truncated valley placed at the footwall of the Campo Felice fault. The truncation onset
 317 predates the till deposition (age 0.25 Myr) and postdates the inception of the central Apennines
 318 extension (1 Myr). In two other cases (F0011NLGMA08 and F0036NTRCH04) the reported Holocene
 319 throws predate a fixed age, but for these estimates, it seems unrealistic that they formed since the
 320 inception of the extension. To avoid any artifacts resulting in implausibly low rates, we assumed that any
 321 stated minimum age Δt_{\min} also provides an order-of-magnitude estimation of the maximum age (
 322 $\approx 10 \cdot \Delta t_{\min}$) defining probability density function of age t' as

$$p_{\Delta t}(t') = N \left\{ \begin{array}{ll} 0; & t' \leq 0 \\ \frac{1}{2} \left[1 + \operatorname{erf} \left(\frac{t' - \Delta t_{\min}}{\sqrt{2} \cdot \sigma_{\Delta t_{\min}}} \right) \right] \lambda \inf \left\{ \begin{array}{l} \exp[-\lambda(t' - \Delta t_{\min})] \\ 1 \end{array} \right\} \frac{1}{2} \operatorname{erfc} \left(\frac{t' - \Delta t_{\max}}{\sqrt{2} \cdot \sigma_{\Delta t_{\max}}} \right); & t' > 0 \end{array} \right\} \quad (8)$$

323 with $\lambda = -\ln(1/2)/\Delta t_{\min}$ (see Figure 2).

324 4.3 Definition and estimation of the net elastic throw ΔE

325 Styron (2019) showed that slip rate estimates calculated over time intervals shorter than the joint
326 duration of ten earthquakes (causing slip at the sampled point on the fault trace) have the highest
327 chances of being inaccurate if they use only Δt and ΔF . This issue can be managed (by increasing
328 uncertainties appropriately) if we consider (and add) an equivalent fault offset ΔE needed for
329 recovering the fault shear stress conditions at Δt . However, the exact value of shear stress at Δt for any
330 offset feature is unknown and we can proceed only in a probabilistic way. We can speculate that under
331 the assumption of a quasi-periodic behavior of the studied fault, its shear stress is rather low after each
332 earthquake. Differently, if clustered earthquake recurrence is assumed, the potential for some greater
333 elastic offsets must be considered. One complex way to approach this issue would be to develop a
334 model investigating the variation of the elastic strain in a long-time (several million years) fault history,
335 then sampling its ΔE changes across time windows of varying lengths, with many randomly selected
336 starting times. Then, we would use empirical functions to describe how the probability density functions
337 of the modeled ΔE varies with the time-window length and with the other parameters of the
338 simulation, such as long-term fault rate and coefficient of variation, COV. This lengthy process would
339 necessarily be based on a selective choice of input parameters. However, no definitive evidence for a
340 clustered recurrence interval, and no data to score the model are available. In this differentiation, we
341 follow Styron (2019) in defining the periodic distribution as one of the earthquakes occurring more
342 regularly than randomly. Statistically, a fault with a periodic distribution of the seismicity has its
343 $\text{COV} = \sigma_{RI} / \mu_{RI} \ll 1$, where μ_{RI} is the mean recurrence interval and σ_{RI} is the standard deviation
344 of the recurrence intervals. Instead, a clustered distribution of the fault seismicity (
345 $\text{COV} = \sigma_{RI} / \mu_{RI} \gg 1$) implies the occurrence of both tightly-spaced and widely-separated earthquakes.

346 Different authors have no consensus on the earthquake recurrence interval in the central Apennines,
347 though the majority described quasi-periodic earthquake recurrences for various faults (Galli, 2020; Galli
348 et al., 2008). As an example, Galli (2020) recently reports an exhaustive review of paleoseismological
349 data showing a well-defined quasi-periodic earthquake recurrence. Conversely, slip histories found by
350 Schlagenhauf et al. (2011) and Benedetti et al. (2013) suggest that the observed displacement is not
351 periodic. Notably, in these works, the amount of coseismic slip is not related to the duration of the
352 previous interseismic interval. Hence, we decided to determine ΔE independently for each offset
353 feature, assuming that it should have a distribution resembling the one observed in the past coseismic

354 throws in the central Apennines, which sometimes represents a single event but sometimes represents
355 the total throw in an earthquake doublet or larger cluster. With this assumption, the determination of
356 ΔE is closer to the quasi-periodic distribution of seismicity, but not excluding some influence from the
357 clustered one.

358 To this end, we collected all single-earthquake throws determined with CHEA and TRCH methods. The
359 resulting paleoearthquake throw dataset is provided in Table S2, and it consists of 150 measurements
360 (69 determined with CHEA and 81 with TRCH). Investigators supplied all the needed information to rely
361 on equations (2)-(5) so we treat throw distributions in a probabilistic sense. Thus, we merged the 150
362 distributions, and the resulting PDF and CDF are shown in Figure 3. It can be noted that the two sub-
363 datasets differ substantially with CHEA method detecting bigger throws than TRCH. As CHEA
364 investigators usually sample steep slopes, we suspect that some gravitational mass-wasting effects
365 could amplify the coseismic signal locally, as recently pointed out by Di Naccio et al. (2019). In any case,
366 this marked difference in the final PDFs and CDFs is expected to capture both the stochastic variability of
367 coseismic throws and the potential for some greater elastic throws. In the last step, we fitted the
368 normalized CDF throw distribution with the exponential curve:

$P(C \leq c) \cong 1 - \exp(-\lambda c)$	(9)
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369 where c is the coseismic throw in a single earthquake. We found $\lambda = 1.0 \text{ m}^{-1}$, implying that c_m the
370 mean paleoseismic throw ($c_m = 1/\lambda$) for the central Apennines faults is 1 m. Its standard deviation (
371 $\sqrt{1/\lambda^2}$) is 1 m, too.

372

373 The development of equations for estimating the probability density function of ΔE (to recover the
374 fault shear stress level at Δt) for different measuring times follows Bird (2007). For young offset
375 features (small Δt), the model distribution of ΔE is asymmetric, with a small chance of a large
376 negative change (due to an earthquake, modeled using the distribution of c'), and a larger chance of
377 net elastic strain accumulation without an earthquake (modeled with the current bootstrap estimate of
378 the fault throw rate). For old offset features with Δt greater than typical earthquake recurrence times
379 (1-3 kyr), the distribution of ΔE will be based on mirroring the distribution of c' about zero, and
380 halving it. That is, the chances of net positive and net negative changes in elastic strain become equal,
381 and the magnitude is limited by the distribution of throws in one earthquake cluster, c' . It is necessary
382 to create a smooth transition between these two end-member cases to contemplate all possible cases.

383 We proceeded as follows: given a positive strain ($e > 0$), the probability of an even greater elastic offset (
384 $e < \Delta E$) can be expressed as the product of two factors. The first factor is the probability that e is loaded

385 in Δt according to the prior distribution of long-term offset rates. The second term is the probability
 386 that e could be released in a single pseudo-earthquake according to the exponential curve determined
 387 above. That is:

$$1 - P(\Delta E \leq e | \Delta E > 0) = P(e < L^{\text{prior}} \Delta t) P(e < C) \quad (10)$$

388 which becomes in terms of probability density function:

$$p_{\Delta E}(e') \equiv -P(e' < L^{\text{prior}} \Delta t) \frac{d}{de'} P(e' < C) - P(e' < C) \frac{d}{de'} P(e' < L^{\text{prior}} \Delta t) \quad (11)$$

389 With its complete form being:

$$z(e') = \exp(-\lambda e') \left[\lambda \int_{e'0}^{\infty} p_L^{\text{prior}}(\ell') p_N\left(t' = \frac{e''}{\ell'}\right) \frac{1}{\ell'} d\ell' de'' + \int_0^{\infty} p_L^{\text{prior}}(\ell') p_{\Delta}\left(t' = \frac{e'}{\ell'}\right) \frac{1}{\ell'} d\ell' \right] \quad \text{for} \quad (12)$$

$e' > 0$

390 as the expectation of $p_{\Delta E}$ is zero, the possibility of a negative offset is weighted with the factor k and
 391 jointly normalized in a way that there is an equal chance of having a gain or a loss in elastic offset:

$$p_{\Delta E}(e') \equiv N \left(\begin{cases} kp_c(c' = -e') & e' \leq 0 \\ z(e') & e' > 0 \end{cases} \right) \quad (13)$$

392 We performed some exploratory tests and determined that any reported offset younger than 2 ka is
 393 likely to be strongly affected by such elastic offset (Figure S1) for the studied faults. In our method, for
 394 offset age < 2 ka, the missing or excess elastic offset is mainly dependent on the assumed recurrence
 395 interval distribution, which in turn can be largely discordant with the real one of the single fault. Thus,
 396 we opted for the conservative choice of discarding any measured offset younger than 2 ka (Table S1).
 397 Similarly, we argue that throw rate variations based on offset features younger than 2 ka are probably
 398 dominated by short-term elastic behavior, thus likely to be not consistent.

399 4.4 Prior throw rates and bootstrap

400 In two cases (discussed above) we had to deal with incomplete information: a) when only a lower limit
 401 on the throw is available, $p_{\Delta F}^{\text{prior}}$ is required (see equation 4); b) when we need to estimate the
 402 probability that a missing or exceeding offset e' is higher than the measured throw (see equation 12),
 403 according to the prior distribution p_L^{prior} . In both cases, we can use the available measurements of all
 404 faults to estimate the corresponding p_L^{prior} . To this aim, we performed a two-step bootstrap estimation.
 405 In the first run, uniform "boxcar" distribution ($0 < L^{\text{prior}} \leq 1.5$ mm/a) was set as model of the prior throw-
 406 rate probability density function for each single fault. After we computed all data-based $p_L(\ell')$, they
 407 were averaged into one posterior throw-rate distribution, in which we have excluded the unbounded

408 rates, where the stated offset is a lower limit. In the second step, the posterior distribution was used as
 409 the prior one. After three iterations, the posterior distribution was almost entirely coincident with the
 410 prior one and we considered the datum-based prior distribution p_L^{prior} well-converged. Then, assuming
 411 $\Delta E = 0$, we can calculate $p_{\Delta F}^{\text{prior}}(f')$, the prior distribution for the long-term fault throw

$$p_{\Delta F}^{\text{prior}}(f') = \int_{-\infty}^{\infty} p_L^{\text{prior}}(\ell') p_{\Delta}(t' = \frac{f'}{\ell'}) \frac{1}{\ell'} d\ell' \quad (14)$$

412 and use it in equation 4 (where only a lower limit is available) to determine the distribution of throw
 413 $p_{\Delta F}(f')$.

414 4.5 Long-term offset rate L for a single measurement

415 Once $p_{\Delta F}$ and $p_{\Delta E}$ have been determined, we convolve them to calculate the probability density
 416 function of the far field long-term offset defined as

$$p_{\Delta D}(d') = N \left(\begin{cases} 0 & d' \leq 0 \\ \int_0^{+\infty} p_{\Delta E}(e') p_{\Delta F}(f' = d' - e') de' & d' > 0 \end{cases} \right) \quad (15)$$

417 and finally, we can determine the probability density function of the long-term throw rate p_L for each
 418 reported measurement as

$$p_L(\ell') = \int_0^{+\infty} p_{\Delta}(t') p_{\Delta D}(d' = \ell' t') t' dt' \quad (16)$$

419 This equation allows defining the probability of value ℓ' to be the long-term throw rate based on a
 420 single offset feature.

421 5. Pattern of long-term throw rates across the central Apennines

422 After inspecting the available literature for active normal faults in the central Apennines, we collected
 423 403 different throw measurements performed with four different techniques (Table S1 and Figure 1),
 424 which are plot against reported age for each single fault in the Figure S2. Numerical-modeling methods
 425 applied to ^{36}Cl -cosmogenic radionuclide and earth trace elements concentrations in footwall carbonates
 426 (CHEA) allowed the dating of 73 throws. Last post-glacial (12-20 kyr) geomorphic markers and deposits
 427 recognition (LGMA) resulted in 232 data reporting the vertical separation of originally subplanar or
 428 sublinear features. Field mapping of dislocated Middle-Late Pleistocene stratigraphic layers/geomorphic
 429 features (LONG technique) resulted in 52 offset rates. Paleoseismic trenching studies (TRCH) constrained
 430 46 offsets and related ages.

431 A closer inspection of the sampled ages reveals that the most-investigated time interval of central
432 Apennines fault evolution is the past 25 kyr, and the most sampled time interval being the $\Delta t \approx 15$ kyr
433 (Figure 4). This peak is due to a large number of LGMA measurements. A second peak is found at ~ 3 -5
434 kyr, roughly corresponding with the most abundant estimates of TRCH and CHEA techniques.

435 We found that the median of all fault throw rate estimates is 0.5 mm/yr, with a 90%-confidence range
436 of 0.08-1.48 mm/yr (Figure 5). However, this is the only phase of our study in which rates from different
437 normal faults are combined; in the following phases, we consider and model only rates measured along
438 one fault.

439 Subdividing the rates by measurement technique (but still lumping all faults together), the 90%-
440 confidence range for LONG constrained throw rate is 0.06-1.16 mm/yr (median 0.42 mm/yr). The
441 general pattern of TRCH throw rates (90%-confidence range 0.04-1.54 mm/yr, median 0.36 mm/yr) is
442 similar to the LONG ones, with the main difference resting in higher probabilities for fast rates. This
443 discrepancy occurs because the offsets in normal faults are often reported as a lower limit, potentially
444 allowing higher values, whereas the age of deformed Middle-Pleistocene units or features is usually well
445 constrained. The confidence range is 0.10-1.40 mm/yr (median 0.50 mm/yr) for LGMA estimates,
446 instead, the 90%-confidence range limit for CHEA is 0.14-1.74 mm/yr with median 0.68 mm/yr (Figure
447 5). The throw rates determined for each fault are plotted against age in Figure S3.

448 The cumulative distribution of the throw rates in the central Apennines can fit a lognormal distribution
449 with a mean of 0.63 mm/yr and a standard deviation of 0.55 mm/yr (red line in Figure 5).

450 5.1 Treatment of uncertainties for ΔF and Δt

451 The uncertainty quantification for both ΔF and Δt is different among techniques and authors. Often a
452 simple experimental uncertainty (observation error) is provided whereas in few cases additional
453 epistemic uncertainty is discussed. However, a homogeneous treatment of uncertainties is required for
454 a coherent data comparison. Following the popular distinction between “aleatory” and “epistemic”
455 uncertainties, we can assume that they are uncorrelated for all throw measurements of the central
456 Apennines and that the “aleatory” uncertainties are mainly defined by the “observational” uncertainties
457 in our paper. Thus the total uncertainty for each measurement would be defined from the Pythagorean
458 sum of both aleatory and epistemic uncertainties.

459 We inspected how the observational uncertainties vary in time for all features of our database by
460 plotting ΔF and Δt 95% Confidence Levels (95% C.L.). We limited our analyses to $\Delta t < 25$ kyr, as most
461 fault throw rates were collected in this time interval (Figure 4).

462 Both confidence intervals expand in time, though the increase is more evident for Δt . With the notable
463 exception of LGMA measurements, $\sigma_{\Delta t}$ was provided by laboratory analyses. The expertise of field

464 investigators may be adequate to avoid or limit any other age uncertainty source, also considering the
465 relatively small Δt . Thus, the observation error reasonably corresponds to the total uncertainty (
466 $\sigma'_{\Delta t} = \sigma_{\Delta t}$), and it is almost 15% of Δt itself (Figure 6). We recall that for many LGMA measurements, we
467 homogenized the Δt uncertainties to the values suggested by Faure Walker et al. (2010). Note that the
468 age of these features is often not experimentally determined but just assumed by regional features
469 correlation. This assumption may be a severe limitation for using these data because it could become a
470 tight constraint on the resulting offset rate when paired with a too-narrow definition of $\sigma_{\Delta t}$. However,
471 we consider the reported $\sigma_{\Delta t} = 3$ kyr a sufficiently large uncertainty to further use these data for long-
472 term estimates.

473 The 95% C.L. of ΔF is less pronounced and more variable than the Δt one and can be approximately fit
474 with a linear increase with age of the feature. The smaller increase as age function of ΔF uncertainties
475 to Δt ones is not surprising because the investigators often report only the observation errors. In this
476 case, it is questionable to assume that epistemic uncertainties are negligible. For example, in LGMA
477 measurements, ΔF is often provided with the precision of a metric ruler. Still, exhaustive discussion
478 and quantification of other uncertainty sources define much greater values (as an example, see
479 Papanikolaou et al., 2005). We will show in section 6 how influential is a correct estimation of the total
480 uncertainty in long-term estimates.

481 6. Determining fault throw rate variance

482 6.1 Noisy throw rate profiles

483 The throw and throw-rate profiles measured along several active faults taper out with characteristic
484 gradients towards the fault tips. Different slip (and slip-rate) profiles collected in several tectonic
485 environments seem to be related to the fault dynamics. Inelastic deformation governs linear slip
486 gradients near the tips, and elliptical shapes describe purely elastic slip distributions (Scholz & Lawler,
487 2004). We plotted the throw-rate profiles for some faults sampled with at least ten measurements
488 points and three different techniques to capture a typical profile for the central Apennines active faults.
489 We followed Manighetti et al. (2001), normalizing the distance to the fault length and the throw-rate to
490 the highest median value among the data of each fault (Figure 7 and Figure S4).

491 The visual inspection of along-strike throw-rate profiles clearly shows a high scattering among data,
492 which does not define any clear throw-rate shape. This overdispersion requires a better investigation of
493 both uncertainty sources and data contamination before assigning any specific rheological meaning to
494 the observed throw rate profiles.

495 6.2 Throw rate disagreement curves for the central Apennines

496 Regardless of the slip gradients near unconnected fault tips or the fault history, two features close in
 497 time and space and sampling the same fault activity are expected to provide comparable long-term
 498 estimates. Contrarily, more distant (in space or in time) features are expected to sample more different
 499 rates. This contrast between close and distant data pairs should appear in some form when inspecting
 500 the dataset of active faults if throw rates are uncontaminated by noise. This step is fundamental to
 501 discriminate spatial and temporal variabilities from throw rate variance. The former is a tectonic signal
 502 and the latter is noise to be addressed by further modeling, as an example in PSHA. We can distinguish
 503 these two-end members if a clear definition of disagreement among throw rate estimates of the same
 504 fault is available.

505 Because the long tails of most individual throw rate distributions (PDFs) at single sites are potentially
 506 overlapping to some extent, we opted for a probabilistic definition of disagreement to a purely binary
 507 classification. Let us consider two throw-rate probability distributions of two independent
 508 measurements A and B. If we define a value x (Movie S1), a sufficient condition for the two
 509 independent measurements to be different is that A and B are always separated by x . Thus we can use
 510 the scalar quantity δ :

$\delta \equiv \sup_{x=0}^{\infty} \{P(A < x \leq B) + P(B < x \leq A)\}$	(17)
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511 which becomes in our case

$\delta \equiv \sup_{x=0}^{\infty} \{P_L^A(A < x)[1 - P_L^B(B < x)] + P_L^B(B < x)[1 - P_L^A(A < x)]\} =$	(18)
$\sup_{x=0}^{\infty} \{P_L^A(A < x) - 2P_L^A(A < x)P_L^B(B < x) + P_L^B(B < x)\}$	

512 The possible values of disagreement δ span between 0.5 and 1. If the median of the two distributions is
 513 coincident, then $\delta = 0.5$ regardless of the stated standard deviations. However, if the two distributions
 514 are completely discordant without any overlapping interval, then $\delta = 1$ (Movie S2).

515 Calculating the disagreement δ makes sense only if we consider data pairs sampling the rate of the
 516 same fault. To this aim, we assigned each throw rate to its fault and determined the disagreement δ of
 517 it to all other throw rates collected along the same fault. This operation resulted in 3466 data pairs for
 518 the 59 fault traces. As we are interested in the overall trends of our dataset, we merged and sorted the
 519 3466 disagreements building the cumulative distribution of disagreements shown in Figure 8.
 520 Furthermore, we split the disagreement dataset into several subsets to inspect the variance among
 521 different methods in sampling the fault throw rates. The resulting curves are also plotted in Figure 8,
 522 where we also report along the distribution curve the value β describing the fraction of highly-
 523 disagreeing data pairs ($\delta \geq 0.95$)

524 It can be noted in Figure 8 that β is never the highest for data collected with the same method. The
 525 only exception is LONG measurements ($\beta = 0.28$), which examine larger time intervals and likely sample
 526 actual long-term variations. However, their scarce number, along any studied fault, doesn't allow
 527 investigating this conjecture in greater detail. The worst pairing for the remaining methods is between
 528 data pairs comparing LGMA measurements ($\beta = 0.16$). Also for CHEA and TRCH the worst pairing is with
 529 LGMA ones ($\beta = 0.11$ for CHEA-LGMA; and $\beta = 0.07$ for TRCH-LGMA). This increase of β when pairing
 530 different methods implies that the variability in throw rates is most apparent when the fault has been
 531 investigated with alternative techniques.

532 To discriminate the noise from signal, we sorted the 3466 disagreements according to spatial separation
 533 of the data pair. We also sorted the δ dataset by time-span between the ages of the two offset features
 534 in each data pair. If spatial separation were important, then disagreements between data should
 535 increase systematically with distance, regardless of the specific fault kinematic model. Likewise, if time-
 536 separation were important, then disagreements between data should increase systematically with the
 537 difference in ages of the offset features, regardless of the specific time-history proposed. Figure 9 shows
 538 the graphical analysis of β (fraction of high disagreement data) versus age-differences and relative
 539 along-fault positions between all possible data pairs sampling the throw rate for the same fault. These
 540 graphs do not support either space or time as critical independent variables. Instead, they seem to
 541 suggest a high degree of randomness as a distinctive characteristic of the dataset, as already shown by
 542 the throw-rate profiles along certain well-studied fault traces (Figure 7). These spatial and temporal
 543 patterns also seem to be confirmed at the single-fault scale (Figure S3 and Figure S4). Furthermore, the
 544 slight reduction in disagreements for separations under 1 km probably results from the increased
 545 likelihood that such pairs used the same measurement technique and even came from the same
 546 published study.

547 6.3 Determination of ΔF epistemic uncertainties ($\sigma'_{\Delta F}$) with throw rate forward modeling and 548 disagreement curve comparison

549 The disagreement curves recovered for the central Apennines have strong implications for throw rates
 550 estimates and uncertainty quantification. In general, when the disagreement of a data pair approaches
 551 one ($\delta \geq 0.95$), three alternatives become equally valid and mutually exclusive: a) both data are erratic
 552 (non-tectonic) measurements; b) the data pair is formed from one correct datum and one erratic one; c)
 553 the two values are correct and the high disagreement comes from the measurement uncertainties.
 554 Considering that β is the percentage of pairs with $\delta \geq 0.95$ in our real dataset, these three alternative
 555 possibilities can be merged and used to determine the fraction of incorrect data α solving the
 556 equation:

$\beta = \alpha \cdot \alpha + (1 - \alpha) \cdot \alpha + \alpha \cdot (1 - \alpha) + \gamma \cdot (1 - \alpha) \cdot (1 - \alpha)$	(19)
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557 where γ represents the possibility that $\delta \geq 0.95$ results from sampling the tails of two correct data pair
558 distributions.

559 For the disagreement curve of real data, we have determined $\beta = 0.13$ (Figure 8). If we assume that field
560 investigators were able to detect tectonic features and distinguish them from non-tectonic, then we
561 would expect $\alpha \approx 0$ and $\beta \approx \gamma$ in equation (19). Differently, for $\beta > \gamma$ and $\alpha > 0$ we have to suppose
562 that field investigators misinterpreted some non-tectonic features as tectonic ones. Thus, the
563 calculation of γ is crucial for discriminating between these two alternatives.

564 To this end, we need to build different datasets characterized by the same uncertainties in both ΔF and
565 Δt measurements. In section 5, we have found that the 95% C.L. of both ΔF and Δt can be modeled by
566 linear functions of age of the offset feature. Assuming that the 95% C.L. roughly corresponds to 4σ (as it
567 does in a normal distribution), we can assume that for each datum (real or synthetic) the observational
568 $\sigma_{\Delta F}$ can be replaced with $\sigma'_{\Delta F}$ having the value of $\sigma'_{\Delta F} = \Delta \tilde{t} \cdot 0.05$ mm/yr where $\Delta \tilde{t}$ represents the
569 median of the Δt distribution. Instead, the observational age uncertainty is set to $\sigma'_{\Delta t} = \Delta \tilde{t} \cdot 0.15$.

570 In a first model we assumed $\sigma'_{\Delta t}$ and $\sigma'_{\Delta F}$ to coincide with the nominal observational (aleatory)
571 uncertainties and we used the ΔF and Δt measurements as in Table S1. In Figure 10 we show the
572 similarity of this modeled disagreement curve (black line in Figure 10A) to the real one (dotted red line
573 in Figure 10A). These similar trends suggest that the linear functions to the feature age are good proxy
574 of observational uncertainties. This allows performing a synthetic test to determine γ from the
575 theoretical distribution curves of disagreements between any purely tectonic offset rates.

576 To this end, we created 135 slip (throw) histories using the observed throw rates and coseismic throws
577 (Table S1 and Table S2). In detail, the cumulative distributions indicate that the throw rate dataset
578 describes the activity of a series of active faults with 0.63 mm/yr (and standard deviation 0.55 mm/yr,
579 Figure 5) long-term throw rate and 1 m (standard deviation 1 m) mean earthquake displacement (Figure
580 3). Considering the uncertainties on both the throw rate and throw distributions, we assumed a return
581 period of 1600 ± 1600 yr for each of the 135 modeled fault histories.

582 We sampled ten times each fault history in the 2-25 kyr time interval, which contains 87% of the
583 observed throw rate measurements. Then we perturbed each age and throw with a Gaussian noise of
584 $\sigma'_{\Delta t} = \Delta \tilde{t} \cdot 0.15$ and $\sigma'_{\Delta F} = \Delta \tilde{t} \cdot 0.05$ mm/yr. Finally, we randomly selected Δt and ΔF in the perturbed
585 distribution intervals and built a synthetic feature dataset. We meta-analyzed this synthetic dataset in
586 the same way as the real one to determine the Δt , ΔF and ΔE distribution curves of each synthetic
587 measurement and then its long-term throw rate distribution L. This synthetic dataset contains ideal
588 tectonic measurements perturbed only by observation errors. Thus, we determined the disagreement

589 curve for this ideal dataset as done for real data and shown in Figure 8, by comparing the long-term
590 throw rates of each data pair sampling the same active fault. From the resulting disagreement curve,
591 shown in Figure 10A, we determined $\gamma=0.03$.

592 Recalling equation 19 and the fraction of high disagreeing data observed in our dataset ($\beta=0.13$), then
593 considering $\gamma=0.03$, we obtain $\alpha\approx 0.05$. This result implies two different scenarios: if we assume that the
594 observation uncertainties ($\sigma_{\Delta t}$ and $\sigma_{\Delta F}$) coincide with the total ones ($\sigma'_{\Delta t}$ and $\sigma'_{\Delta F}$), then several data
595 (5%) sampled non-tectonic features along the fault traces. Alternatively, if we maintain that field
596 investigators have the expertise to detect tectonic features and distinguish them from non-tectonic
597 ones, then the ΔF total (observational and epistemic) uncertainties ($\sigma'_{\Delta F}$) are higher than the
598 observational errors $\sigma_{\Delta F}$. We opted for this latter alternative.

599 We focused on $\sigma'_{\Delta F}$ because the short length of the investigation period (<25 kyr), the steady-state of
600 long-term geodynamic settings for the central Apennines, and the laboratory uncertainties suggest that
601 total and observation uncertainties may not differ significantly for the age of observed features thus we
602 set always $\sigma'_{\Delta t} \approx 0.15 \cdot \Delta t$. Then, we tested different values of $\sigma'_{\Delta F}$, used it to determine the Δt , ΔF and
603 ΔE distribution curves of each datum and then the long-term throw rate distributions L . We built the
604 disagreement curve and compared it to the synthetic one searching for the $\sigma'_{\Delta F}$ value minimizing the
605 difference between β and γ of two curves. Finally, we found that $\beta \approx \gamma$ ($\alpha \approx 0$) for $\sigma'_{\Delta F} = \Delta \tilde{t} \cdot 0.10$
606 mm/yr (Figure 10C). In Figure 6 we have plotted all tested $\sigma'_{\Delta F}$ and in red the optimal value
607 $\sigma'_{\Delta F} = \Delta \tilde{t} \cdot 0.10$ mm/yr. It can be observed in Figure 6 that this total uncertainty doubles or triples the
608 observational one ($\sigma_{\Delta F}$) of LGMA estimates, which are largely predominant in for $10 \text{ kyr} < \Delta t < 20 \text{ kyr}$.
609 For graphical inspection, we report single throw rate estimates L for both nominal (observational) and
610 total ($\sigma'_{\Delta F} = \Delta \tilde{t} \cdot 0.10$ mm/yr) uncertainties in Figure S5.

611

612 7. Fault long-term throw rates

613 The meta-analyses performed on our throw rate dataset suggest that it is affected by noise
614 contamination causing overdispersion with respect to the computed resultants of nominal observational
615 uncertainties. For this reason, we propose a base model (null hypothesis) of uniform throw rate along
616 each fault trace (except tapering to zero at unconnected ends). Also, we consider that temporal
617 variations in throw-rate remain unproven, and may constitute over-fitting. To this aim, we recalculated
618 the prior throw rate distribution p_L^{prior} and the single-measurement throw rate distributions $p_L(\ell')$

619 using the new age-dependent epistemic uncertainty $\sigma'_{\Delta F}$ for each datum. Then, we determined the
 620 fault long-term throw rates combining all the measurements I of each fault as

$$\left\{ p_L^{combined}(\ell^i) \right\}^i \cong N \left[\left\{ p_L^{combined}(\ell^i) \right\}^{(i-1)} \left\{ p_L^i(\ell^i) \right\} \right] \quad (20)$$

621 for all i in the range $1 \leq i < I$ with $\left\{ p_L^{combined}(\ell^i) \right\}^0 = 1$. The 95% C.L. and the median values of the
 622 probability density function of the combined throw rate $p_L^{combined}$ are reported in Table 1 and shown in
 623 Figure S6. Median values are plotted on map in Figure 11.

624 The most prominent spatial pattern of the resulting estimates is the almost-uniform throw rate for the
 625 faults in the central sector of the chain, which roughly corresponds with the area characterized by the
 626 highest geodetic strain rates (Carafa et al., 2020).

627 The fastest slipping faults (> 0.97 mm/yr) are Liri, Fiamignano, Barisciano-San Pio delle Camere and
 628 Assergi. These faults are not adjacent and are located in different sectors of the chain. Given that the
 629 maximum extension of the central Apennines migrated eastward since Pliocene (Calamita & Pizzi, 1994;
 630 Piacentini & Miccadei, 2014), the distribution of these fast-slipping faults could be linked to the
 631 geodynamic evolution of the chain. However, this hypothesis is weak because these fast-slipping faults
 632 are close to other slow-to-intermediate slipping ones.

633 For assessing the relative precision of each estimate, we defined the coefficient of variation (COV) as the
 634 ratio of the 95% confidence interval to the median of the throw rate distribution (Figure 11). COV values
 635 lower than one intuitively suggest a pretty well-defined throw rate in the number of measurements and
 636 their compatibility. However, COV values close or above one characterize scarce and/or mostly
 637 scattered offset features, as the Assergi fault. Figure 11 shows that the north-western and easternmost
 638 active faults are less investigated, and more effort is needed to narrow the uncertainty on their activity
 639 rates.

640 8. Discussion

641 We applied a meta-analysis to dated throw features on the active normal faults of the central
 642 Apennines, and determined their long-term throw rates. We found a large overdispersion of rates with
 643 respect to nominal observational uncertainties which implies that total uncertainties are two to three
 644 times bigger than the stated observational uncertainties. These results agree with the findings of
 645 McPhillips and Scharer (2018) and indicate that natural (but non-tectonic) or anthropogenic landscape
 646 modifications justify the augmentation of the observational errors by epistemic errors when computing
 647 long-term throw rate estimates. Here we discuss some possible causes.

648 8.1 Method-related noise

649 We have shown that variations in fault throw rate are apparent when inspecting data collected with
650 different techniques applied to the same fault, and any method-specific systematic bias is revealed by
651 comparing measurements taken with alternative methods. For example, in ^{36}Cl dating, the sampling bias
652 favors slip estimates fitting the bottom of the fault scarp, the portion with higher data density (Goodall
653 et al., 2021). This method limitation can be unveiled only by comparing different techniques and their
654 individual assumptions.

655 Different strategies for selecting the measurement sites along the same fault seem to further increase
656 the probability of sampling slip rate randomness or non-tectonic effects. Thus, it is not surprising that
657 well-trained investigators contested the alleged tectonic source of some previously reported offset
658 features. Goodall et al. (2021) supposed that some fast throw rates on the Middle Aterno Valley fault
659 (F0015NCHEA01, F0015NCHEA02, and F0015NCHEA03 offset measurements in Table S1) are due to net
660 erosive slope processes that likely removed material from the hangingwall slope resulting in a higher
661 scarp. Kastelic et al. (2017) proved an equivalent non-tectonic short-term exposure rate for several
662 faults in the central Apennines and some measurement points along the Campo Felice fault scarp
663 investigated by Goodall et al. (2021).

664 Another severe method-related limitation is the minimum sampled age. Using a bootstrap strategy to
665 estimate the size of shear stress variations around a typical fault during its simulated earthquake
666 history, we showed that offset features younger than 2 ka must be considered to give only an “apparent
667 tectonic rate”. This estimate may be quite different from the long-term one, because the state of shear
668 stress surrounding the fault may have changed, and this effect is material for small/young offsets (Figure
669 S1). For these estimates, some increase in the throw uncertainty is needed before determining the long-
670 term throw rates (which, by definition, would not include any short-term elastic-strain effects). This
671 point confirms our findings on method-related variance not being marginal; it may have been
672 overlooked by other investigators focused on proving spatial and temporal variability of fault slip rates.

673 8.2 Surface processes noise

674 In the coseismic phase, gravity and rheological variations are well-documented to amplify the tectonic
675 signal at the ground surface (Chen et al., 2012; Di Naccio et al., 2019; Roback et al., 2018). That is, the
676 height of a tectonic normal fault scarp may be amplified by incipient landsliding and compaction in the
677 hangingwall, especially when it is composed of incoherent talus. In these circumstances, even extensive
678 field investigations can hardly distinguish non-tectonic from tectonic deformation in the near-surface
679 due to the geomorphic convergence of the alternative or coexisting processes. Thus, it is likely that in
680 some peculiar situations, the reported throw-rate feature is magnified (or reduced) by the effect of
681 unrelated parameters or concurrent processes (Kastelic et al., 2017; Stemberk et al., 2019).

682 We suspect that such spurious amplifications of the tectonic signal are most likely at measurement sites
683 located on steep slopes. Some form of non-tectonic gravity-driven scarp growth is a significant
684 complicating factor in such settings, amplifying the tectonic signal. Previous analysis of coseismic offsets
685 measured after two main earthquakes that occurred in the central Apennines (Di Naccio et al., 2019)
686 have proved that above the critical slope angle of 30° slope failures may occur, and rock scarps are
687 expected to be amplified as a result. The extra accelerations and extra stresses due to seismic wave
688 passage can perturb the static equilibrium and cause the hanging wall to move short distances in short
689 times.

690 Also, the formation and evolution of the bedrock scarp may have been controlled by erosional
691 processes, which would act in the other direction, to reduce or hide the tectonic signal. This process has
692 been suggested by Galadini (1999) for the Liri fault (F0029N measurements in our dataset), which places
693 the highly-eroded Miocene flysch in contact with Mesozoic carbonate bedrock without tectonic
694 evidence for Late Pleistocene deposits. Furthermore, Papanikolaou et al. (2005) argued that, even if the
695 site selection was very careful, the fault scarp has inevitably been altered to some extent by weathering
696 and deposition, resulting in an underestimation of post-glacial throw.

697 8.3 Future developments

698 Our meta-analysis shows that overdispersion is a distinctive data pattern in the central Apennines.
699 Consequently, we argue that the throw rate dataset cannot support any general and sophisticated fault-
700 growth model. For this reason, we attempted to calculate a model with the minimum number of
701 degrees of freedom needed for seismic hazard analysis: one uniform throw rate for each fault. Following
702 the “Occam’s razor” principle, the most defensible estimate would be only a uniform mean throw rate
703 with a clear definition of its uncertainty. And this is valuable because, up to now, no throw (or slip) rates
704 have been provided for the central Apennines that merged different estimates in a rigorous and
705 probabilistic framework, as we have.

706 Specifically, we are not contesting that fault throws and throw rates taper to zero at some unconnected
707 fault tips. Instead, (now) available data in the central Apennines do not allow defining the shape of this
708 tapering, which could be sharp at the tips or gentle and pervasive from the center of the fault. This
709 issue could be resolved in the central Apennines with a dramatic increase of measurements and
710 analyses on both coseismic throws and long-term throw rates. In this sense, our analyses support a
711 further increase in geological studies as recently done by Schirripa Spagnolo et al. (2021), paired with
712 deep seismic profiles (Buttinelli et al. 2021; Patruno & Scisciani 2021,) and geodetic data (Carafa et al.,
713 2020) for a better definition of single-fault throw rate, possibly varying in space and time.

714 Only three earthquake sequences in the past century (the 1915 Fucino earthquake sequence, the 2009
715 L'Aquila earthquake sequence, and the 2016-2017 Amatrice-Norcia seismic sequence) left continuous

716 and measurable marks. Thus, any long-term throw rate model based on fault trace estimates currently
717 lacks enough data to be validated without any circular reasoning. Consequently, extreme caution (and
718 complete definition of observational and epistemic uncertainties) is needed for seismic hazard modelers
719 in using fault data.

720 Furthermore, we do not exclude that active faults could appear on the ground surface as arrays of en-
721 echelon segments whose offsets are governed by laterally-variable rheology. In this context, non-
722 characteristic and clustered earthquakes can amplify offset rate variability in space and time. Meta-
723 analysis on the regional dataset using fault trace data alone cannot solve these issues due to data
724 scarcity and noise contamination. Instead, lateral variability of fault activity can be better assessed if
725 fault trace data are paired with mineralogical analyses on the composition of carbonate-hosted faults
726 (Smeraglia et al., 2018; Smeraglia et al., 2017), because lateral variability of fault friction is expected to
727 become evident. As in Doglioni et al. (2011) or Carafa and Barba (2011), patterns in fault rheology can
728 then be included in rheological models to better predict and understand along-strike variability. Also,
729 geochemical studies on fault-trace CO₂ degassing rate (Chiodini et al., 2020; Chiodini et al., 2004) and the
730 secular variation in geodetic time series can perhaps be related to the date of the most recent
731 earthquakes (detected with paleoseismology and reported in Table S2) to discriminate whether faults
732 exhibit any measurable time-dependent pattern(s).

733 9. Conclusions

734 We found that fluctuations in throw rates on active faults in the central Apennines are larger than the
735 expected resultants of nominal observational errors. We also proved that their real tectonic significance
736 remains highly questionable because these fluctuations are dominated by data overdispersion. If the
737 apparent throw rate variability results from a long-term tectonic process, that would have crucial
738 implications, justifying more complex time-dependent models in seismic hazard assessment. However,
739 independent data (e.g., known climate or geodynamic changes) must parallel and thus confirm any
740 apparent slip rate changes, which, otherwise, solely reflect both the inherited stochasticity of the
741 earthquake processes and noise contamination.

742 Given the current understanding of temporal and spatial fault throw rate variability in the central
743 Apennines and the epistemic uncertainties connected to the throw rate estimates, it is not yet feasible
744 to produce more complex PSHA models. A base model with a uniform throw rate along the trace
745 (tapering to zero at unconnected fault tips), and merging information from offset features of different
746 ages to constrain a single time-independent rate, is still the most reasonable one.

747

748

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754 policies.
755

756 **Data Availability Statement**

757 - All estimates have been obtained with a modified version of Slippery code (available at
758 <http://peterbird.name/oldFTP/Slippery/Slippery.f90.txt>) adapted for the central Apennines
759 kinematics and attached as Text S1.
760

761 - All data are cited and referenced in the text and are included in the following papers:
762 Benedetti et al., 2013; Blumetti et al., 2013; Boncio et al., 2010; Bubeck et al. 2005; Calamita et
763 al., 2000; Carcaillet et al. 2008; Cinti et al., 1992; Cinti et al., 2011; Cowie et al., 2017; D'Addezio
764 et al., 2001; Falcucci et al., 2015; Faure Walker et al., 2009; Faure Walker et al., 2010; Faure
765 Walker et al., 2019; Galadini and Galli, 1999; Galadini and Galli, 2000; Galadini et al., 1998;
766 Galli et al. 2010; Galli et al. 2012; Galli et al. 2015; Giraudi and Frezzotti, 1995; Giraudi, 1988;
767 Giraudi, 1989; Giraudi, 1995; Goodall et al., 2020; Gori et al. 2007; Gori et al., 2011; Gori et al.,
768 2017; Manighetti et al. 2010; Miccadei et al., 2002; Morewood and Roberts, 2000; Moro et al.,
769 2013; Pantosti et al, 1996; Papanikolaou et al., 2005; Patruno and Scisciani, 2021; Piccardi et al.,
770 1999; Pizzi and Pugliese, 2004; Pizzi et al., 2002; Pizzi et al., 2010; Pucci et al., 2016; Pucci et al.,
771 2019; Roberts and Michetti, 2004; Roberts et al., 2010; Salvi et al., 2003; Saroli et al., 2008;
772 Schlagenhauf et al., 2009; Schlagenhauf et al., 2011; Tesson et al., 2016; Villani et al., 2015;
773 Wilkinson et al., 2015.
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1085 Captions

1086 **Figure 1.** Map of the active faults and measurement points across the central Apennines. The
1087 measurement points are distinguished by used technique: LGMA = Last Glacial Maximum; CHEA =
1088 geochemical analysis in scarps; LONG = long-term geologic offsets; TRCH = paleoseismic trenching. Refer
1089 to the Table S1 for fault code, offset, and age of each measurement.

1090 **Figure 2.** Plots of samples of different types of probability density functions (PDFs) used for both throw
1091 (ΔF) and age (Δt) distributions.

1092 **Figure 3.** Frequency distribution of paleoearthquake throws in the central Apennines. The black lines
1093 show the cumulative distributions; grey circles indicate 5th, 50th and 95th percentiles of each distribution.
1094 The red line is the exponential function fitting the cumulative distribution of paleoearthquake throws.

1095 **Figure 4.** Frequency distribution of sampled age for the complete dataset (top) and for different
1096 techniques (bottom). The black lines show the cumulative distributions; grey circles indicate 5th, 50th and
1097 95th percentiles of each distribution. Throw-rate measurement techniques: numerical-modeling
1098 methods applied to trace elements and cosmogenic radionuclides (CHEA), geomorphic markers exposed

1099 since the Last Glacial Maximum (LGMA), long-term (>20 kyr) estimates (LONG), and paleoseismic
1100 trenching (TRCH).

1101 **Figure 5.** Frequency distribution of throw rates for the complete dataset (top) and each method
1102 subdataset (bottom). The black lines show the cumulative distributions; grey circles indicate 5th, 50th and
1103 95th percentiles of each distribution. Throw-rate measurement techniques: numerical-modeling
1104 methods applied to trace elements and cosmogenic radionuclides (CHEA), geomorphic markers exposed
1105 since the Last Glacial Maximum (LGMA), long-term (>20 kyr) estimates (LONG), and paleoseismic
1106 trenching (TRCH). The red line is the lognormal function fitting the cumulative distribution of throw
1107 rates.

1108 **Figure 6.** 95% Confidence Level (C.L.) for both ΔF and Δt . Top: measured 95% C.L. for ΔF and tested total
1109 uncertainties $\sigma'_{\Delta F}$ (see Figure 10). Bottom: 95% C.L. for Δt and tested total uncertainties σ'_Δ .

1110 **Figure 7.** Along-strike variability for different faults (see Figure 1 and Table S1 for location). The
1111 suggested throw-rate profile used in our calculation for the fault throw rate is also shown.

1112 **Figure 8.** Cumulative distribution curves of disagreement δ . Bottom left: the cumulative distributions for
1113 all possible data pairs (on the same fault) are shown. Top right: distribution curves of disagreement δ
1114 disaggregated by methods. The red circles describe the fraction of highly-disagreeing data pairs
1115 ($\delta \geq 0.95$), defined in the text with the parameter β .

1116 **Figure 9.** The variability of β (fraction of $\delta \geq 0.95$ data pairs) in space and time. The average value of
1117 $\beta = 0.13$ (see Figure 8) is shown with a dotted black line, whereas blue points report the interval-
1118 average fractions.

1119 **Figure 10.** Forward modeling of throw rates using different total uncertainties $\sigma'_{\Delta F}$. The optimal model
1120 B ($\sigma'_{\Delta F} = \Delta t \cdot 0.10$ mm/yr) minimizes the differences between real-data (black) and synthetic-data (grey)
1121 disagreement curves at fixed $\sigma'_{\Delta F}$. This minimization implies that collected measurements are tectonic
1122 but affected by a total throw uncertainty $\sigma'_{\Delta F}$ greater than the observational one $\sigma_{\Delta F}$. Instead, the
1123 difference between the black and red lines indicates the effect of using $\sigma'_{\Delta F}$ (total uncertainty) to $\sigma_{\Delta F}$
1124 (observation error) in determining the probability density function of the long-term throw rates (p_L).

1125 **Figure 11.** Top: Resulting uniform throw rates (Top) and relative coefficient of variation COV (Bottom).
1126 For the fault name refer to Table 1.

1127 **Table 1.** Throw-rate median, 95% Confidence Interval and Coefficient of Variation (COV) for the central
1128 Apennines. See Figure 11 for linking fault ID to its location.

Figure 1.

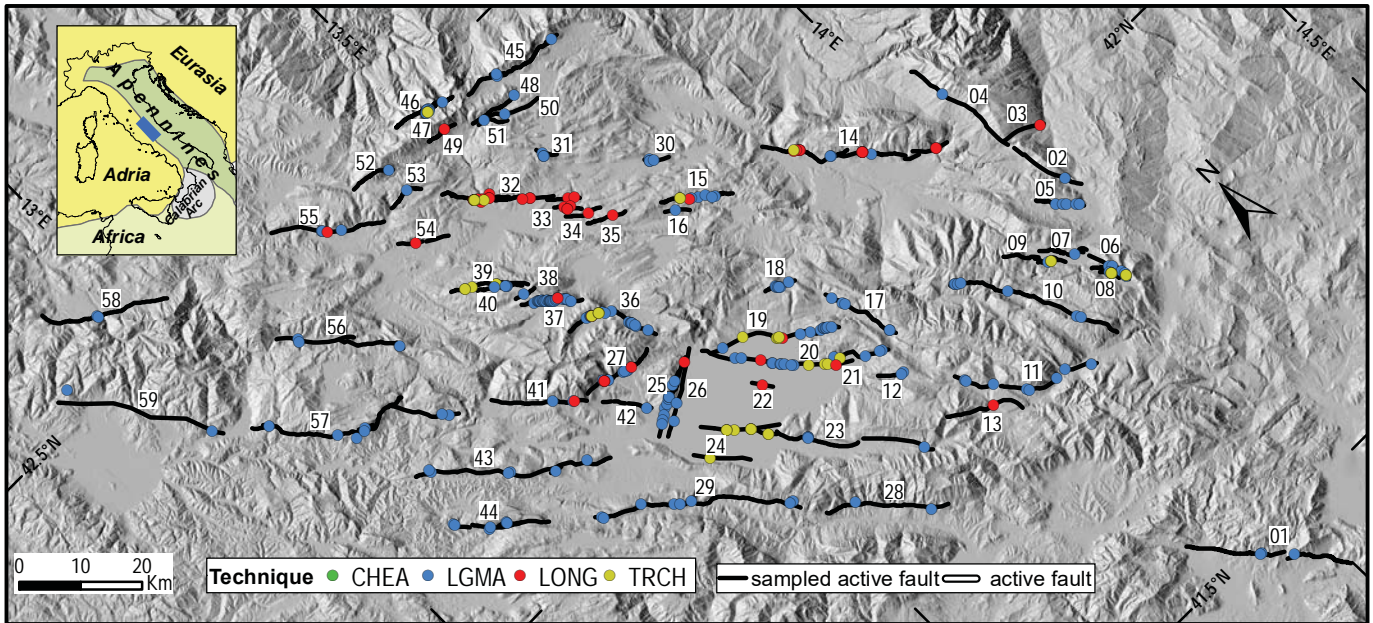
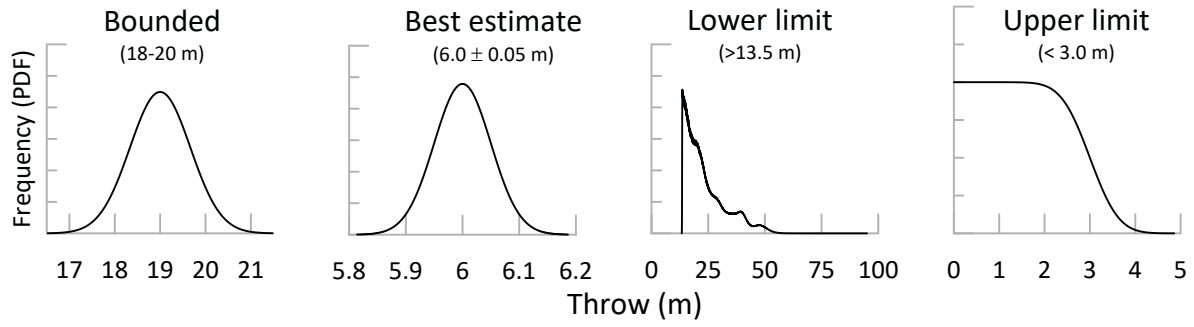


Figure 2.

THROW (ΔF) DISTRIBUTIONS



AGE (Δt) DISTRIBUTIONS

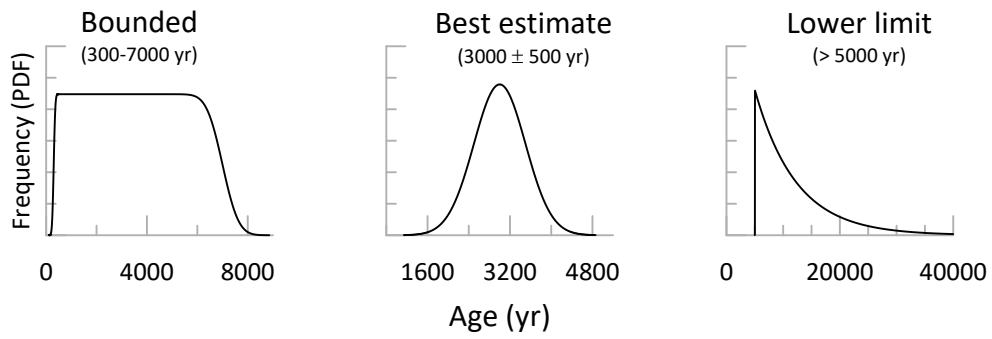


Figure 3.

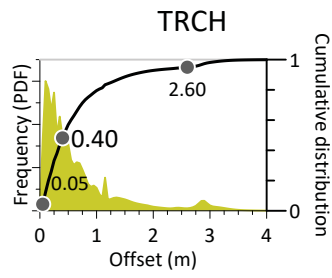
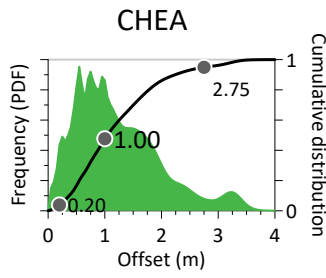
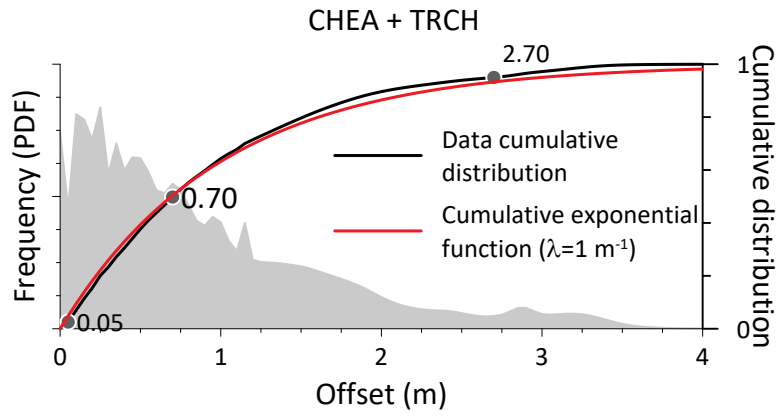


Figure 4.

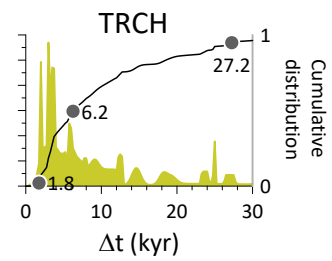
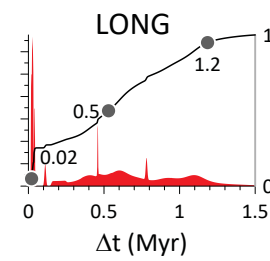
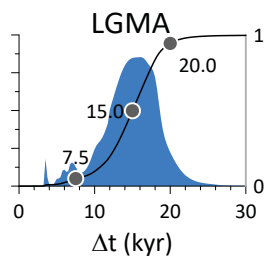
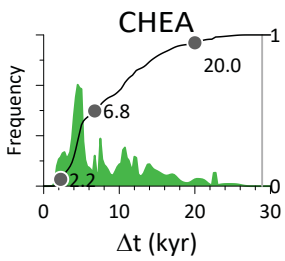
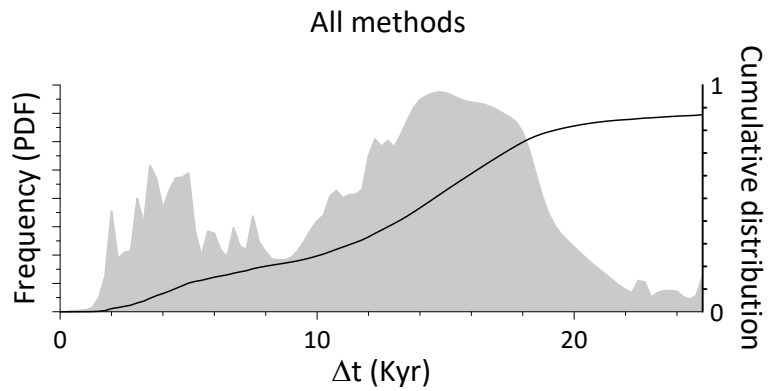


Figure 5.

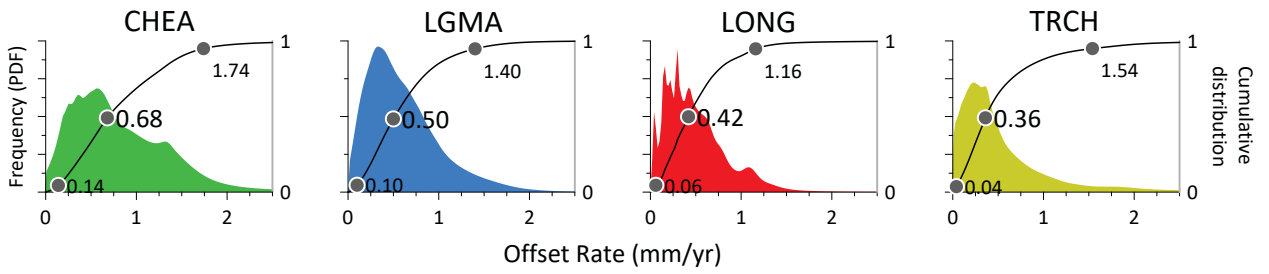
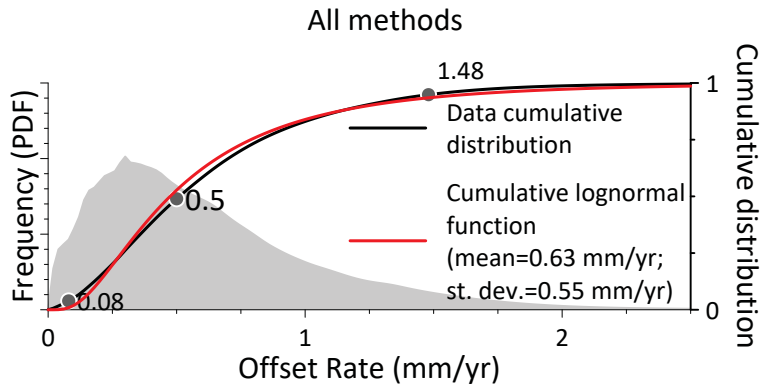


Figure 6.

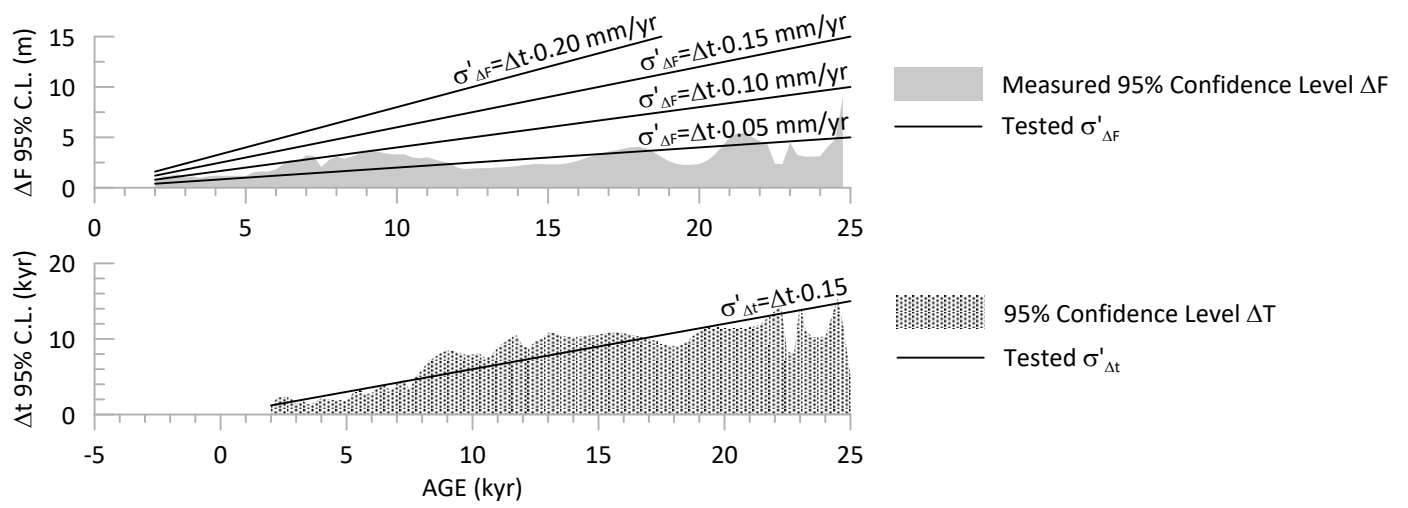


Figure 7.

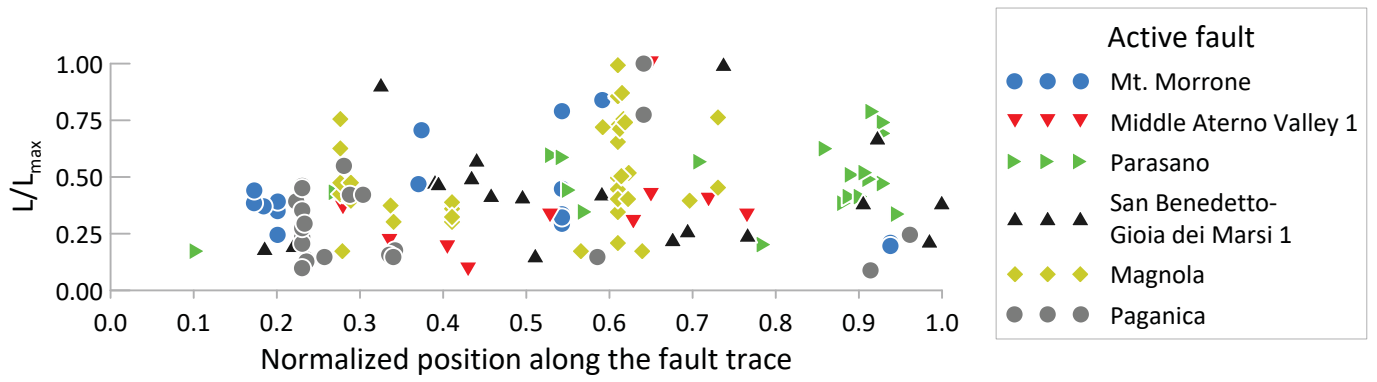


Figure 8.

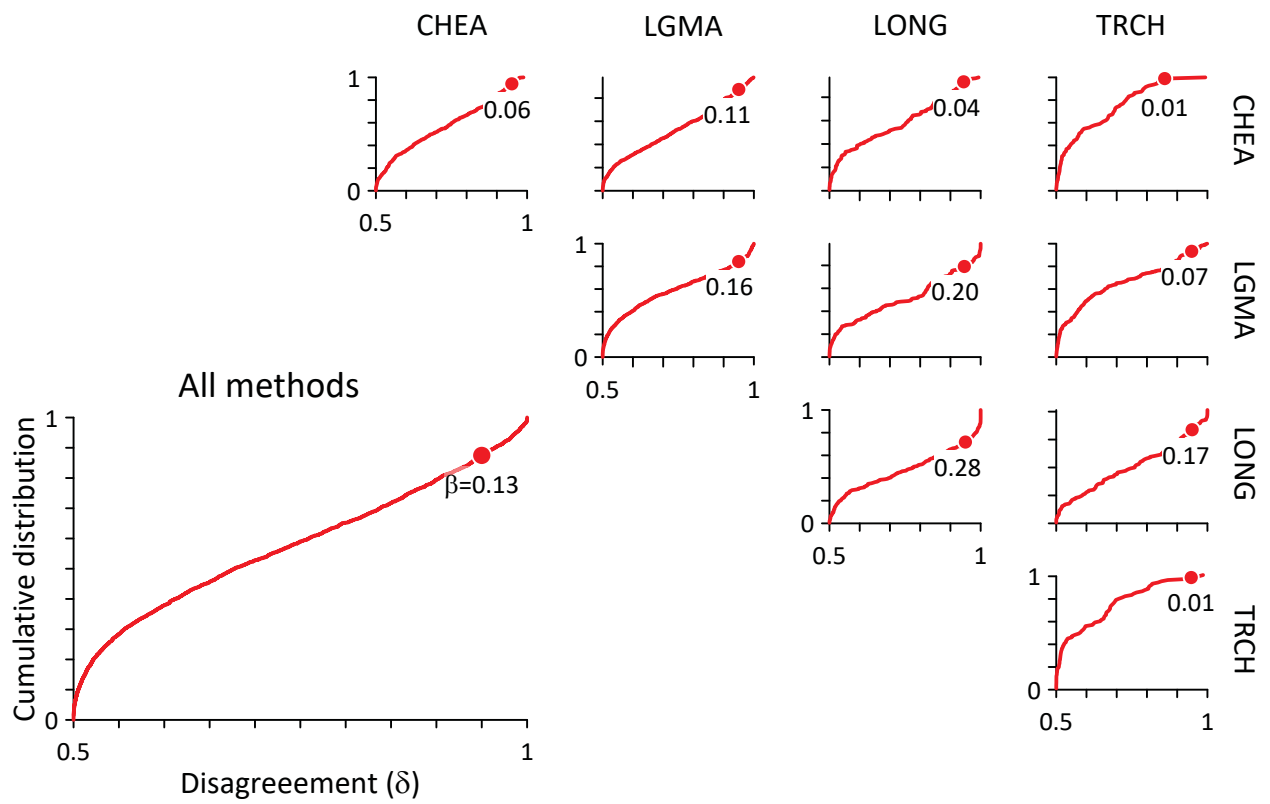


Figure 9.

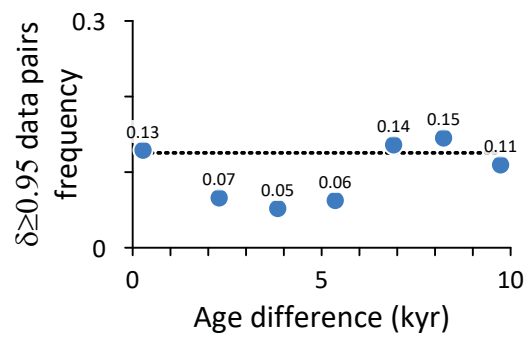
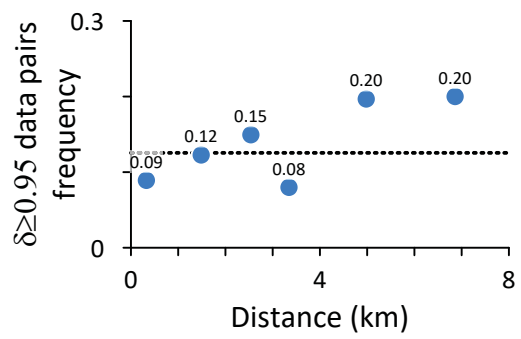


Figure 10.

A) $\sigma'_{\Delta F} = \Delta t \cdot 0.05$ mm/yr B) $\sigma'_{\Delta F} = \Delta t \cdot 0.10$ mm/yr C) $\sigma'_{\Delta F} = \Delta t \cdot 0.15$ mm/yr D) $\sigma'_{\Delta F} = \Delta t \cdot 0.20$ mm/yr

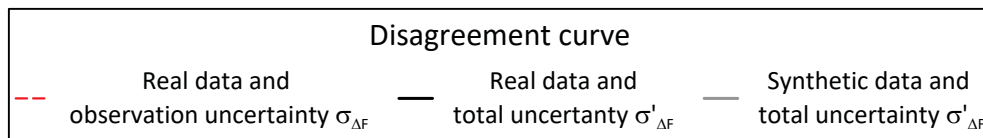
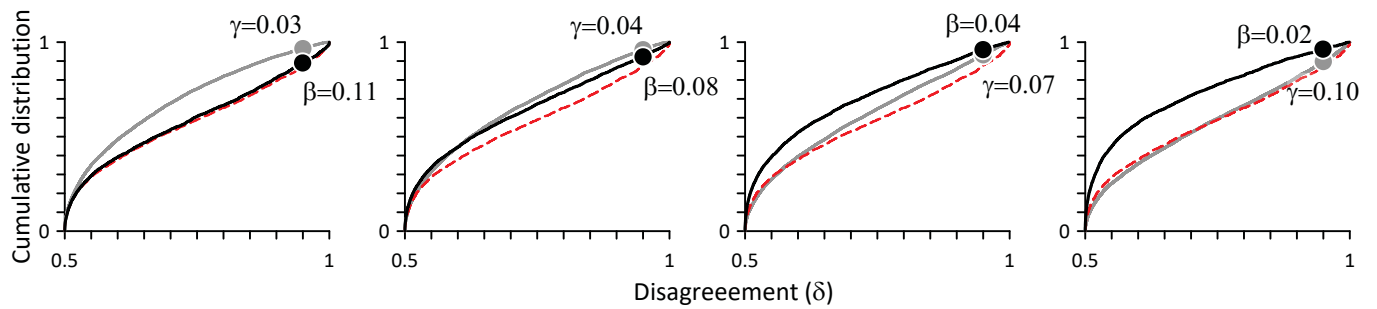


Figure 11.

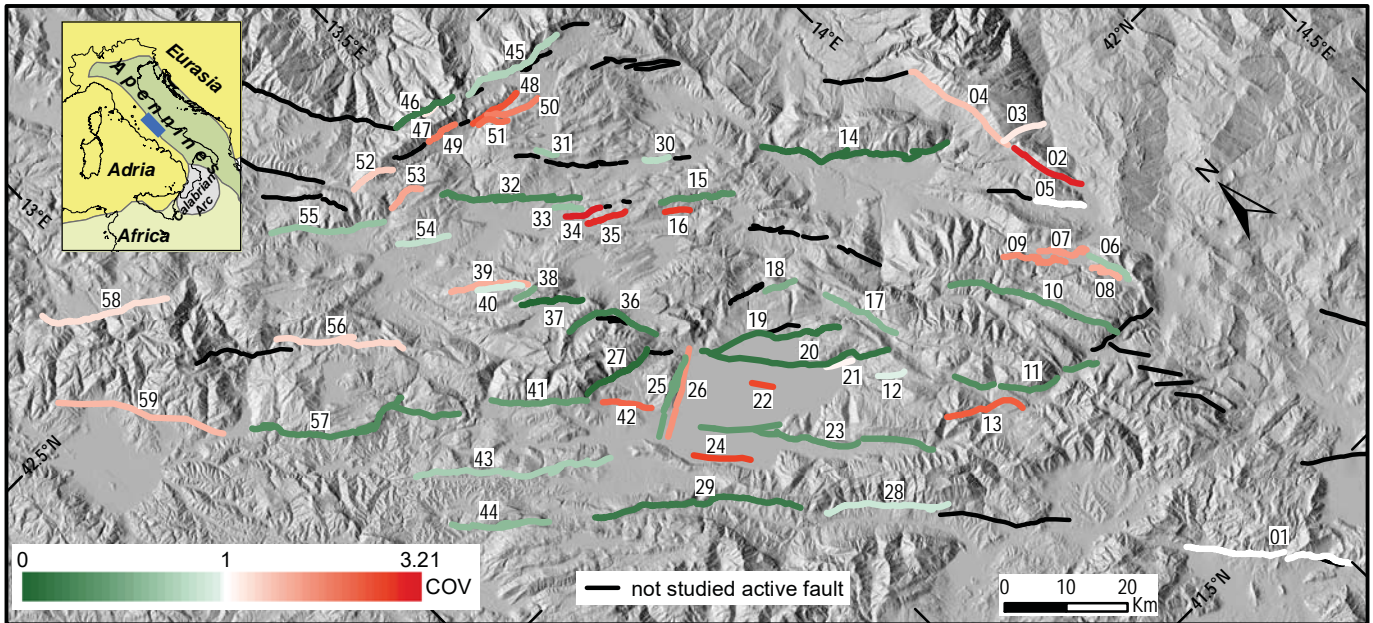
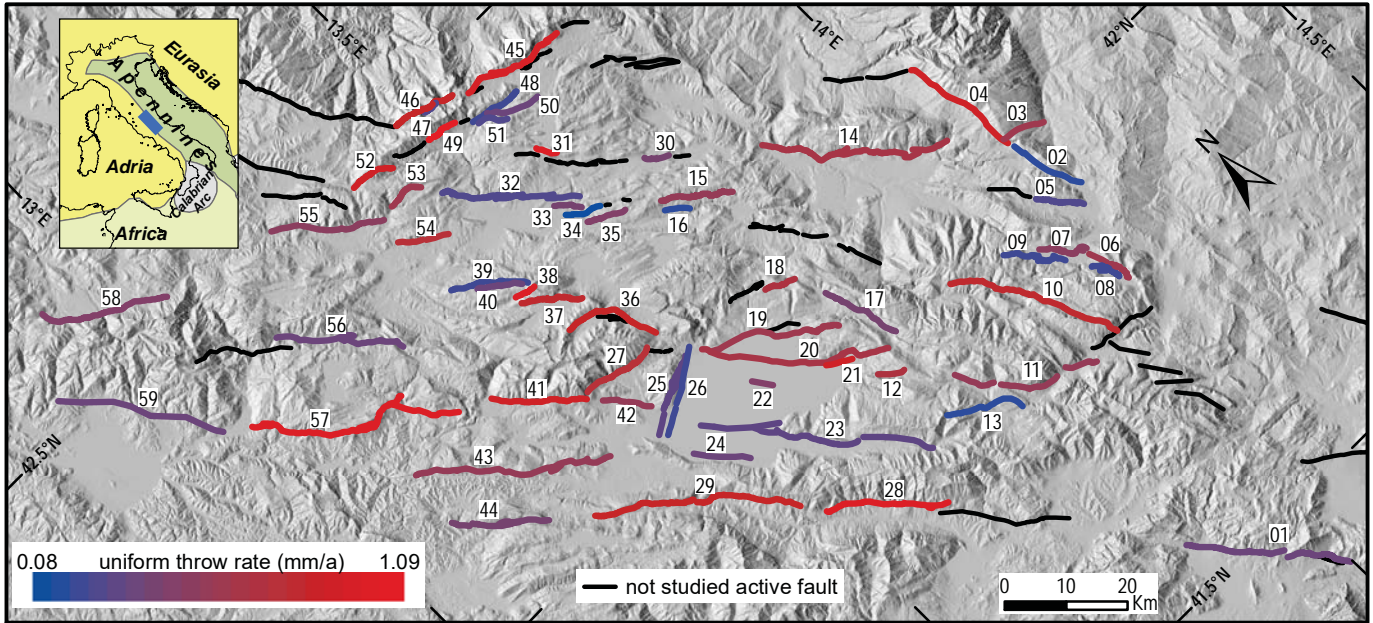


Table 1. Throw-rate median, 95% Confidence Interval and Coefficient of Variation (COV) for the central Apennines. See Figure 11 for linking fault ID to its location.

Fault ID	Fault name	Median (mm/yr)	95% Confidence Interval (mm/yr)	COV	Measurements (#)
1	Cassino	0.32	0.17-0.50	1.00	3
2	Porrara	0.14	0.01-0.44	2.99	1
3	Maiella-Palena	0.47	0.26-0.75	1.05	1
4	Maiella-Caramanico	0.83	0.49-1.46	1.17	1
5	Pizzalto	0.24	0.13-0.37	1.00	8
6	Aremogna-Cinque Miglia 1	0.46	0.31-0.63	0.70	7
7	Aremogna-Cinque Miglia 2	0.41	0.10-0.73	1.55	1
8	Aremogna-Cinque Miglia 3	0.21	0.06-0.37	1.47	5
9	Aremogna-Cinque Miglia 4	0.21	0.05-0.41	1.69	3
10	Scanno	0.75	0.58-0.95	0.48	8
11	Sangro-Mt. Marsicano 1	0.44	0.34-0.55	0.46	9
12	Sangro-Mt. Marsicano 2	0.52	0.31-0.80	0.94	2
13	Sangro-Mt. Marsicano 3	0.16	0.02-0.40	2.44	1
14	Mt. Morrone	0.49	0.41-0.57	0.32	16
15	Middle Aterno Valley 1	0.43	0.33-0.53	0.47	14
16	Middle Aterno Valley 2	0.16	0.01-0.47	2.80	1
17	San Sebastiano	0.35	0.26-0.45	0.55	11
18	Ventrino	0.55	0.41-0.71	0.54	7
19	Parasano	0.46	0.38-0.54	0.34	20
20	San Benedetto-Gioia dei Marsi 1	0.51	0.43-0.60	0.34	19
21	San Benedetto-Gioia dei Marsi 2	0.83	0.51-1.35	1.01	1
22	Bacinetto	0.38	0.08-1.10	2.66	1
23	Trasacco	0.29	0.22-0.36	0.48	20
24	Luco dei Marsi	0.31	0.03-0.88	2.74	1
25	Tre Monti 1	0.29	0.22-0.36	0.50	17
26	Tre Monti 2	0.20	0.05-0.37	1.62	2
27	Magnola	0.64	0.58-0.71	0.19	42
28	Liri 1	0.98	0.65-1.43	0.80	2
29	Liri 2	0.82	0.68-0.98	0.37	8
30	Barisciano-San Pio delle Camere 1	0.37	0.24-0.52	0.77	4
31	Barisciano-San Pio delle Camere 2	1.02	0.69-1.40	0.70	3
32	Paganica	0.23	0.18-0.28	0.44	23
33	San Demetrio	0.37	0.25-0.49	0.64	4
34	Middle Aterno northern splay 1	0.08	0.00-0.26	3.21	1
35	Middle Aterno northern splay 2	0.39	0.08-1.17	2.82	1
36	Ovindoli-Pezza	0.81	0.65-0.99	0.42	26
37	Campo Felice	0.74	0.68-0.81	0.18	39
38	Campo Felice-northern splay	0.97	0.71-1.23	0.53	1
39	Campo Felice-Colle Cerasitto 1	0.21	0.08-0.39	1.47	5
40	Campo Felice-Colle Cerasitto 2	0.31	0.18-0.46	0.89	4
41	Velino 1	0.85	0.59-1.01	0.49	7
42	Velino 2	0.43	0.15-1.09	2.21	1
43	Scurcola	0.44	0.29-0.60	0.72	6

44	Carsoli	0.36	0.26-0.49	0.63	6
45	Campo Imperatore 1	0.94	0.62-1.31	0.73	3
46	Campo Imperatore 2	0.80	0.65-0.94	0.37	5
47	Campo Imperatore 3	0.26	0.13-0.40	1.05	5
48	Assergi 1	0.21	0.03-0.55	2.45	1
49	Assergi 2	1.09	0.75-3.08	2.15	1
50	Montecristo 1	0.32	0.07-0.69	1.97	1
51	Montecristo 2	0.26	0.04-0.61	2.20	1
52	Mt. San Franco	0.84	0.50-1.47	1.16	1
53	Stabiata	0.47	0.19-0.91	1.54	1
54	Pettino	0.64	0.40-0.90	0.79	1
55	Pizzoli	0.42	0.29-0.57	0.64	4
56	Sella di Corno	0.31	0.14-0.49	1.13	3
57	Fiamignano	0.98	0.77-1.23	0.46	9
58	Leonessa	0.39	0.21-0.64	1.09	2
59	Rieti	0.32	0.14-0.55	1.26	2